

Part 1: Information Theory Basics

L02: Basics of Probability (1/2)

[Random experiment, independence, conditional probability, chain rule, Bayes' theorem, random variables]

Wolfgang Gatterbauer

cs7840 Foundations and Applications of Information Theory (fa25)

<https://northeastern-datalab.github.io/cs7840/fa25/>

9/11/2025

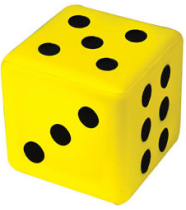
Pre-class conversations

- Last class recapitulation
- Organizational matters: Does every one get Piazza messages?
- First scribe on Piazza. Awesome! I will look over the weekend
- Office hours: Usually right after class, or via email / Teams. Also feel free talk to me regularly during break/ after class about project ideas, or papers/research directions to include towards the end
- New class arrivals
- Today:
 - The basics of probability theory

Basics of probability theory

Following slides are built upon examples by Jay Aslam from earlier editions of this class. For a more extensive cover of the basics, I recommend "Bertsekas, Tsitsiklis. Introduction to Probability, 2008." It's a solid textbook on probability theory and I regularly find myself going back to this book to look up basic concepts. Working by yourself through chapter 1 on "Sample Space and Probability" is a good investment of your time.

Sample space / outcomes / events / probability



PROBABILITY

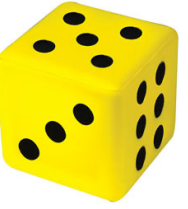
- A random experiment
- Sample space Ω : set of all possible outcomes

EXAMPLE 1:

roll a fair die with 6 sides

$\Omega =$?

Sample space / outcomes / events / probability



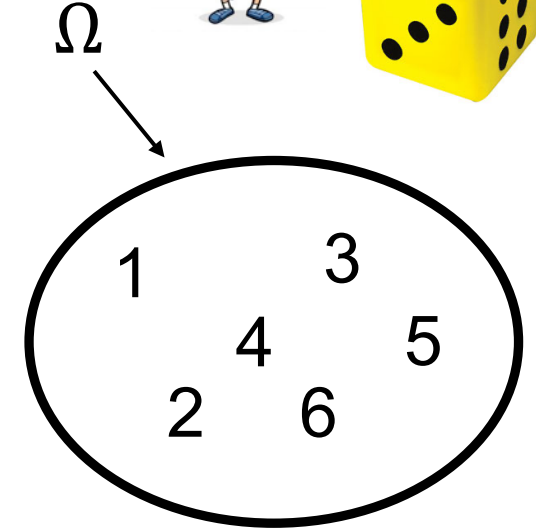
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- A random experiment
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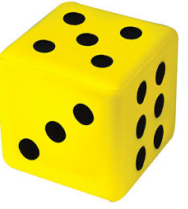
EXAMPLE 1:

roll a fair die with 6 sides

$$\Omega = \{1, 2, 3, 4, 5, 6\}$$



Sample space / outcomes / events / probability



PROBABILITY

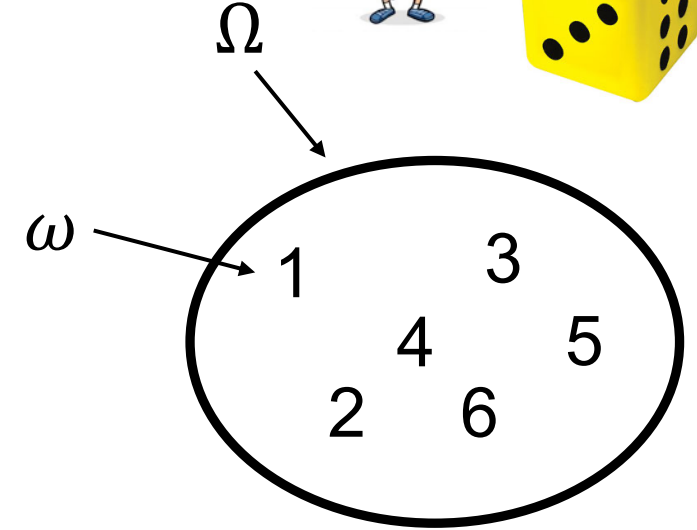
- A random experiment
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$$\Omega = \{1, 2, 3, 4, 5, 6\}$$

Say the outcome is a 1



Sample space / outcomes / events / probability



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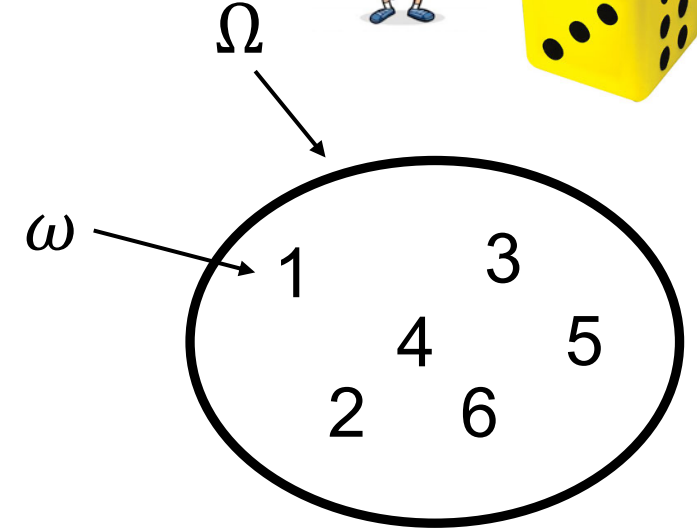
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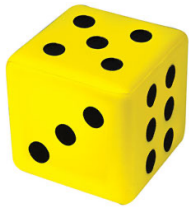
Say the outcome is a 1

$$E_1 = \text{"even"}$$

$$E_2 = \text{"} \geq 3 \text{"}$$



Sample space / outcomes / events / probability



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- Probability measure:

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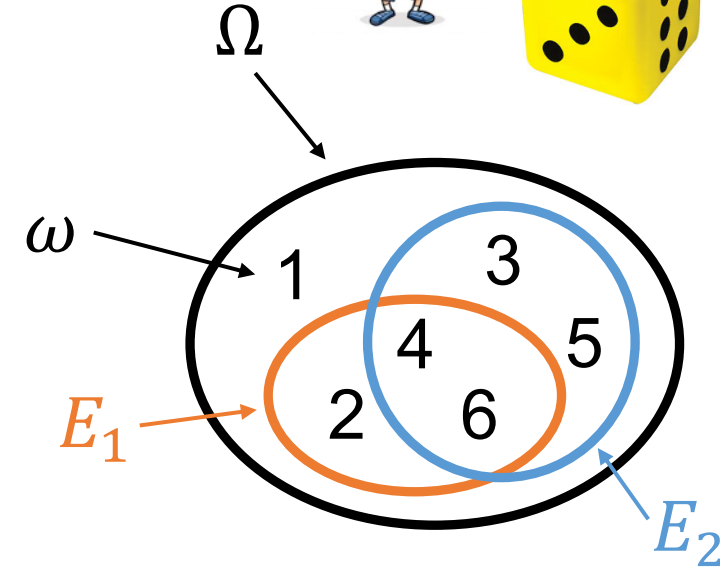
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$$\Omega = \{1, 2, 3, 4, 5, 6\}$$

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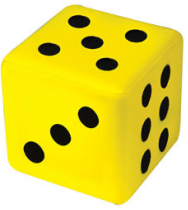
$$E_1 = \text{"even"} = \{2, 4, 6\}$$

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?

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 ~~$\mathbb{P}: \Omega \rightarrow [0, 1]$~~

?

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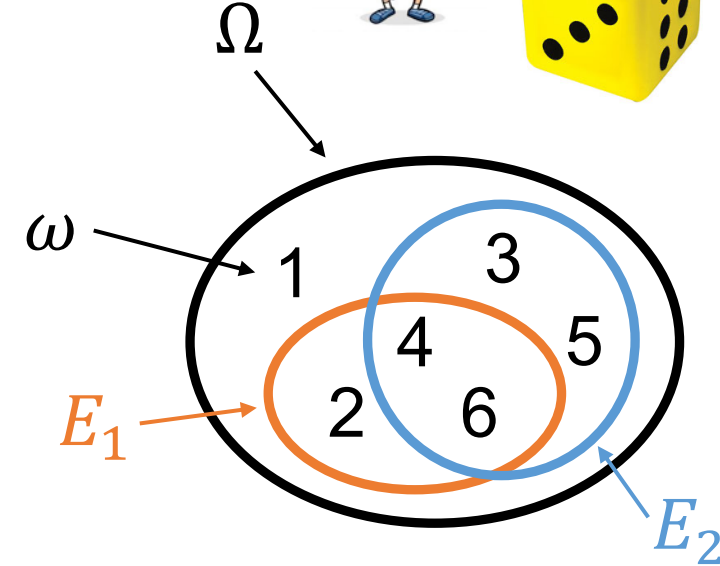
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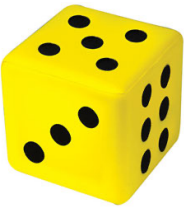
Say the outcome is a 1

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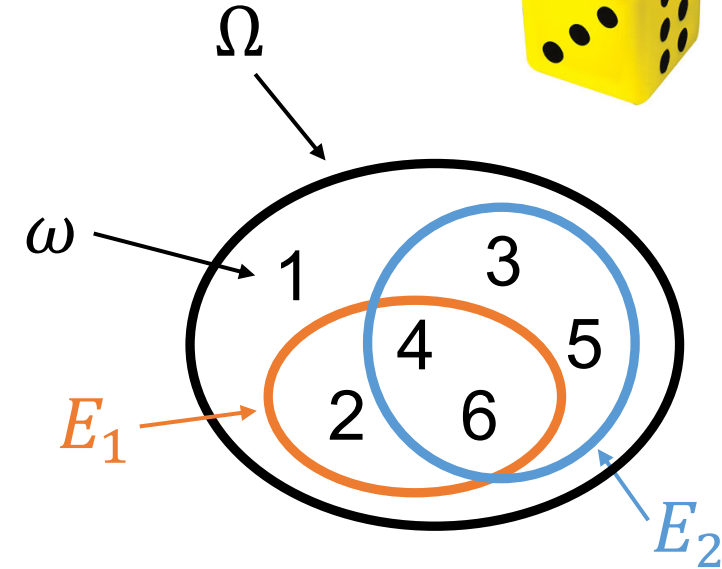
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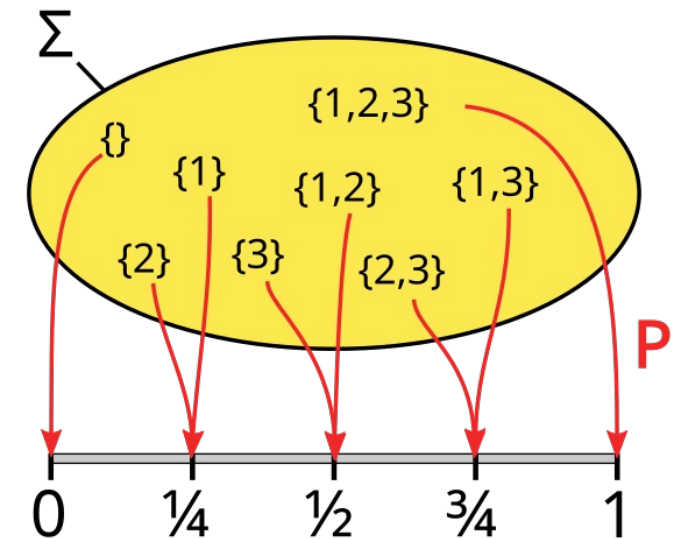
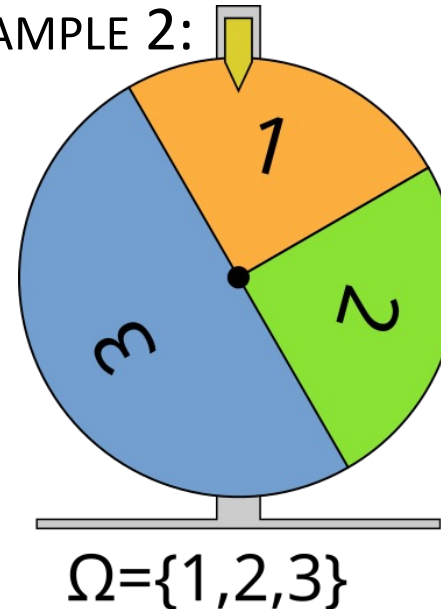
Say the outcome is a 1

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EXAMPLE 2:



The set Σ of all events $E \in \Sigma$ is a σ -algebra, basically the same as the power set 2^Ω , except for some pathological non-measurable sets like the Vitali-Set (https://www.youtube.com/watch?v=hs3eDa3_DzU). See also: <https://en.wikipedia.org/wiki/%CE%A3-algebra>

Source of figure to the right: https://en.wikipedia.org/wiki/Probability_measure

Gatterbauer. Foundations and Applications of Information Theory: <https://northeastern-datalab.github.io/cs7840/>

Sample space / outcomes / events / probability



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- Probability measure:

~~$$\mathbb{P}: \Omega \rightarrow [0, 1]$$~~

$$\mathbb{P}: \Sigma \rightarrow [0, 1]$$

also $\sum_{\omega \in \Omega} \mathbb{P}(\{\omega\}) = 1$

- 1) Nonnegativity: $\mathbb{P}(E) \geq 0, \forall E \in \Sigma$
- 2) Normalization: $\mathbb{P}(\Omega) = 1$
- 3) Additivity: if E_1 and E_2 are disjoint events, then $\mathbb{P}(E_1 \cup E_2) = \mathbb{P}(E_1) + \mathbb{P}(E_2)$,

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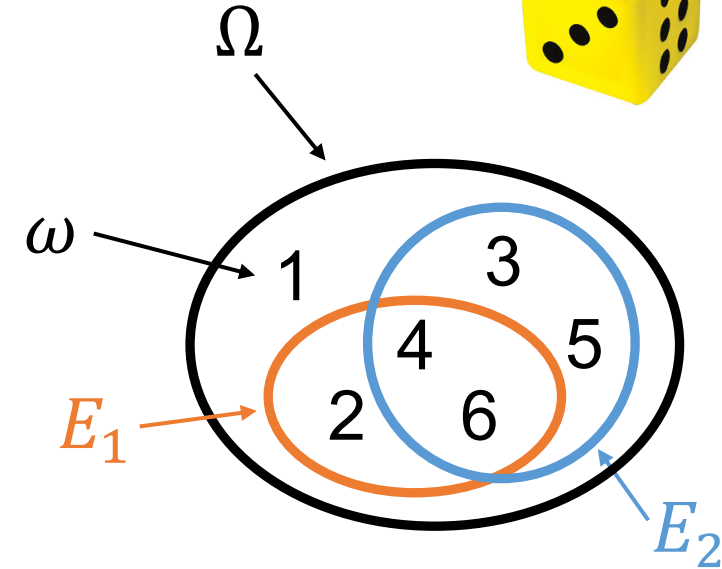
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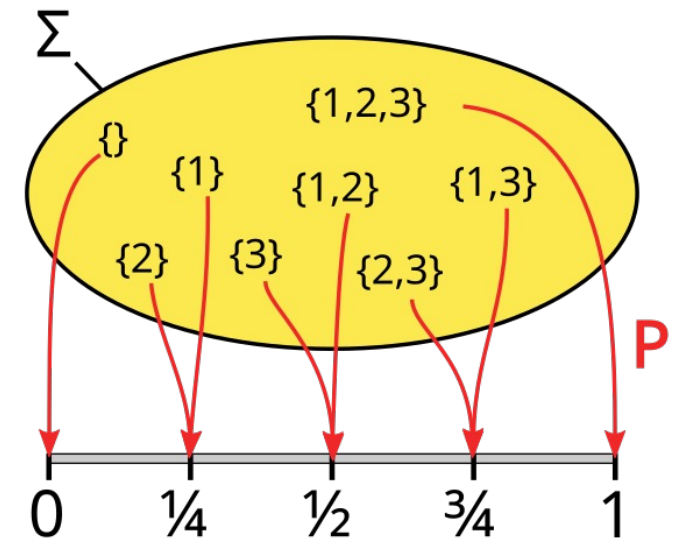
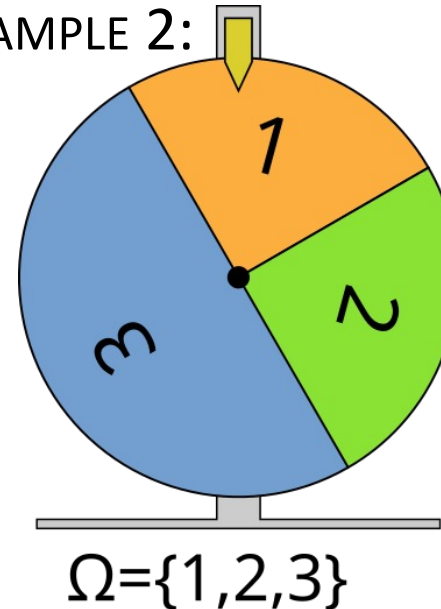
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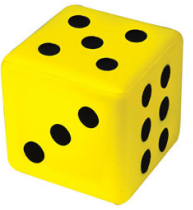
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EXAMPLE 2:



Sample space / outcomes / events / probability



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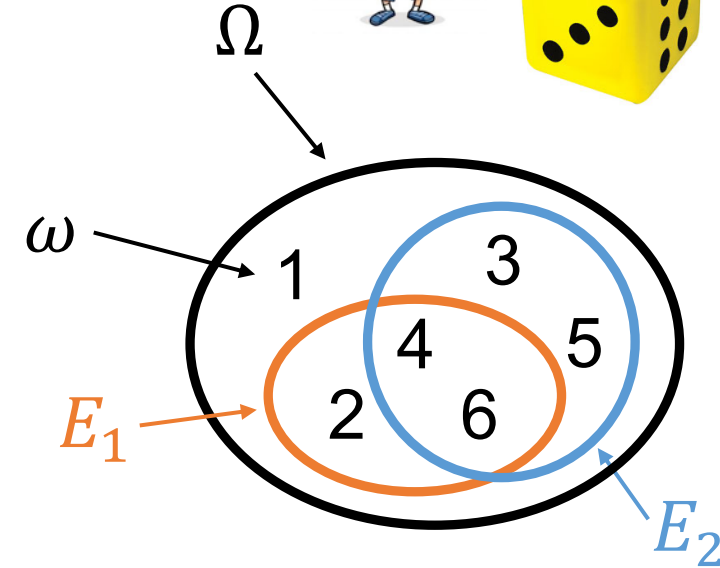
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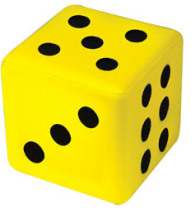
$$E_2 = \text{"} \geq 3 \text{"} = \{3, 4, 5, 6\}$$



short for $\mathbb{P}(\{\omega\})$

- If $\mathbb{P}(\omega) = \frac{1}{|\Omega|}, \forall \omega \in \Omega$, then $\mathbb{P}(E) = ?$

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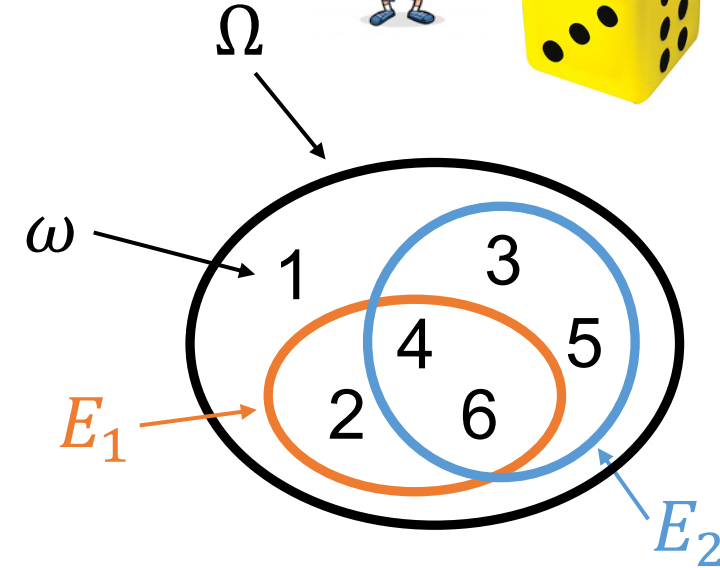
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- If $\mathbb{P}(\omega) = \frac{1}{|\Omega|}, \forall \omega \in \Omega$, then $\mathbb{P}(E) = \frac{|E|}{|\Omega|}$

$$\mathbb{P}(1) = \mathbb{P}(2) = \dots = \mathbb{P}(6) = \frac{1}{6}$$

$$\mathbb{P}(E_1) =$$

$$\mathbb{P}(E_2) =$$

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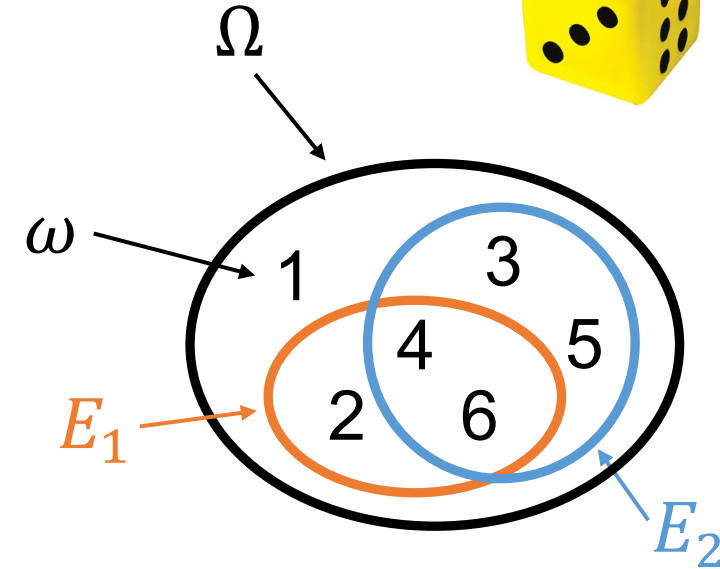
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$$\mathbb{P}(1) = \mathbb{P}(2) = \dots = \mathbb{P}(6) = \frac{1}{6}$$

$$\mathbb{P}(E_1) = \mathbb{P}(2) + \mathbb{P}(4) + \mathbb{P}(6) = \frac{3}{6} = \frac{1}{2}$$

$$\mathbb{P}(E_2) = \frac{4}{6} = \frac{2}{3}$$

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roll two fair dice with 6 sides

$\Omega =$

?

$|\Omega| =$

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EXAMPLE 3:

roll two fair dice with 6 sides

$$\begin{aligned}\Omega &= \{1, 2, 3, 4, 5, 6\} \times \{1, 2, 3, 4, 5, 6\} \\ &= \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}\end{aligned}$$

$$|\Omega| = 36$$

		second die					
		1	2	3	4	5	6
first die	1						
	2						
	3						
	4						
	5						
	6						

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		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

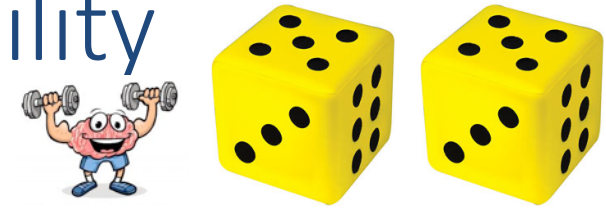
$$E_1 = \text{"sum is 7"}$$

$$|E_1| =$$

$$\mathbb{P}(E_1) =$$



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$$|\Omega| = 36$$

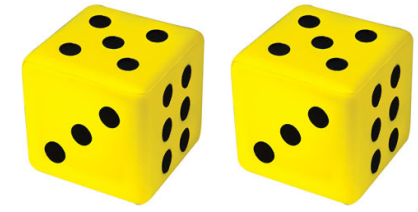
		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
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	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

$$\begin{aligned}E_1 &= \text{"sum is 7"} \\ &= \{(1,6), (2,5), (3,4), (4,3), (5,2), (6,1)\}\end{aligned}$$

$$|E_1| = 6$$

$$\mathbb{P}(E_1) = \frac{|E_1|}{|\Omega|} = \frac{6}{36} = \frac{1}{6}$$

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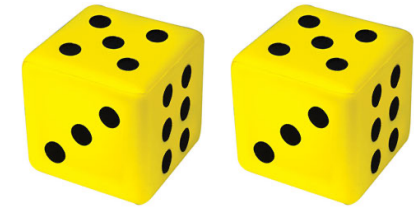
$$E_2 = \text{"total is greater than 8"}$$

$$|E_2| =$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|}$$



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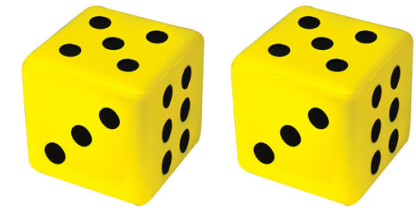
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	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

$$\begin{aligned}E_2 &= \text{"total is greater than 8"} \\ &= \text{"total is 9 or 10 or 11 or 12"}\end{aligned}$$

$$|E_2| =$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|}$$

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$$|\Omega| = 36$$

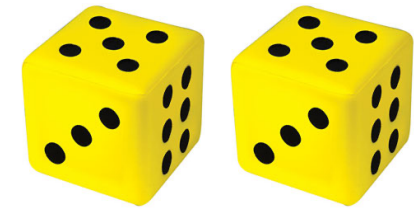
		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

$$\begin{aligned}E_2 &= \text{"total is greater than 8"} \\ &= \text{"total is 9 or 10 or 11 or 12"}\end{aligned}$$

$$|E_2| = 4$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|}$$

Sample space / outcomes / events / probability



PROBABILITY

- A random experiment
- Sample space Ω : set of all possible outcomes
- An outcome $\omega \in \Omega$ of the exp.
- Event E : a subset of the sample space $E \subseteq \Omega$
- Probability measure:
 $\mathbb{P}: \Sigma \rightarrow [0, 1]$

EXAMPLE 3:

roll two fair dice with 6 sides

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$$|\Omega| = 36$$

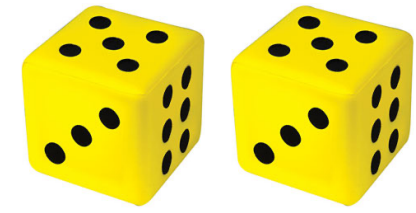
		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
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	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

$$\begin{aligned}E_2 &= \text{"total is greater than 8"} \\ &= \text{"total is 9 or 10 or 11 or 12"}\end{aligned}$$

$$|E_2| = \begin{array}{c} \downarrow \quad \downarrow \\ 4 + 3 \end{array}$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|}$$

Sample space / outcomes / events / probability



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$$|\Omega| = 36$$

		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
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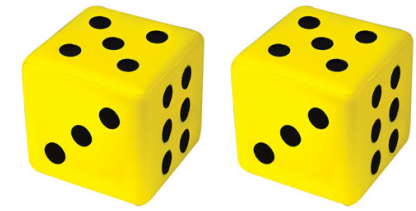
E_2 = "total is greater than 8"

= "total is 9 or 10 or 11 or 12"

$$|E_2| = \begin{array}{c} \downarrow \quad \downarrow \quad \downarrow \\ 4 + 3 + 2 \end{array}$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|}$$

Sample space / outcomes / events / probability



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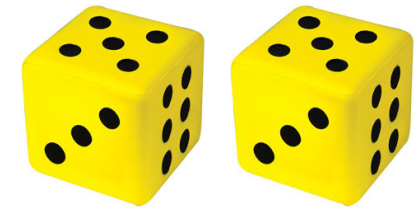
E_2 = "total is greater than 8"

= "total is 9 or 10 or 11 or 12"

$$|E_2| = \begin{array}{ccccccc} & & \downarrow & \downarrow & \downarrow & \downarrow & \\ & & 4 & + & 3 & + & 2 & + & 1 \end{array}$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|}$$

Sample space / outcomes / events / probability



PROBABILITY

- A random experiment
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E_2 = "total is greater than 8"

= "total is 9 or 10 or 11 or 12"

additivity of
disjoint events

$$|E_2| = \begin{array}{ccccccc} & \downarrow & \downarrow & \downarrow & \downarrow & & \\ & 4 & + & 3 & + & 2 & + & 1 & = & 10 \end{array}$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|} = \frac{10}{36} = \frac{5}{18}$$

Conditional Probabilities / Independence

Conditional probability

DEFINITION

Conditional probability $\mathbb{P}(A|B)$ is the probability that event A occurred, given that event B occurred

EXAMPLE 4
(FAIR 16-SIDED DIE):

Ω



1	2	3	5	6	7	10		11
13				16				
14	4	12		8	9		15	

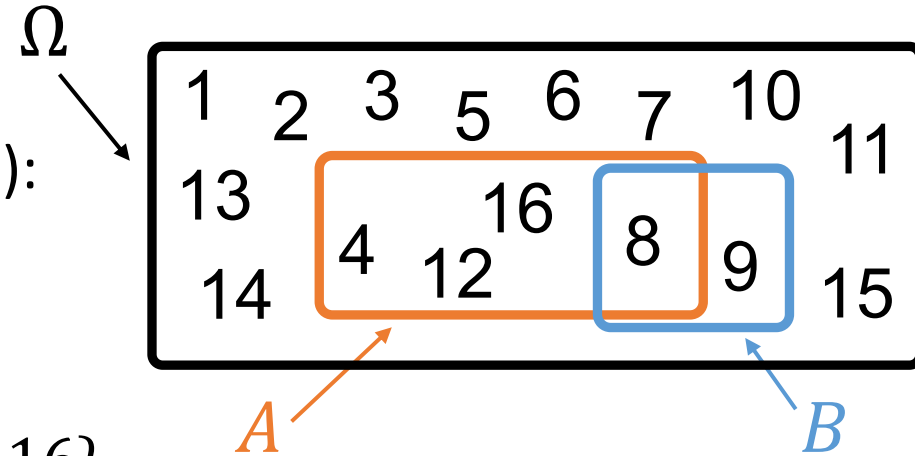
$$\Omega = \{1, 2, \dots, 16\}$$

Conditional probability

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Conditional probability $\mathbb{P}(A|B)$ is the probability that event A occurred, given that event B occurred

EXAMPLE 4
(FAIR 16-SIDED DIE):

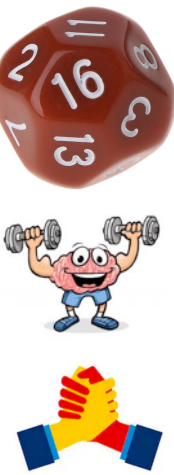


$$\Omega = \{1, 2, \dots, 16\}$$

$$A = \text{"divisible by 4"} = \{4, 8, 12, 16\}$$

$$B = \text{"center numbers"} = \{8, 9\}$$

$$\mathbb{P}(A|B) = ?$$

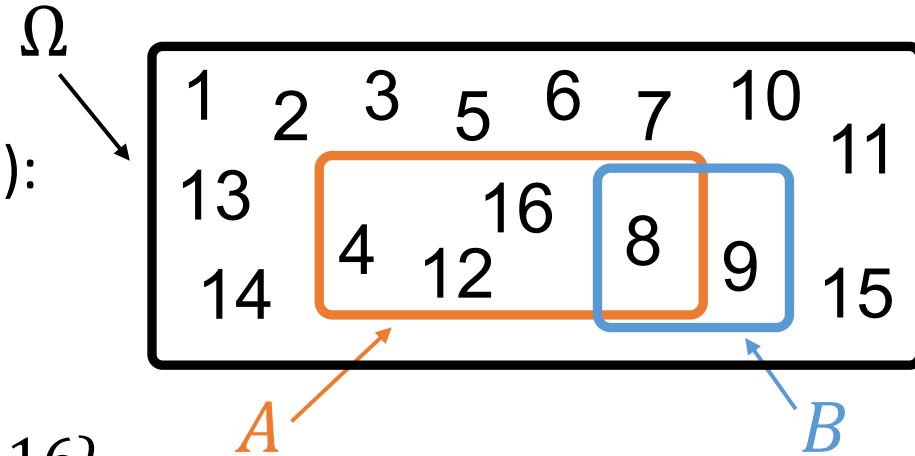


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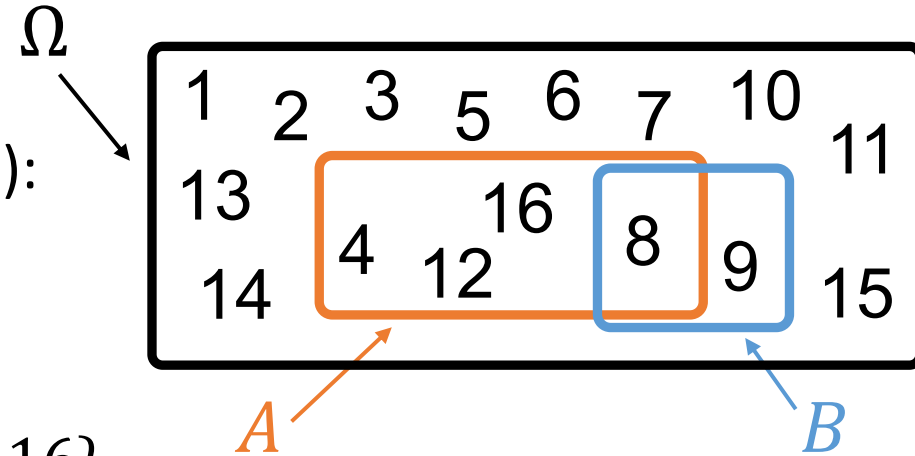
$$\mathbb{P}(A|B) = \frac{1}{2}$$

Conditional probability

DEFINITION

Conditional probability $\mathbb{P}(A|B)$ is the probability that event A occurred, given that event B occurred

EXAMPLE 4
(FAIR 16-SIDED DIE):



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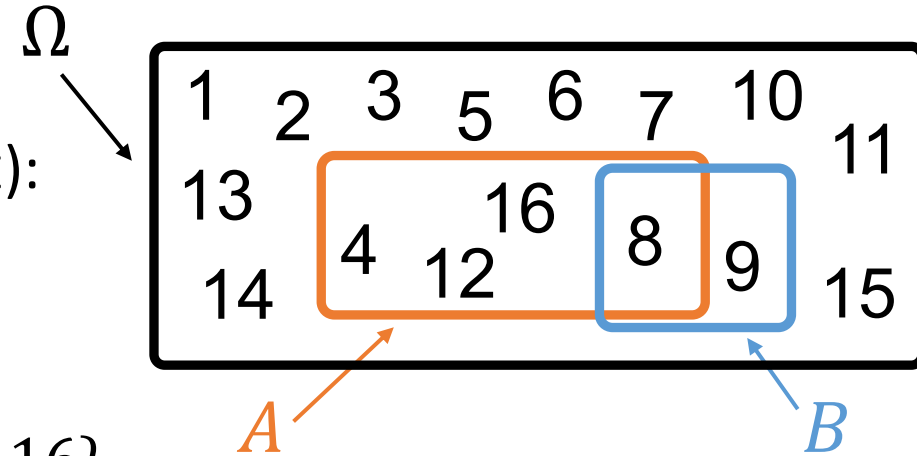
$$\mathbb{P}(A|B) = \frac{\text{Number of outcomes in } A \cap B}{\text{Number of outcomes in } B}$$

Conditional probability

DEFINITION

Conditional probability $\mathbb{P}(A|B)$ is the probability that event A occurred, given that event B occurred

EXAMPLE 4
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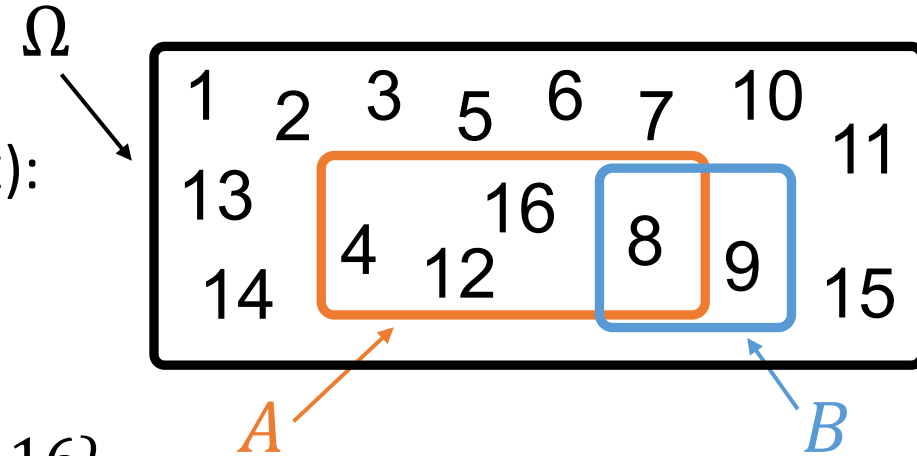
$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

Conditional probability

DEFINITION

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Conditional probability

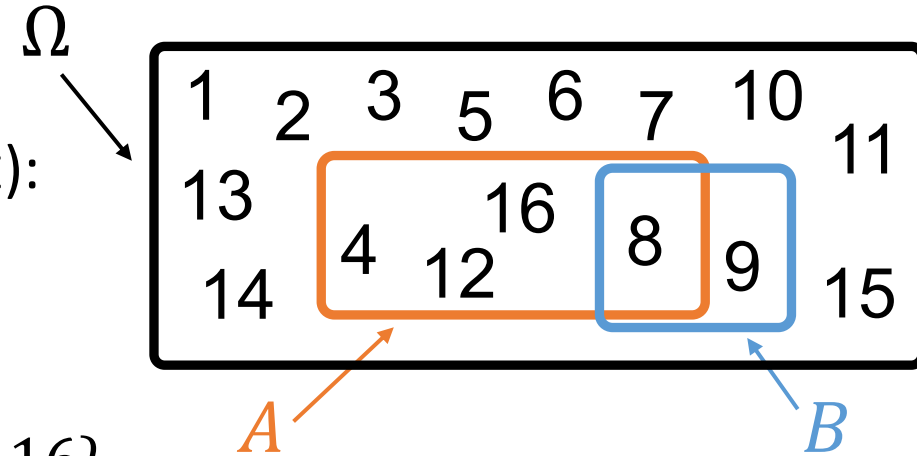
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$\mathbb{P}(A \cap B)$

EXAMPLE 4
(FAIR 16-SIDED DIE):



$$\Omega = \{1, 2, \dots, 16\}$$

$$A = \text{"divisible by 4"} = \{4, 8, 12, 16\}$$

$$B = \text{"center numbers"} = \{8, 9\}$$

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)} = \frac{\frac{1}{16}}{\frac{2}{16}} = \frac{1}{2}$$

Independence

DEFINITION:

Two events are independent if the probability that one occurred is not affected by knowledge that the other occurred.

$$\mathbb{P}(A|B) = \mathbb{P}(A) \quad (\text{if } \mathbb{P}(B) \neq 0)$$

$$\mathbb{P}(B|A) = \mathbb{P}(B) \quad (\text{if } \mathbb{P}(A) \neq 0)$$

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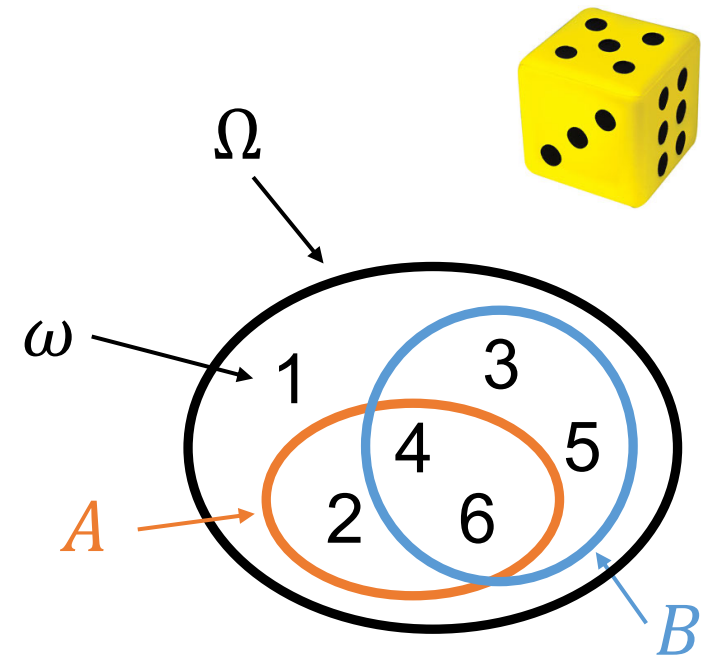
EXAMPLE 1 (CONTINUED):

roll a fair die with 6 sides
the outcome is a 1

$$\Omega = \{1, 2, 3, 4, 5, 6\}$$

$$A = \text{"even"} = \{2, 4, 6\}$$

$$B = \text{"}\geq 3\text{"} = \{3, 4, 5, 6\}$$



Are A and B independent ?

Independence

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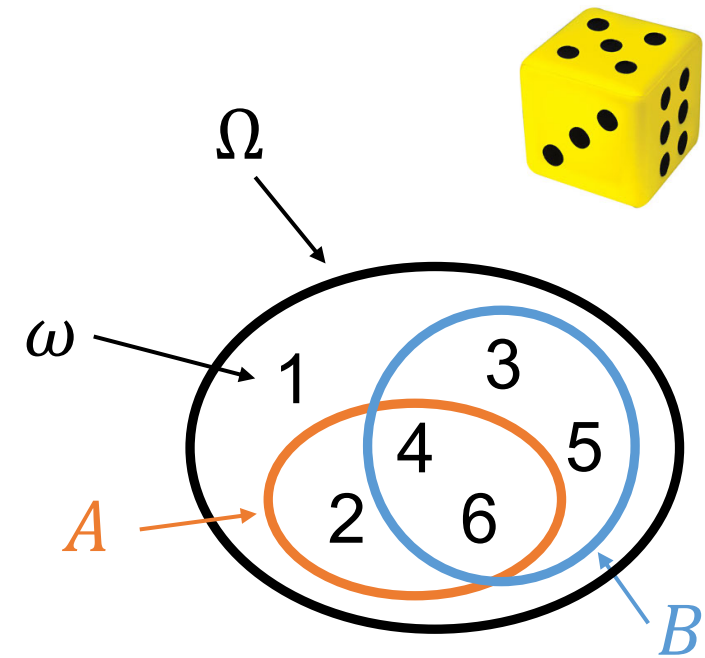
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A and B are independent !

$$\mathbb{P}(A) = \frac{1}{2}$$

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)} = \frac{1}{2}$$

1	3	5
2	4	6

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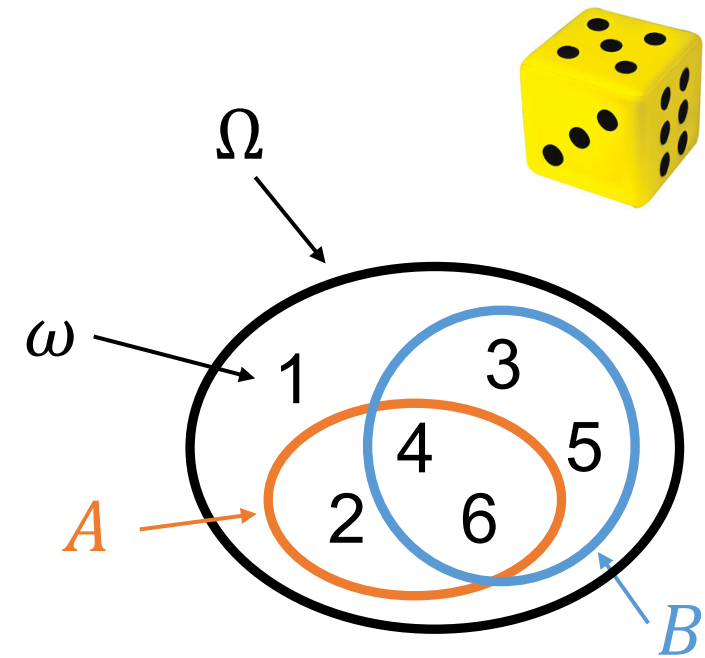
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A and B are independent !

$$\mathbb{P}(A) = \frac{1}{2}$$

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)} = \frac{1}{2}$$

$$\mathbb{P}(A) = \frac{1}{2}$$

		$\mathbb{P}(B) = \frac{2}{3}$	
	1	3	5
	2	4	6

*Think about a coordinate system:
x and y are independent*

Independence

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Recall the definition of cond. prob.:

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

CLAIM

Two events A and B are independent if and only if $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$

PROOF (assuming $\mathbb{P}(A) \neq 0, \mathbb{P}(B) \neq 0$)



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PROOF (assuming $\mathbb{P}(A) \neq 0, \mathbb{P}(B) \neq 0$)

($\Rightarrow$) If $\mathbb{P}(A|B) = \mathbb{P}(A)$:

?

($\Leftarrow$) If $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$:

?

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PROOF (assuming $\mathbb{P}(A) \neq 0, \mathbb{P}(B) \neq 0$)

($\Rightarrow$) If $\mathbb{P}(A|B) = \mathbb{P}(A)$:

$$\text{Then } \mathbb{P}(A, B) = \mathbb{P}(A|B) \cdot \mathbb{P}(B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$

($\Leftarrow$) If $\mathbb{P}(A, B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$:

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)} = \frac{\mathbb{P}(A) \cdot \cancel{\mathbb{P}(B)}}{\cancel{\mathbb{P}(B)}} = \mathbb{P}(A)$$

$\nearrow \neq 0$

Independence

DEFINITION:

Two events are independent if the probability that one occurred is not affected by knowledge that the other occurred.

$$\mathbb{P}(A|B) = \mathbb{P}(A) \quad (\text{if } \mathbb{P}(B) \neq 0)$$

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Recall the definition of cond. prob.:

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

Commonly used as definition of independence because it is slightly more general: if one event has probability 0 then the statement holds vacuously (while one conditional probability is undefined). And if both are not 0, then we just proved

DEFINITION: that to be equivalent to the earlier definition.

$$\text{Two events } A \text{ and } B \text{ are independent} \Leftrightarrow \mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$

PROOF (assuming $\mathbb{P}(A) \neq 0, \mathbb{P}(B) \neq 0$)

($\Rightarrow$) If $\mathbb{P}(A|B) = \mathbb{P}(A)$:

$$\text{Then } \mathbb{P}(A, B) = \mathbb{P}(A|B) \cdot \mathbb{P}(B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$

($\Leftarrow$) If $\mathbb{P}(A, B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$:

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)} = \frac{\mathbb{P}(A) \cdot \cancel{\mathbb{P}(B)}}{\cancel{\mathbb{P}(B)}} = \mathbb{P}(A)$$

$\nearrow$
 $\neq 0$

Independence

DEFINITION:

Two events A and B are independent if
 $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$

EXAMPLE 3 (CONTINUED):

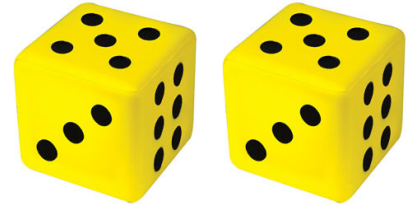
roll two fair dice with 6 sides

$$\Omega = \{1, 2, 3, 4, 5, 6\} \times \{1, 2, 3, 4, 5, 6\}$$

A = "1st roll is 1"

B = "sum is 7"

Are A and B independent? ?



Independence

DEFINITION:

Two events A and B are independent if
 $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$

EXAMPLE 3 (CONTINUED):

roll two fair dice with 6 sides

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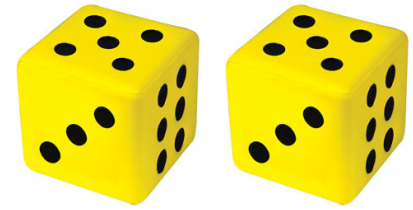
B = "sum is 7"

Are A and B independent?

$$\mathbb{P}(A) = ?$$

$$\mathbb{P}(B) = ?$$

$$\mathbb{P}(A, B) = ?$$



Independence

DEFINITION:

Two events A and B are independent if
 $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$

EXAMPLE 3 (CONTINUED):

roll two fair dice with 6 sides

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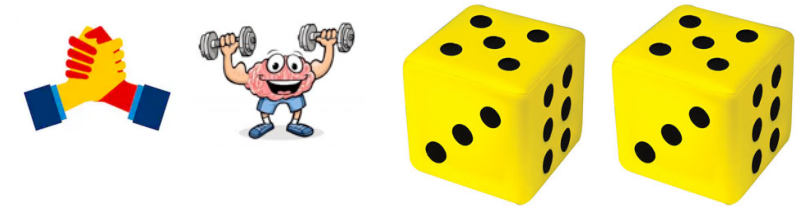
B = "sum is 7"

Are A and B independent?

$$\mathbb{P}(A) = \frac{1}{6}$$

$$\mathbb{P}(B) = \frac{1}{6}$$

$$\mathbb{P}(A, B) = ?$$



second die

	1	2	3	4	5	6
1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

first die

Independence

DEFINITION:

Two events A and B are independent if
 $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$

EXAMPLE 3 (CONTINUED):

roll two fair dice with 6 sides

$$\Omega = \{1, 2, 3, 4, 5, 6\} \times \{1, 2, 3, 4, 5, 6\}$$

A = "1st roll is 1"

B = "sum is 7"

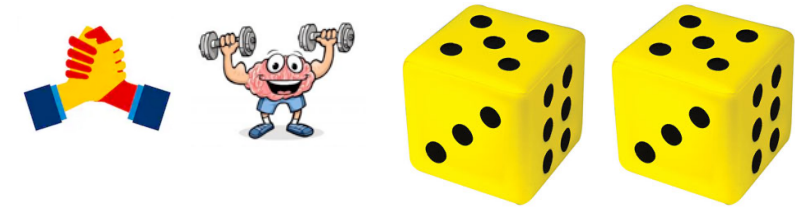
Are A and B independent?

$$\mathbb{P}(A) = \frac{1}{6}$$

$$\mathbb{P}(B) = \frac{1}{6}$$

$$\mathbb{P}(A, B) = \frac{1}{36}$$

Yes: $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$



second die

	1	2	3	4	5	6
1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

first die

Handwritten annotations: An orange circle highlights the first row (1,*) and the first column (*,1). A blue diagonal line runs from (1,1) to (6,6). A red circle highlights the cell (1,6). Handwritten fractions: 1/6 next to the first row, 1/6 next to the first column, and 1/36 next to the cell (1,6).

Chain rules

Chain Rule

DEFINITION:

$$\mathbb{P}(A, B) = \mathbb{P}(A) \cdot \mathbb{P}(B|A)$$

$$\mathbb{P}(A, B, C) = \text{?}$$

Recall the definition of cond. prob.:

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

Chain Rule

DEFINITION:

$$\mathbb{P}(A, B) = \mathbb{P}(A) \cdot \mathbb{P}(B|A)$$

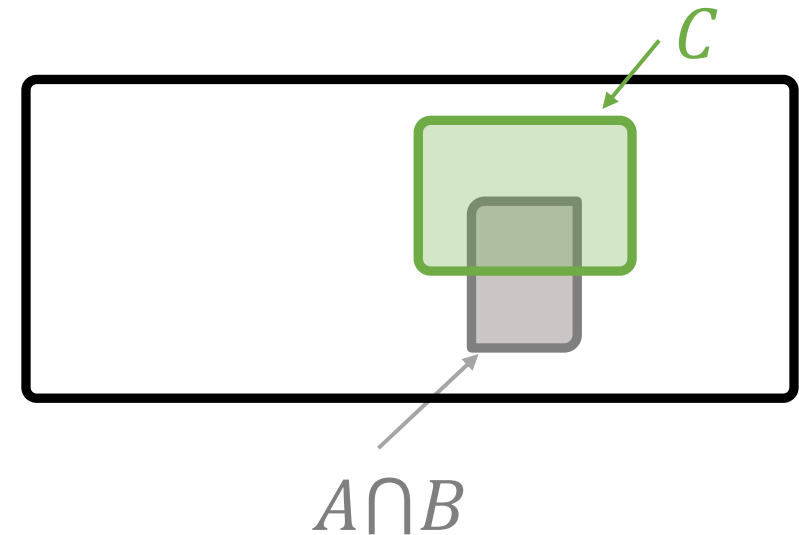
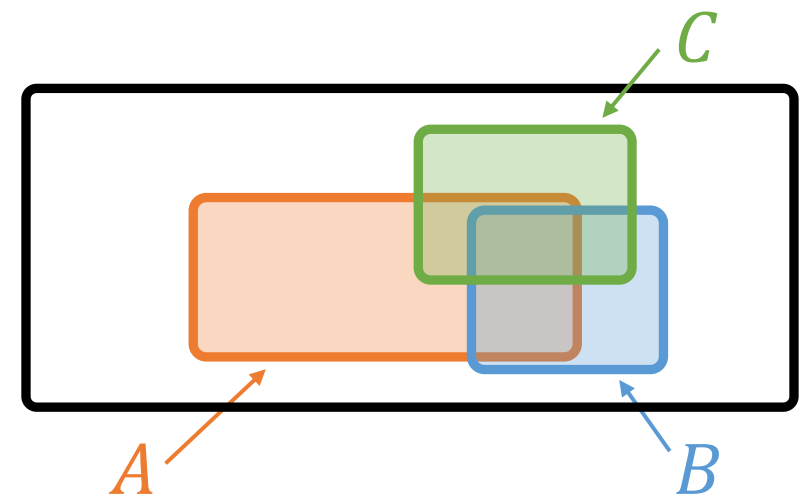
$$\mathbb{P}(\underbrace{A, B}_E, C) = \mathbb{P}(\underbrace{A, B}_E) \cdot \mathbb{P}(C|\underbrace{A, B}_E)$$

Just treat $A \cap B$ as a new event E

$$\mathbb{P}(A, B, C) = \mathbb{P}(A) \cdot \mathbb{P}(B|A) \cdot \mathbb{P}(C|A, B)$$

Recall the definition of cond. prob.:

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$



Bayes law

Bayes Law

DEFINITION

Conditional probability $\mathbb{P}(A|B)$ is the probability that event A occurred, given that event B occurred (defined if $\mathbb{P}(B) \neq 0$)

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

$$\mathbb{P}(B|A) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(A)}$$

$$\mathbb{P}(A, B) = \mathbb{P}(B|A) \cdot \mathbb{P}(A)$$

$$= \mathbb{P}(A|B) \cdot \mathbb{P}(B)$$

BAYES LAW

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(B|A) \cdot \mathbb{P}(A)}{\mathbb{P}(B)}$$

Bayes Law

DEFINITION

Conditional probability $\mathbb{P}(A|B)$ is the probability that event A occurred, given that event B occurred (defined if $\mathbb{P}(B) \neq 0$)

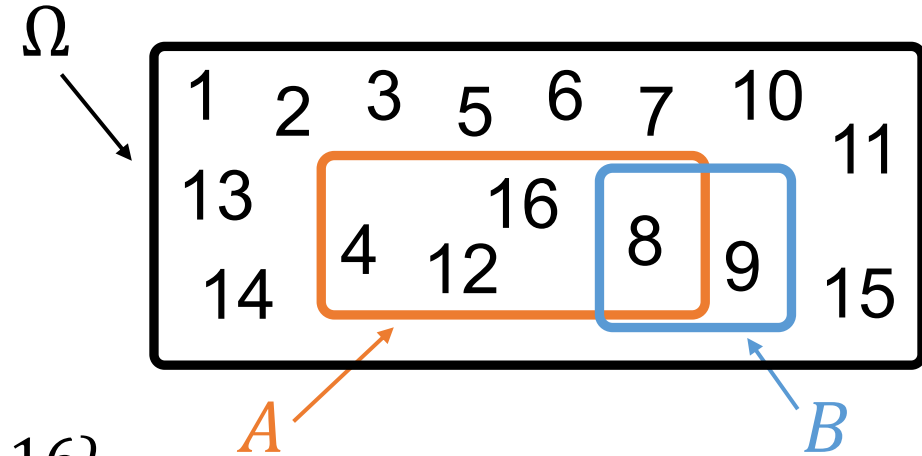
$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

$$\mathbb{P}(B|A) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(A)}$$

$$\begin{aligned}\mathbb{P}(A, B) &= \mathbb{P}(B|A) \cdot \mathbb{P}(A) \\ &= \mathbb{P}(A|B) \cdot \mathbb{P}(B)\end{aligned}$$

EXAMPLE 4:

$$\Omega = \{1, 2, \dots, 16\}$$



BAYES LAW

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(B|A) \cdot \mathbb{P}(A)}{\mathbb{P}(B)}$$

?

Bayes Law

DEFINITION

Conditional probability $\mathbb{P}(A|B)$ is the probability that event A occurred, given that event B occurred (defined if $\mathbb{P}(B) \neq 0$)

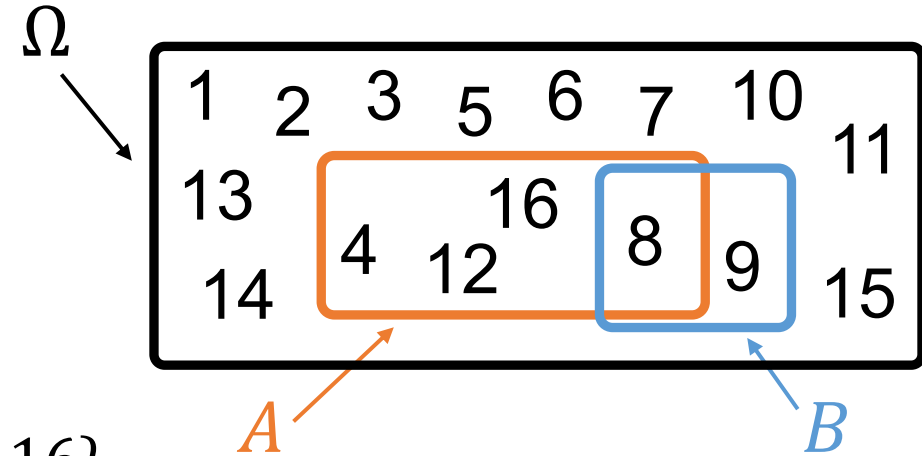
$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

$$\mathbb{P}(B|A) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(A)}$$

$$\begin{aligned}\mathbb{P}(A, B) &= \mathbb{P}(B|A) \cdot \mathbb{P}(A) = \frac{1}{4} \cdot \frac{1}{4} = \frac{1}{16} \\ &= \mathbb{P}(A|B) \cdot \mathbb{P}(B) = \frac{1}{2} \cdot \frac{1}{8} = \frac{1}{16}\end{aligned}$$

EXAMPLE 4:

$$\Omega = \{1, 2, \dots, 16\}$$



BAYES LAW

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(B|A) \cdot \mathbb{P}(A)}{\mathbb{P}(B)}$$

Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika (Z) in Florida is 10^{-5} (1 in 100k)
- accuracy of blood test is 99%
- A patient has a positive test.

What is the chance they have Zika?



		blood test positive (B)	
		B	$\bar{B}$
patient has Zika (Z)	Z	$\mathbb{P}(B Z) = 0.99$	$\mathbb{P}(\bar{B} Z) = 0.01$
	$\bar{Z}$	$\mathbb{P}(B \bar{Z}) = 0.01$	$\mathbb{P}(\bar{B} \bar{Z}) = 0.99$

Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika (Z) in Florida is 10^{-5} (1 in 100k) = **base rate** = $\mathbb{P}(Z)$

- accuracy of blood test is 99%

- A patient has a positive test.

What is the chance they have Zika?



patient has Z
Zika (Z)

blood test positive (B)	
B	$\bar{B}$
$\mathbb{P}(B Z) = 0.99$	$\mathbb{P}(\bar{B} Z) = 0.01$
$\mathbb{P}(B \bar{Z}) = 0.01$	$\mathbb{P}(\bar{B} \bar{Z}) = 0.99$

false negative

false positive

We don't want $\mathbb{P}(B|Z) = 0.99$, but rather $\mathbb{P}(Z|B)$

Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika (Z) in Florida is 10^{-5} (1 in 100k) = base rate = $\mathbb{P}(Z)$
- accuracy of blood test is 99%
- A patient has a positive test.
What is the chance they have Zika?

		blood test positive (B)	
		B	$\bar{B}$
patient has Zika (Z)	Z	$\mathbb{P}(B Z) = 0.99$	$\mathbb{P}(\bar{B} Z) = 0.01$
	$\bar{Z}$	$\mathbb{P}(B \bar{Z}) = 0.01$	$\mathbb{P}(\bar{B} \bar{Z}) = 0.99$

We don't want $\mathbb{P}(B|Z) = 0.99$, but rather $\mathbb{P}(Z|B)$

$$\mathbb{P}(Z|B) = \frac{\mathbb{P}(B, Z)}{\mathbb{P}(B)} = \frac{\mathbb{P}(B|Z) \cdot \mathbb{P}(Z)}{?}$$



Bayes Law ("Base rate fallacy problem")

EXAMPLE 5: Zika in Florida in 2016

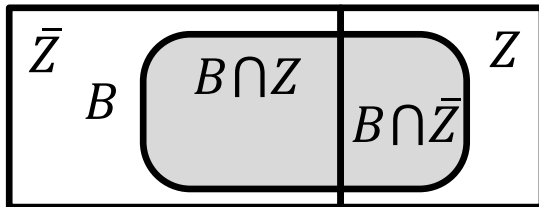
- prevalence of Zika (Z) in Florida is 10^{-5} (1 in 100k) = base rate = $\mathbb{P}(Z)$
- accuracy of blood test is 99%
- A patient has a positive test.
What is the chance they have Zika?

		blood test positive (B)	
		B	$\bar{B}$
patient has Zika (Z)	Z	$\mathbb{P}(B Z) = 0.99$	$\mathbb{P}(\bar{B} Z) = 0.01$
	$\bar{Z}$	$\mathbb{P}(B \bar{Z}) = 0.01$	$\mathbb{P}(\bar{B} \bar{Z}) = 0.99$

We don't want $\mathbb{P}(B|Z) = 0.99$, but rather $\mathbb{P}(Z|B)$

$$\mathbb{P}(Z|B) = \frac{\mathbb{P}(B, Z)}{\mathbb{P}(B)} = \frac{\mathbb{P}(B|Z) \cdot \mathbb{P}(Z)}{\mathbb{P}(B, Z) + \mathbb{P}(B, \bar{Z})}$$

total probability theorem



$\mathbb{P}(B|Z) \cdot \mathbb{P}(Z)$?



Bayes Law ("Base rate fallacy problem")

EXAMPLE 5: Zika in Florida in 2016

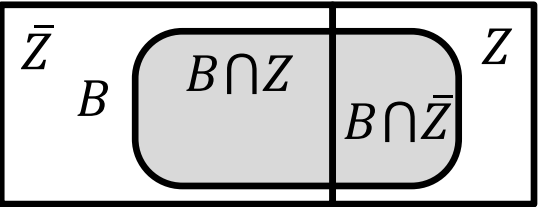
- prevalence of Zika (Z) in Florida is 10^{-5} (1 in 100k) = base rate = $\mathbb{P}(Z)$
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We don't want $\mathbb{P}(B|Z) = 0.99$, but rather $\mathbb{P}(Z|B)$

$$\mathbb{P}(Z|B) = \frac{\mathbb{P}(B, Z)}{\mathbb{P}(B)} = \frac{\mathbb{P}(B|Z) \cdot \mathbb{P}(Z)}{\mathbb{P}(B, Z) + \mathbb{P}(B, \bar{Z})} = \frac{0.99 \cdot 10^{-5}}{0.99 \cdot 10^{-5} + 0.01 \cdot (1 - 10^{-5})}$$

total probability theorem



$\mathbb{P}(B|Z) \cdot \mathbb{P}(Z)$

$\mathbb{P}(B|\bar{Z}) \cdot \mathbb{P}(\bar{Z})$

$$\approx \frac{10^{-5}}{10^{-5} + 0.01} \approx \frac{10^{-5}}{10^{-2}} = 10^{-3} = 0.1\%$$

Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika (Z) in Florida is 10^{-5} (1 in 100k) = base rate = $\mathbb{P}(Z)$
- accuracy of blood test is 99%
- Assume 10M people in FL. **?**
Complete the numbers!

		blood test positive (B)	
		B	$\bar{B}$
patient has Zika (Z)	Z	?	?
	$\bar{Z}$	?	?

$$\mathbb{P}(Z|B) \approx 0.1\%$$

Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika (Z) in Florida is 10^{-5} (1 in 100k) = **base rate** = $\mathbb{P}(Z)$
- accuracy of blood test is 99%
- Assume 10M people in FL.
Complete the numbers!

		blood test positive (B)		
		B	$\bar{B}$	
patient has Zika (Z)	Z			100
	$\bar{Z}$			9,999,900

$$\mathbb{P}(Z|B) \approx 0.1\%$$

Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika (Z) in Florida is 10^{-5} (1 in 100k) = base rate = $\mathbb{P}(Z)$
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- Assume 10M people in FL.
Complete the numbers!

		blood test positive (B)		
		B	$\bar{B}$	
patient has Zika (Z)	Z	99	1	100
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Complete the numbers!

		blood test positive (B)		
		B	$\bar{B}$	
patient has Zika (Z)	Z	99	1	100
	$\bar{Z}$	99,999	9,899,901	9,999,900

$$\mathbb{P}(Z|B) \approx 0.1\%$$

Bayes Law ("Base rate fallacy problem")



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- Assume 10M people in FL. Complete the numbers!

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		B	$\bar{B}$	
patient has Zika (Z)	Z	99	1	100
	$\bar{Z}$	99,999	9,899,901	9,999,900
		100,098	9,899,902	

$\mathbb{P}(Z|B) \approx 0.1\%$
probability that:
a positive test is correct

Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika (Z) in Florida is 10^{-5} (1 in 100k) = **base rate** = $\mathbb{P}(Z)$

- accuracy of blood test is 99%

- Assume 10M people in FL.
Complete the numbers!

$$\mathbb{P}(B|Z) = 99\%$$
$$\mathbb{P}(\bar{B}|\bar{Z}) = 99\%$$

patient has Z
Zika (Z) $\bar{Z}$

blood test positive (B)		
B	$\bar{B}$	
99	1	100
99,999	9,899,901	9,999,900
100,098	9,899,902	

$$\mathbb{P}(B|Z) \cdot \mathbb{P}(Z) + \mathbb{P}(\bar{B}|\bar{Z}) \cdot \mathbb{P}(\bar{Z}) = 99\%$$

probability that a random
test is correct

$\neq$

$$\mathbb{P}(Z|B) \approx 0.1\%$$

probability that:
a positive test is correct



Bayes Law ("Base rate fallacy problem")

EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika (Z) in Florida is 10^{-5} (1 in 100k) = **base rate** = $\mathbb{P}(Z)$
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- Assume 10M people in FL. Complete the numbers!

		blood test positive (B)		
		B	$\bar{B}$	
patient has Zika (Z)	Z	99	1	100
	$\bar{Z}$	99,999	9,899,901	9,999,900
		100,098	9,899,902	

$\mathbb{P}(B|Z) \cdot \mathbb{P}(Z) + \mathbb{P}(\bar{B}|\bar{Z}) \cdot \mathbb{P}(\bar{Z}) = 99\%$
probability that a random
test is correct

$\neq$

$\mathbb{P}(Z|B) \approx 0.1\%$
probability that:
a positive test is correct

$\bigcirc \rightarrow \mathbb{P}(Z) = 10^{-5}$



Bayes Law ("Base rate fallacy problem")

EXAMPLE 5: Zika in Florida in 2016

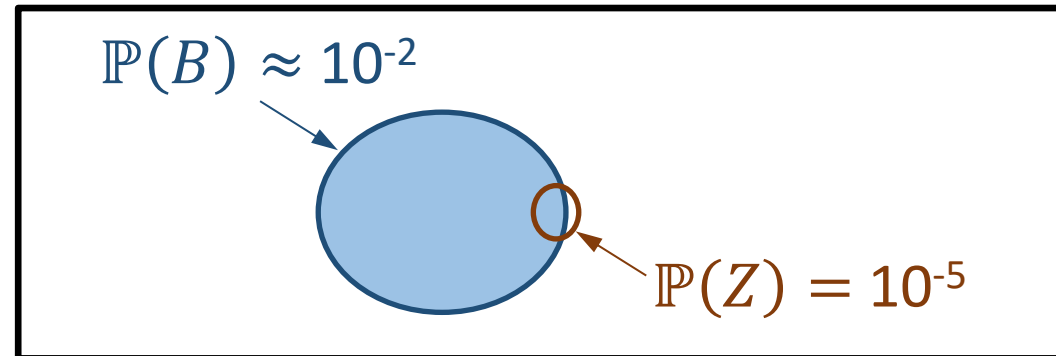
- prevalence of Zika (Z) in Florida is 10^{-5} (1 in 100k) = **base rate** = $\mathbb{P}(Z)$
- accuracy of blood test is 99%
- Assume 10M people in FL. Complete the numbers!

		blood test positive (B)		
		B	$\bar{B}$	
patient has Zika (Z)	Z	99	1	100
	$\bar{Z}$	99,999	9,899,901	9,999,900
		100,098	9,899,902	

$\mathbb{P}(B|Z) \cdot \mathbb{P}(Z) + \mathbb{P}(\bar{B}|\bar{Z}) \cdot \mathbb{P}(\bar{Z}) = 99\%$
probability that a random test is correct

$\neq$

$\mathbb{P}(Z|B) \approx 0.1\%$
probability that:
a positive test is correct





Bayes Law ("Base rate fallacy problem")

EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika (Z) in Florida is 10^{-5} (1 in 100k) = **base rate** = $\mathbb{P}(Z)$
- accuracy of blood test is 99%

- Assume 10M people in FL.
Complete the numbers

$$\mathbb{P}(B|Z) = 99\%$$
$$\mathbb{P}(\bar{B}|\bar{Z}) = 99\%$$

patient has Z
Zika (Z) $\bar{Z}$

blood test positive (B)		
B	$\bar{B}$	
99	1	100
99,999	9,899,901	9,999,900
100,098	9,899,902	

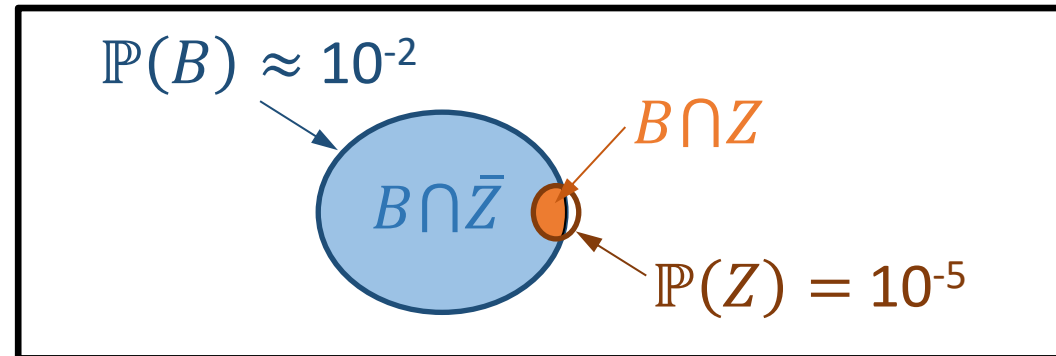
$$\mathbb{P}(B|Z) \cdot \mathbb{P}(Z) + \mathbb{P}(\bar{B}|\bar{Z}) \cdot \mathbb{P}(\bar{Z}) = 99\%$$

probability that a random test is correct

$\neq$

$$\mathbb{P}(Z|B) = 0.1\%$$

probability that:
a positive test is correct

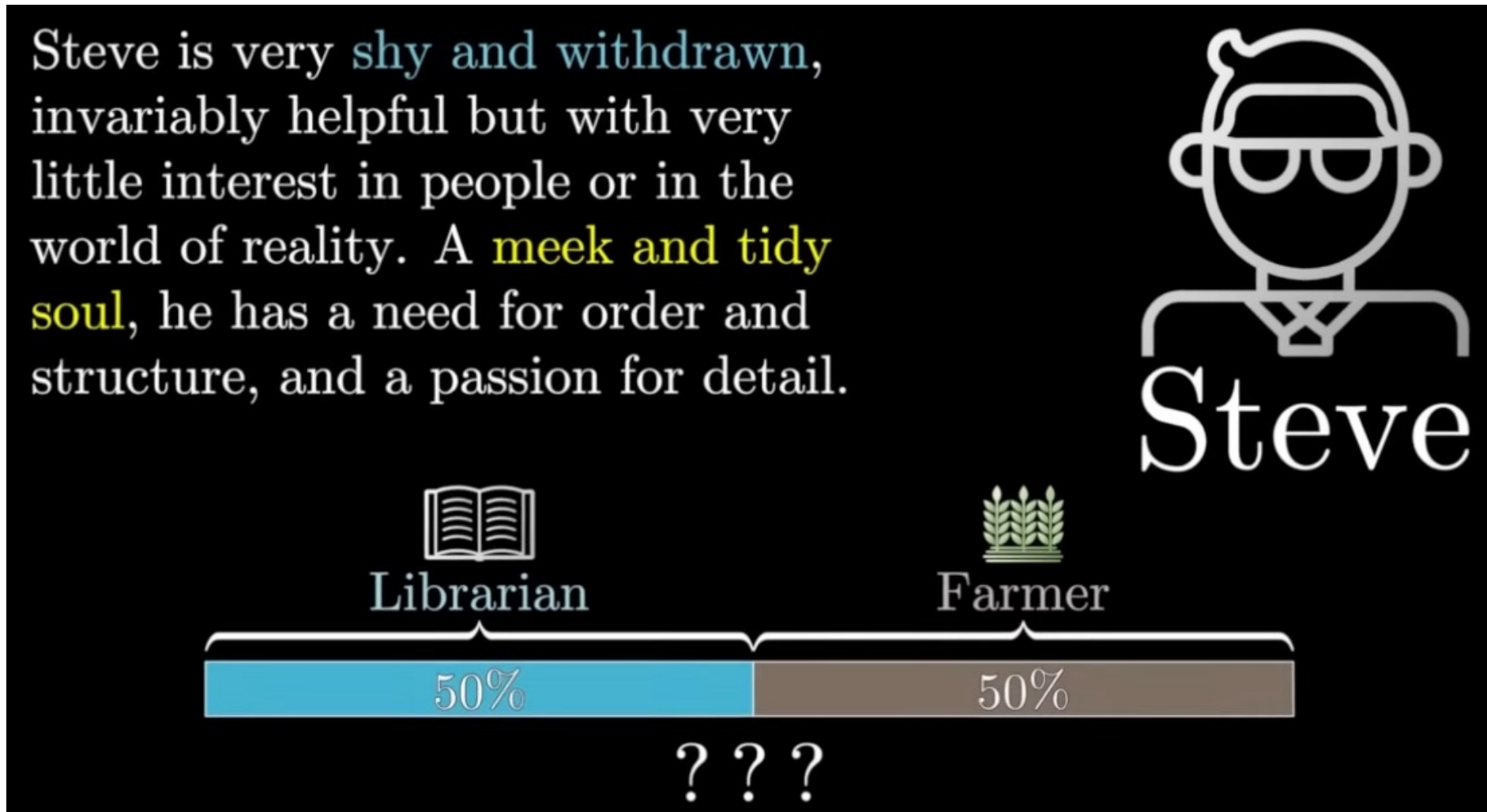


An unrelated (?) question:

Steve is very shy and withdrawn, invariably helpful but with very little interest in people or in the world of reality. A meek and tidy soul, he has a need for order and structure, and a passion for detail.



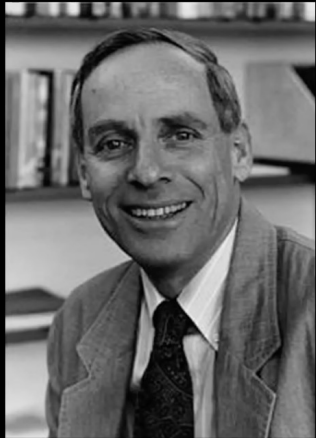
An unrelated (?) question:



Bayes theorem illustrated by 3Blue1Brown

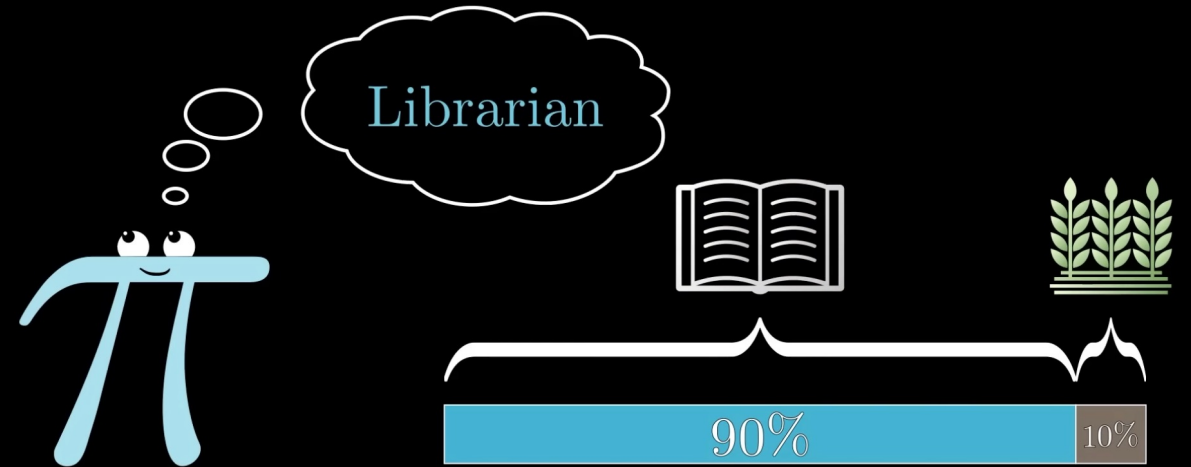


Daniel
Kahneman

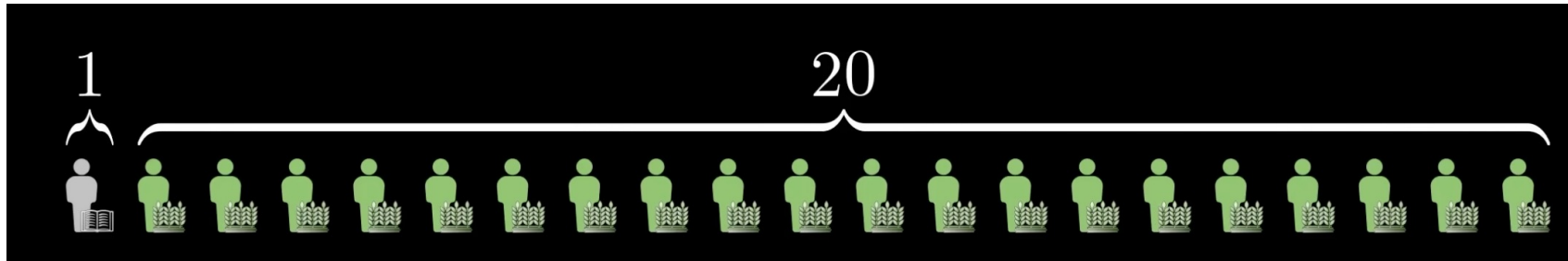


Amos
Tversky

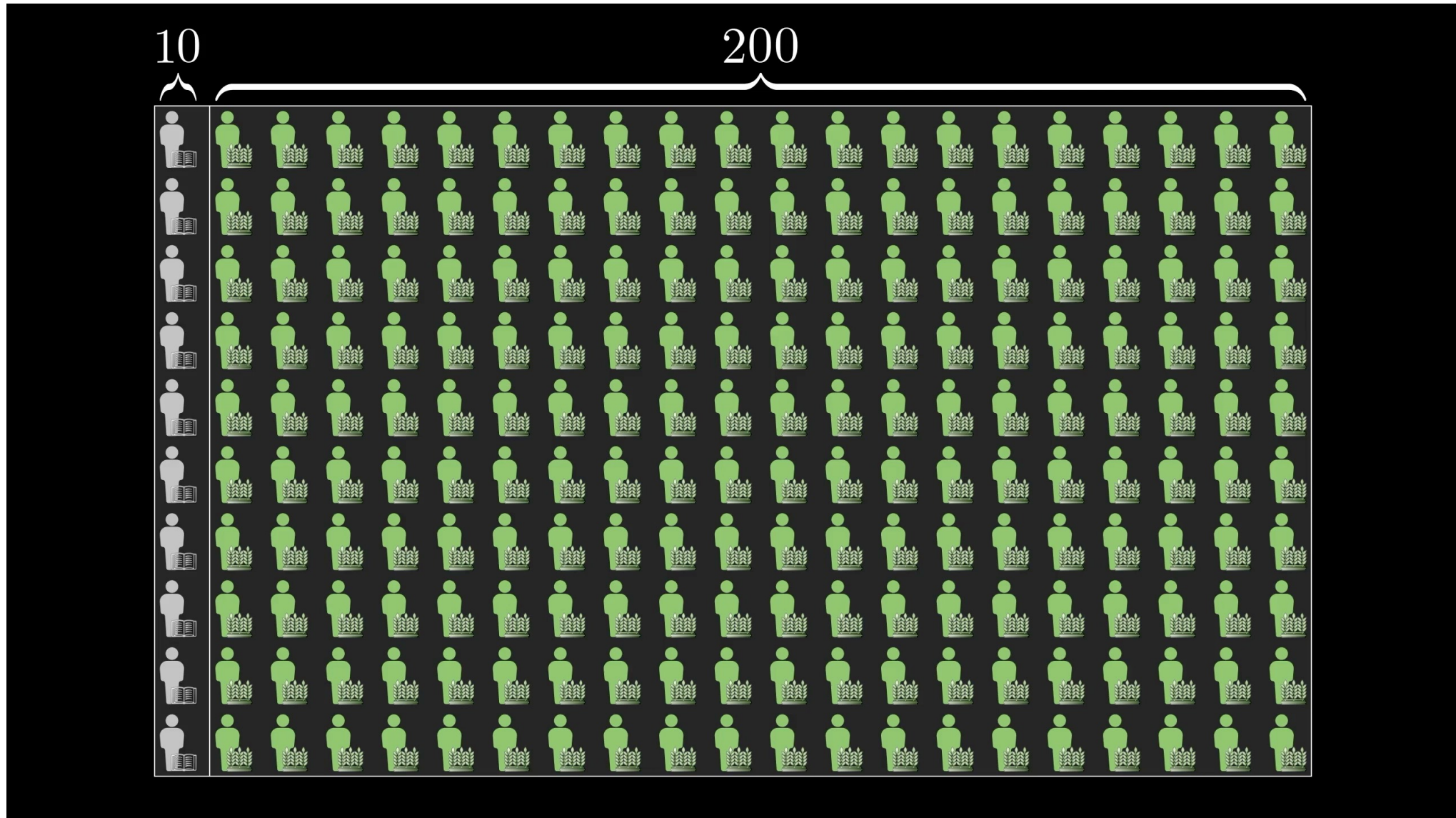
Steve is very shy and withdrawn, invariably helpful but with very little interest in people or in the world of reality. A meek and tidy soul, he has a need for order and structure, and a passion for detail.



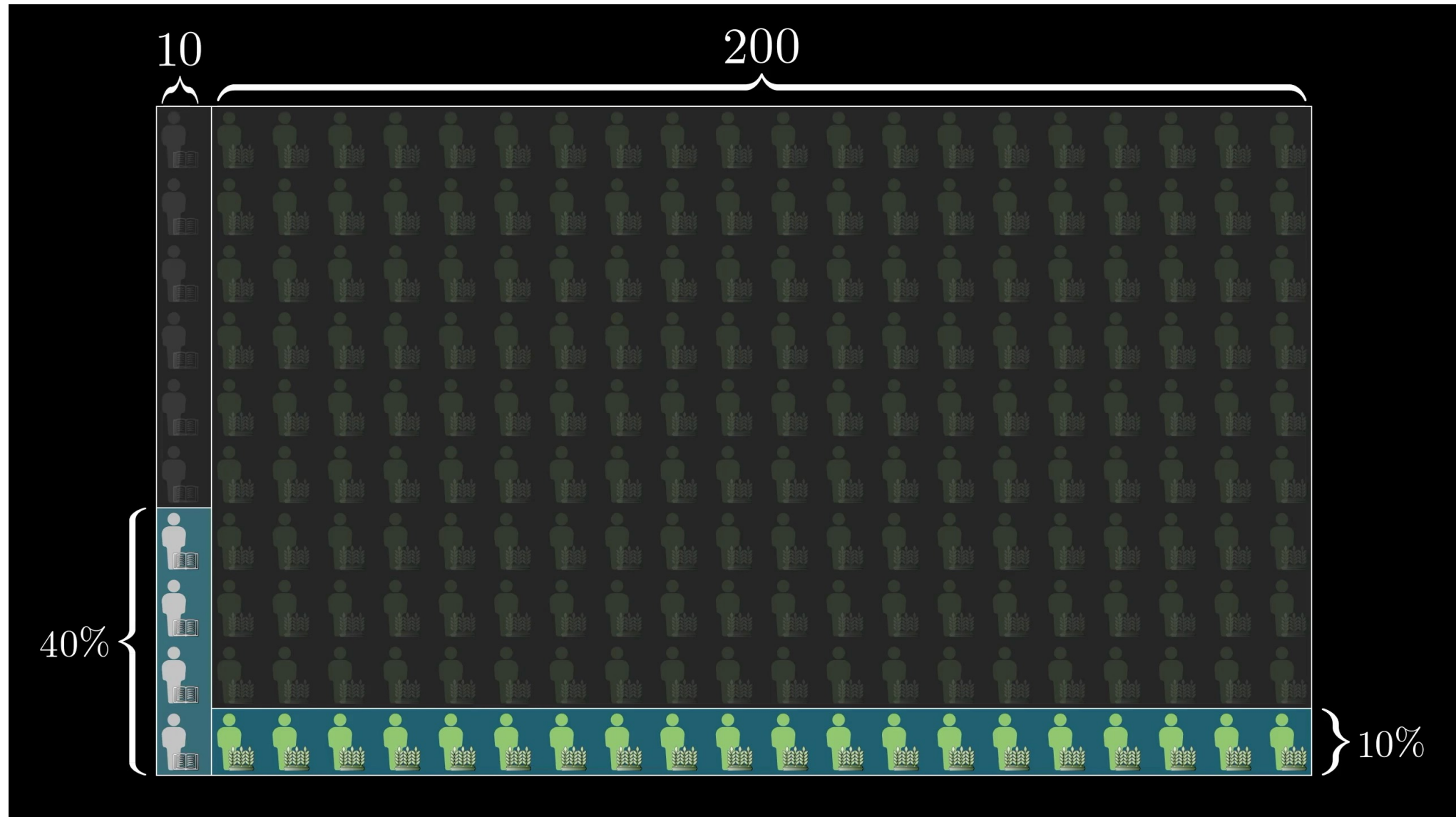
Bayes theorem illustrated by 3Blue1Brown



Bayes theorem illustrated by 3Blue1Brown

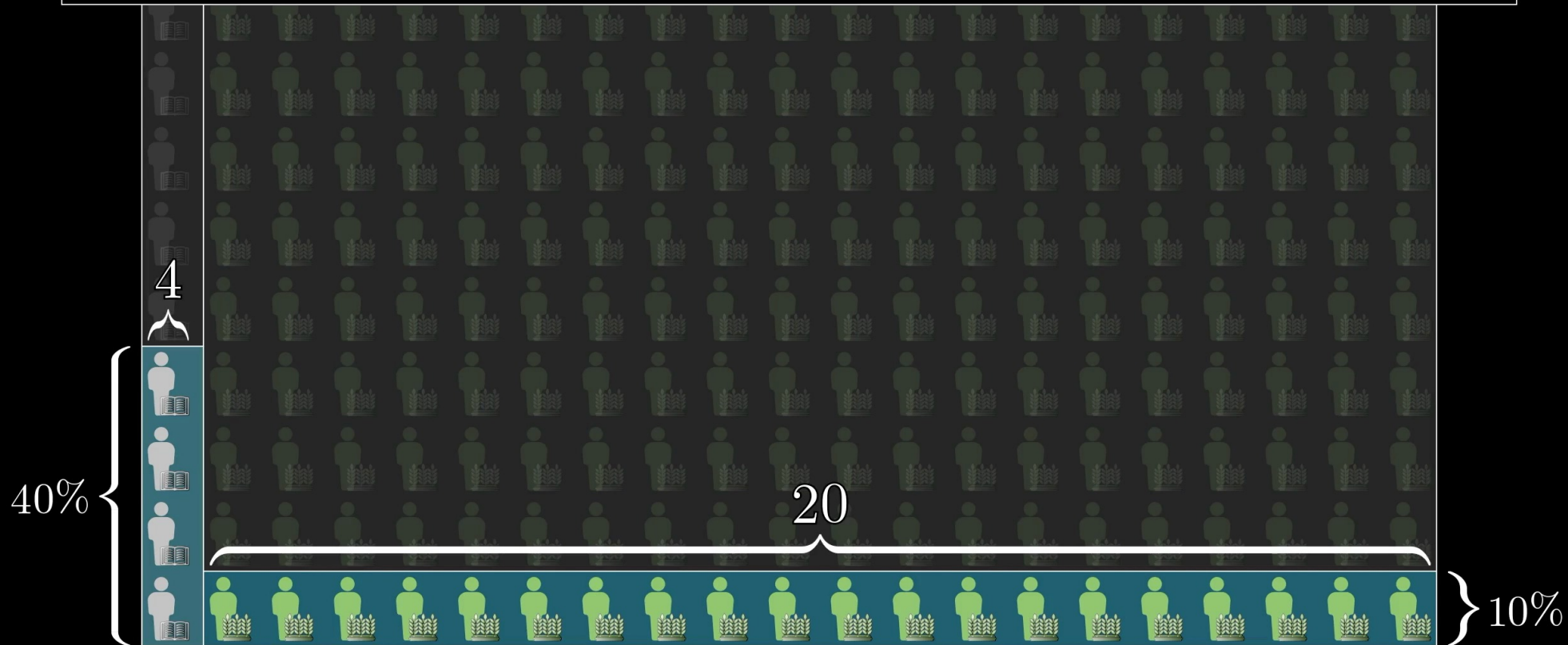


Bayes theorem illustrated by 3Blue1Brown



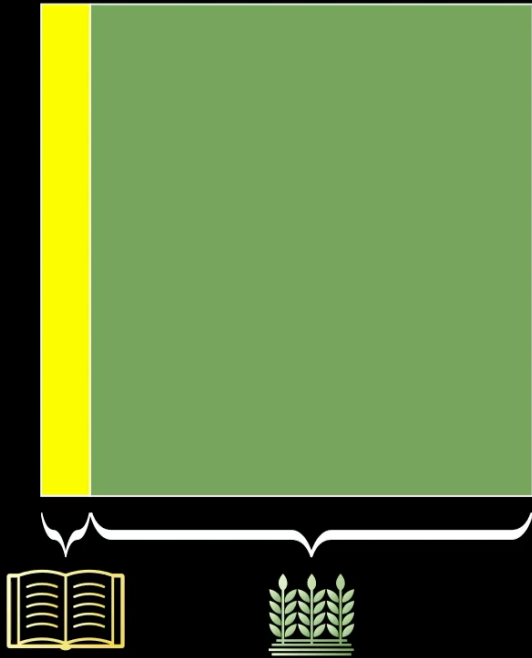
Bayes theorem illustrated by 3Blue1Brown

$$P(\text{Librarian given description}) = \frac{4}{4 + 20} \approx 16.7\%$$

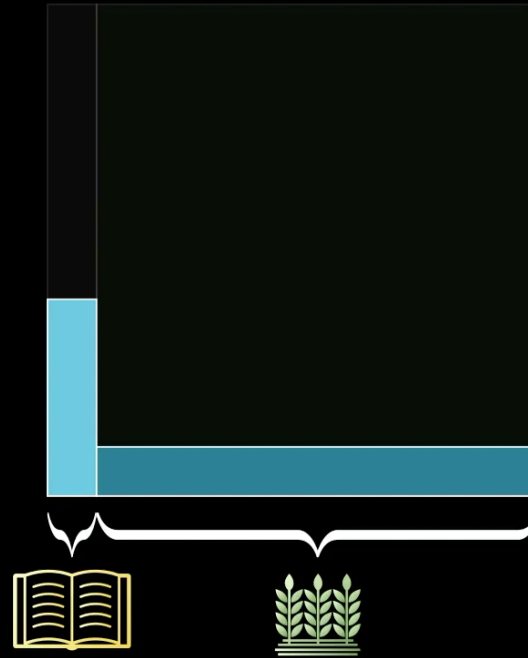


Bayes theorem illustrated by 3Blue1Brown

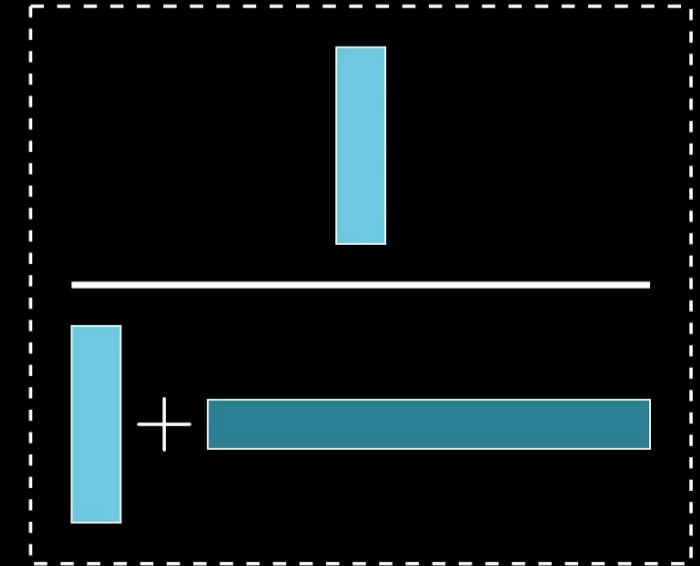
All possibilities



All possibilities
fitting the evidence



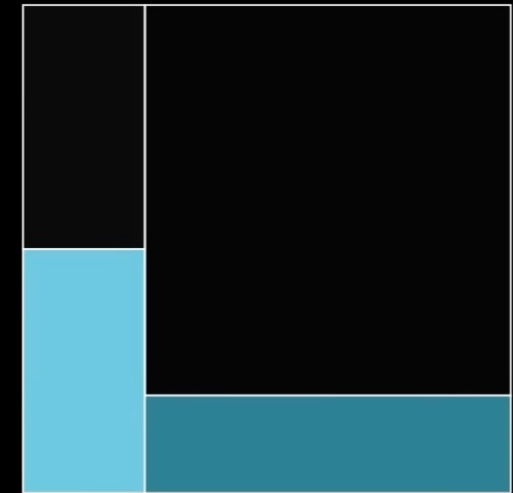
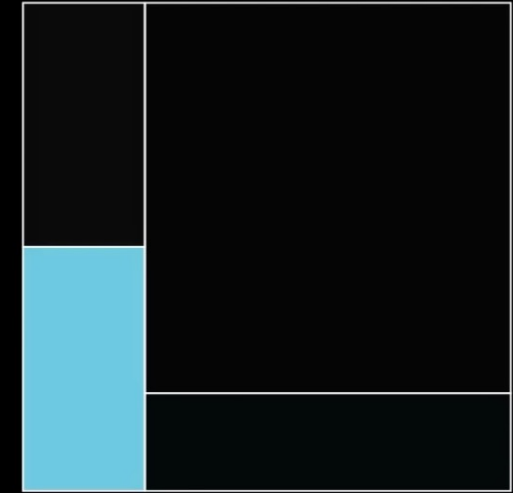
P (Librarian given
the evidence)



Bayes theorem illustrated by 3Blue1Brown

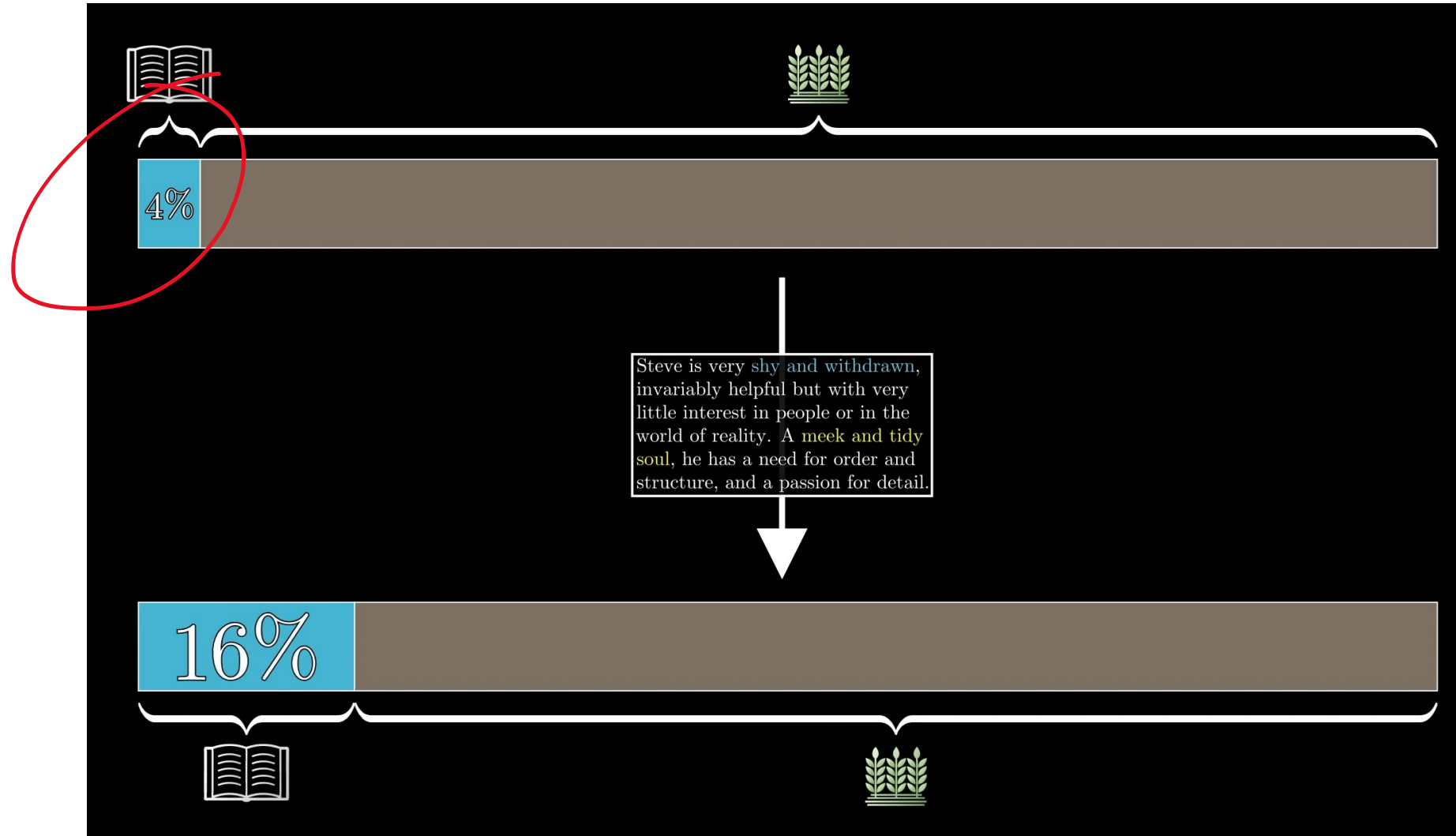
Bayes' theorem

$$P(H|E) = \frac{P(H)P(E|H)}{P(E)} = \frac{\text{Area of top-left rectangle}}{\text{Area of bottom-left rectangle}}$$



Bayes theorem illustrated by 3Blue1Brown

New evidence updates prior beliefs! Evidence does not exist in a vacuum



Part 1: Information Theory Basics

L03: Basics of Probability (2/2)

[Random experiment, independence, conditional probability, chain rule, Bayes' theorem, random variables]

Wolfgang Gatterbauer

cs7840 Foundations and Applications of Information Theory (fa25)

<https://northeastern-datalab.github.io/cs7840/fa25/>

9/15/2025

Pre-class conversations

- Last class recapitulation
- Suggestions on small class projects = "scribes" (e.g. parallelizing compression)
- New class arrivals
- Questions

- Today:
 - The basics of probability theory
 - Start of information theory basics

Random variable, expectation

Random variable: pmf $p_X(x)$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

Two types:

1. numerical

e.g. $X(\omega)$ lottery win

2. indicator

e.g. $X(\omega) = \begin{cases} 1 & \text{if win} \\ 0 & \text{if not} \end{cases}$

X is called a "random variable" (RV) because it depends on the outcome of a random experiment. But the mapping $X: \Omega \rightarrow \mathbb{R}$ is deterministic.

The underlying probability measure $\mathbb{P}: \Sigma \rightarrow [0, 1]$ induces a pmf p_X (probability mass function) over the range of the RV $X: p_X: \mathbb{R} \rightarrow [0, 1]$

$$p_X(x) = \mathbb{P}(\{X = x\})$$

also written as $p(x) = \mathbb{P}(X = x)$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let X be the sum of the rolls.

Then X is a RV.

What is the pmf?



	Ω	second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

Random variable: pmf $p_X(x)$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

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2. indicator

e.g. $X(\omega) = \begin{cases} 1 & \text{if win} \\ 0 & \text{if not} \end{cases}$

X is called a "random variable" (RV) because it depends on the outcome of a random experiment. But the mapping $X: \Omega \rightarrow \mathbb{R}$ is deterministic.

The underlying probability measure $\mathbb{P}: \Sigma \rightarrow [0, 1]$ induces a pmf p_X (probability mass function) over the range of the RV $X: p_X: \mathbb{R} \rightarrow [0, 1]$

$$p_X(x) = \mathbb{P}(\{X = x\})$$

also written as $p(x) = \mathbb{P}(X = x)$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let X be the sum of the rolls.

Then X is a RV.

What is the pmf?



		Ω second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)
		X					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
	6	7	8	9	10	11	12

Random variable: pmf $p_X(x)$

Random variable (RV)

$$X: \Omega \rightarrow \mathbb{R}$$

Two types:

1. numerical

e.g. $X(\omega)$ lottery win

2. indicator

$$\text{e.g. } X(\omega) = \begin{cases} 1 & \text{if win} \\ 0 & \text{if not} \end{cases}$$

X is called a "random variable" (RV) because it depends on the outcome of a random experiment. But the mapping $X: \Omega \rightarrow \mathbb{R}$ is deterministic.

The underlying probability measure $\mathbb{P}: \Sigma \rightarrow [0, 1]$ induces a pmf p_X (probability mass function) over the range of the RV $X: p_X: \mathbb{R} \rightarrow [0, 1]$

$$p_X(x) = \mathbb{P}(\{X = x\})$$

also written as $p(x) = \mathbb{P}(X = x)$

EXAMPLE 3 (CONT.):

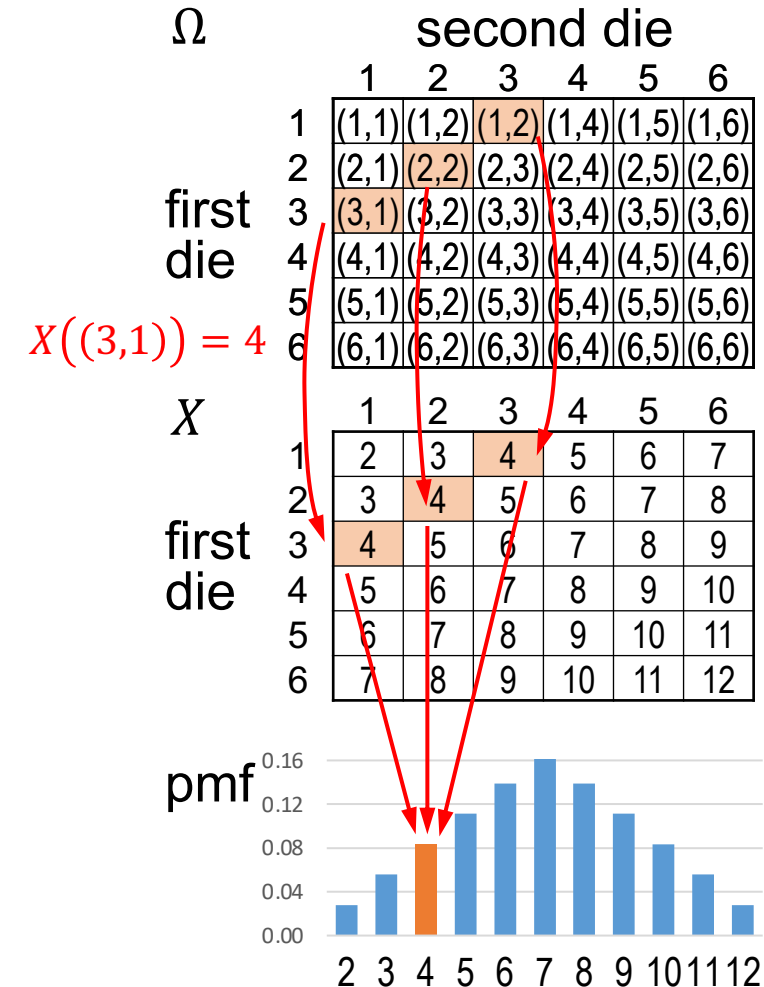
roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let X be the sum of the rolls.

Then X is a RV.

What is the pmf?



Random variable: expectation $\mathbb{E}$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

Expectation: a weighted average
(in proportion to the probabilities)
of the possible values of X

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

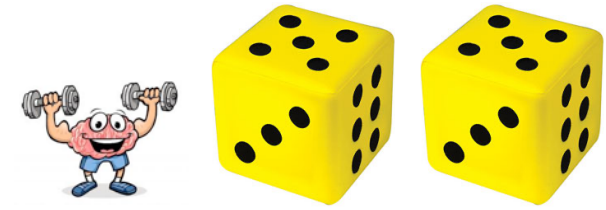
$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$

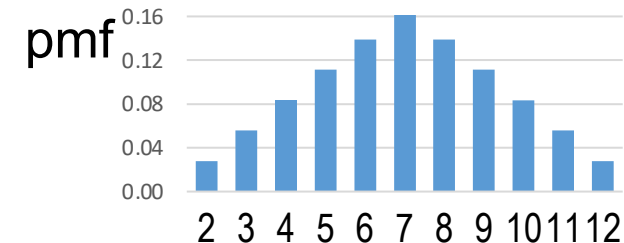
Let X be the sum of the rolls.



What is $\mathbb{E}[X]$?

		Ω second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		X					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
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Random variable: expectation $\mathbb{E}$

Random variable (RV)

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Expectation: a weighted average
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$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let X be the sum of the rolls.



What is $\mathbb{E}[X]$?

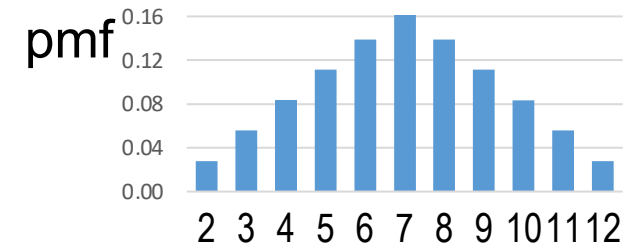
Variant 1:

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

?

		Ω second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		X					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
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Random variable: expectation $\mathbb{E}$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

Expectation: a weighted average
(in proportion to the probabilities)
of the possible values of X

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

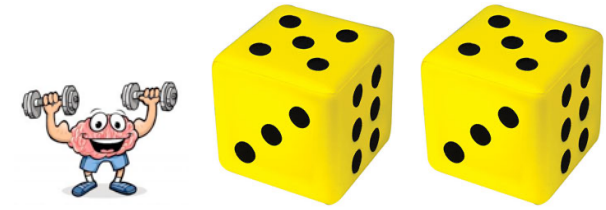
$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$

Let X be the sum of the rolls.



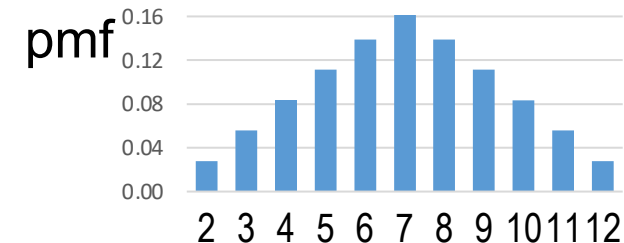
What is $\mathbb{E}[X]$?

Variant 1:

$$\begin{aligned} \mathbb{E}[X] &= \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\}) \\ &= \frac{1}{36} \cdot \sum_{\omega \in \Omega} X(\omega) \\ &= \frac{1}{36} \cdot (\text{"sum of table"}) \\ &= 7 \end{aligned}$$

		Ω second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		X					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
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Random variable: expectation $\mathbb{E}$

Random variable (RV)

$$X: \Omega \rightarrow \mathbb{R}$$

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$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

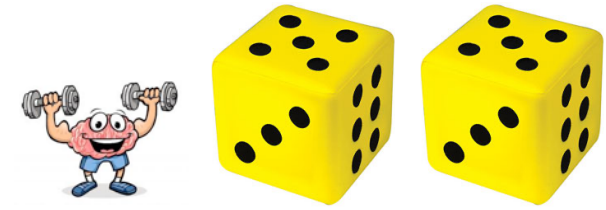
$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let X be the sum of the rolls.



What is $\mathbb{E}[X]$?

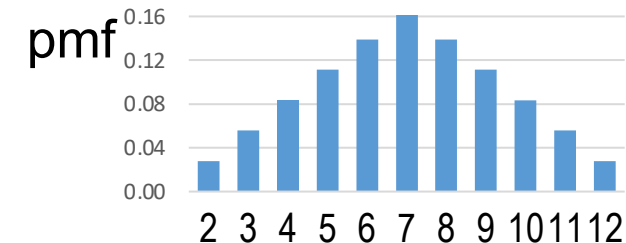
Variant 2:

$$\mathbb{E}[X] = \sum_x x \cdot \mathbb{P}(X = x)$$

?

		Ω second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		X					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
	6	7	8	9	10	11	12



Random variable: expectation $\mathbb{E}$

Random variable (RV)

$$X: \Omega \rightarrow \mathbb{R}$$

Expectation: a weighted average
(in proportion to the probabilities)
of the possible values of X

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

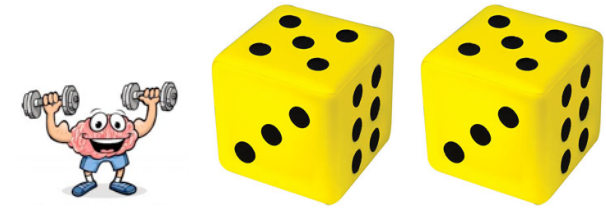
$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let X be the sum of the rolls.



What is $\mathbb{E}[X]$?

Variant 2:

$$\mathbb{E}[X] = \sum_x x \cdot \mathbb{P}(X = x)$$

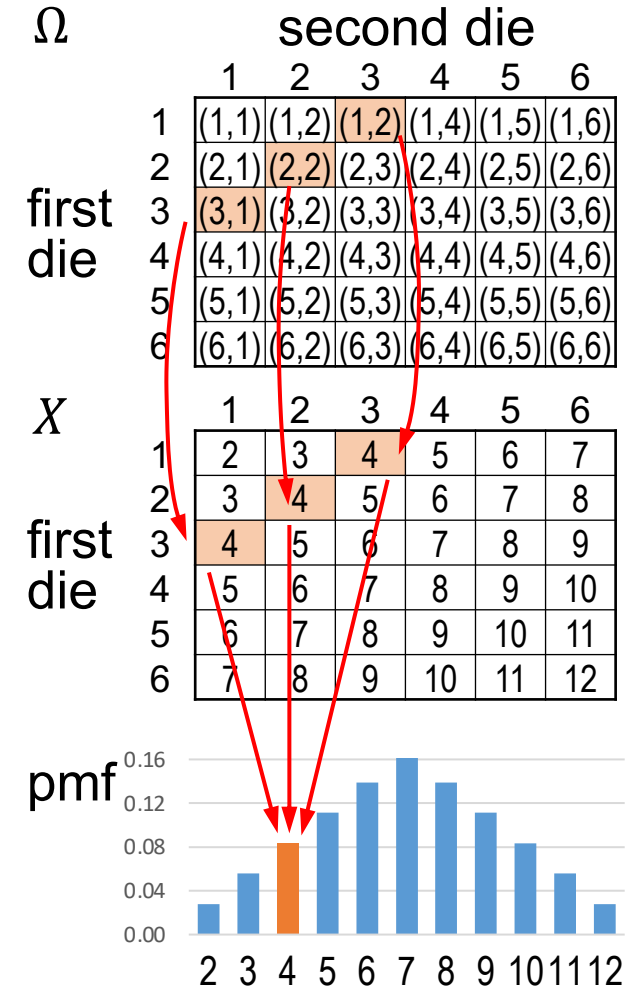
$$X = 2 \quad p_X(2) = \frac{1}{36}$$

$$X = 3 \quad p_X(3) = \frac{2}{36}$$

...

$$X = 12 \quad p_X(12) = \frac{1}{36}$$

$$\mathbb{E}[X] = 2 \cdot \frac{1}{36} + 3 \cdot \frac{2}{36} + \dots + 12 \cdot \frac{1}{36} = 7$$



Random variable: expectation $\mathbb{E}$

Random variable (RV)

$$X: \Omega \rightarrow \mathbb{R}$$

Expectation: a weighted average
(in proportion to the probabilities)
of the possible values of X

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let X be the sum of the rolls.

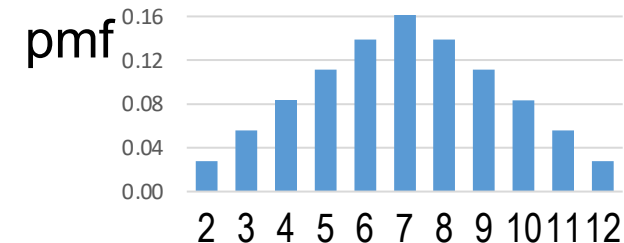


What is $\mathbb{E}[X]$?

Variant 3: *RVs for the outcome of the first and second die rolls*

$$\mathbb{E}[X] = \mathbb{E}[X_1 + X_2] = ?$$

		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)
		X					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
	6	7	8	9	10	11	12



Random variable: expectation $\mathbb{E}$

Random variable (RV)

$$X: \Omega \rightarrow \mathbb{R}$$

Expectation: a weighted average
(in proportion to the probabilities)
of the possible values of X

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let X be the sum of the rolls.



What is $\mathbb{E}[X]$?

Variant 3: *RVs for the outcome of the first and second die rolls*

$$\mathbb{E}[X] = \mathbb{E}[X_1 + X_2] = \mathbb{E}[X_1] + \mathbb{E}[X_2]$$

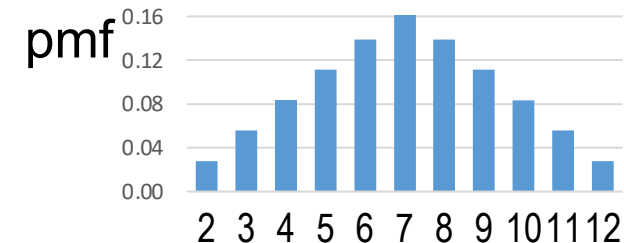
linearity of expectation!

$$\mathbb{E}[X_1] = \mathbb{E}[X_2] = 3.5$$

$$\mathbb{E}[X] = 3.5 + 3.5 = 7$$

		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
	6	7	8	9	10	11	12



Random variable: expectation $\mathbb{E}$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

Expectation: a weighted average
(in proportion to the probabilities)
of the possible values of X

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$

Let Y be the product of the rolls.



What is $\mathbb{E}[Y]$?

		Ω second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)
		Y					
		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
	2	2	4	6	8	10	12
	3	3	6	9	12	15	18
	4	4	8	12	16	20	24
	5	5	10	15	20	25	30
	6	6	12	18	24	30	36

Random variable: expectation $\mathbb{E}$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

Expectation: a weighted average
(in proportion to the probabilities)
of the possible values of X

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

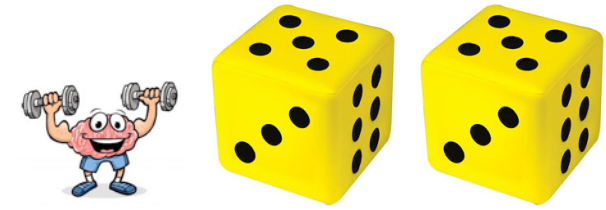
$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$

Let Y be the product of the rolls.

What is $\mathbb{E}[Y]$?

$$\mathbb{E}[Y] = \mathbb{E}[X_1 \cdot X_2] = \mathbb{E}[X_1] \cdot \mathbb{E}[X_2]$$

?



		Ω second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)
		Y					
		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
	2	2	4	6	8	10	12
	3	3	6	9	12	15	18
	4	4	8	12	16	20	24
	5	5	10	15	20	25	30
	6	6	12	18	24	30	36

Random variable: expectation $\mathbb{E}$

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Expectation: a weighted average
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of the possible values of X

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$

Let Y be the product of the rolls.



What is $\mathbb{E}[Y]$?

$$\mathbb{E}[Y] = \mathbb{E}[X_1 \cdot X_2] = \mathbb{E}[X_1] \cdot \mathbb{E}[X_2]$$

because the $X_1 \perp X_2$

$$\mathbb{E}[X_1] = \mathbb{E}[X_2] = 3.5$$

$$\mathbb{E}[Y] = 3.5 \cdot 3.5 = 12.25$$

		Ω second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)
		Y					
		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
	2	2	4	6	8	10	12
	3	3	6	9	12	15	18
	4	4	8	12	16	20	24
	5	5	10	15	20	25	30
	6	6	12	18	24	30	36

Random variable: expectation $\mathbb{E}$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

Expectation: a weighted average
(in proportion to the probabilities)
of the possible values of X

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

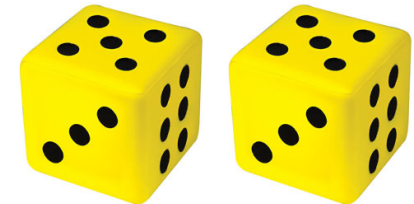
roll two fair dice with 6 sides

$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$

Let Y be the product of the rolls.

Let X be the sum of the rolls.

What is $\mathbb{E}[X + Y]$?



		Ω second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		Y					
		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
	2	2	4	6	8	10	12
	3	3	6	9	12	15	18
	4	4	8	12	16	20	24
	5	5	10	15	20	25	30
	6	6	12	18	24	30	36

		X					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
	6	7	8	9	10	11	12

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EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

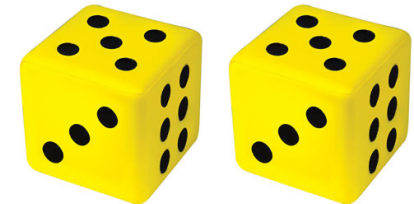
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Let Y be the product of the rolls.

Let X be the sum of the rolls.

What is $\mathbb{E}[X + Y]$?

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y] \quad ?$$



		Ω second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		Y					
		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
	2	2	4	6	8	10	12
	3	3	6	9	12	15	18
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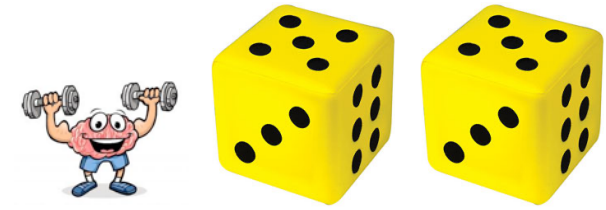
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$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y]$$

X and Y are clearly dependent ☹️



	Ω	second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

	Y	1	2	3	4	5	6
		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
	2	2	4	6	8	10	12
	3	3	6	9	12	15	18
	4	4	8	12	16	20	24
	5	5	10	15	20	25	30
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	X	1	2	3	4	5	6
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
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roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let Y be the product of the rolls.

Let X be the sum of the rolls.

What is $\mathbb{E}[X + Y]$?

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y]$$

X and Y are clearly dependent ☹️

But linearity of expectation still holds
even if the RVs are dependent 😊

$$\mathbb{E}[X] = 7 + 12.25 = 19.25$$



Ω		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,2)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

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		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
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X		1	2	3	4	5	6
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
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$\mathbb{E}[X + Y]$ vs. $\mathbb{E}[X \cdot Y]$



$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y] \text{ (linearity of expectation)}$$

holds even if X and Y are not independent

PROOF

$$\mathbb{E}[X + Y] = \sum_x \sum_y (x + y) \cdot p(x, y)$$

?

$$= \mathbb{E}[X] + \mathbb{E}[Y]$$

$$\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y], \text{ only if } X \perp Y \text{ (independent)}$$

PROOF

$$\mathbb{E}[X \cdot Y] = \sum_x \sum_y x \cdot y \cdot p(x, y)$$

?

$$= \mathbb{E}[X] \cdot \mathbb{E}[Y]$$

$\mathbb{E}[X + Y]$ vs. $\mathbb{E}[X \cdot Y]$

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y] \text{ (linearity of expectation)}$$

holds even if X and Y are not independent

PROOF

$$\begin{aligned}\mathbb{E}[X + Y] &= \sum_x \sum_y (x + y) \cdot p(x, y) \\ &= \sum_x \sum_y x \cdot p(x, y) + \sum_x \sum_y y \cdot p(x, y) \\ &= \sum_x x \cdot \sum_y p(x, y) + \sum_y y \cdot \sum_x p(x, y) \\ &= \sum_x x \cdot p(x) + \sum_y y \cdot p(y) \\ &= \mathbb{E}[X] + \mathbb{E}[Y]\end{aligned}$$

$$\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y], \text{ only if } X \perp Y \text{ (independent)}$$

PROOF

$$\begin{aligned}\mathbb{E}[X \cdot Y] &= \sum_x \sum_y x \cdot y \cdot p(x, y) \\ &= \sum_x \sum_y x \cdot y \cdot p(x) \cdot p(y) \\ &= \left(\sum_x x \cdot p(x) \right) \cdot \left(\sum_y y \cdot p(y) \right) \\ &= \mathbb{E}[X] \cdot \mathbb{E}[Y]\end{aligned}$$

because $X \perp Y$

distributive law of multiplication over addition (each cross-term appears exactly once on both sides)

Linearity of Expectation



EXAMPLE 6 (CONT.):

- The digits 1,2,3, and 4 are randomly arranged to form two two-digit numbers, $\overline{AB}$ and $\overline{CD}$. For example, we could have $\overline{AB} = 42$ and $\overline{CD} = 13$.
- What is the expected value of $\overline{AB} \cdot \overline{CD}$?

?

Linearity of Expectation



EXAMPLE 6 (CONT.):

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- What is the expected value of $\overline{AB} \cdot \overline{CD}$?
- What about $\mathbb{E}[\overline{AB}] \cdot \mathbb{E}[\overline{CD}]$?

?

Linearity of Expectation



EXAMPLE 6 (CONT.):

- The digits 1,2,3, and 4 are randomly arranged to form two two-digit numbers, $\overline{AB}$ and $\overline{CD}$. For example, we could have $\overline{AB} = 42$ and $\overline{CD} = 13$.
- What is the expected value of $\overline{AB} \cdot \overline{CD}$?
- ~~What about $\mathbb{E}[\overline{AB}] \cdot \mathbb{E}[\overline{CD}]$?~~
- **No!** $\mathbb{E}[xy] = \mathbb{E}[x] \cdot \mathbb{E}[y]$ only holds when the RVs are independent.

Clearly, $\overline{AB}$ and $\overline{CD}$ are not independent since each digit can only be used once (e.g., if $\overline{AB} = 42$ then we would know that $\overline{CD}$ can only be 13 or 31).

Linearity of Expectation

EXAMPLE 6 (CONT.):

- The digits 1,2,3, and 4 are randomly arranged to form two two-digit numbers, $\overline{AB}$ and $\overline{CD}$. For example, we could have $\overline{AB} = 42$ and $\overline{CD} = 13$.
- What is the expected value of $\overline{AB} \cdot \overline{CD}$?
- ~~What about $\mathbb{E}[\overline{AB}] \cdot \mathbb{E}[\overline{CD}]$?~~
- Can we instead get some kind of sum?



Linearity of Expectation

EXAMPLE 6 (CONT.):

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- What is the expected value of $\overline{AB} \cdot \overline{CD}$?

- ~~What about $\mathbb{E}[\overline{AB}] \cdot \mathbb{E}[\overline{CD}]$?~~

- Can we instead get some kind of sum?

$$\overline{AB} \cdot \overline{CD} = (10 \cdot A + B) \cdot (10 \cdot C + D) = 100 \cdot A \cdot C + 10 \cdot A \cdot D + 10 \cdot B \cdot C + B \cdot D$$

- Now by linearity of expectation,

?

Linearity of Expectation

EXAMPLE 6 (CONT.):

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$$\mathbb{E}[\overline{AB} \cdot \overline{CD}] = 100 \cdot \mathbb{E}[A \cdot C] + 10 \cdot \mathbb{E}[A \cdot D] + 10 \cdot \mathbb{E}[B \cdot C] + \mathbb{E}[B \cdot D] = \quad ?$$

Linearity of Expectation

EXAMPLE 6 (CONT.):

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The expected value of all of these products are the same since there is symmetry among A, B, C, D .

Linearity of Expectation

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The expected value of all of these products are the same since there is symmetry among A, B, C, D .

$$\mathbb{E}[A \cdot C] = \frac{1 \cdot 2 + 1 \cdot 3 + 1 \cdot 4 + 2 \cdot 3 + 2 \cdot 4 + 3 \cdot 4}{6} = \frac{35}{6}$$

Variance

Measuring variability

EXAMPLE: average size of mice

case 1:

{4 cm, 5 cm, 6 cm}

case 2:

{9 cm, 10 cm, 11 cm}

case 3:

{8 cm, 10 cm, 12 cm}

*What is a reasonable
measure of centrality?*

Measuring variability

EXAMPLE: average size of mice

	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}	What is a reasonable measure of centrality?
$\mathbb{E}[X] =$	5	10	10	$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$ mean = average = "expected value"

Measuring variability

EXAMPLE: average size of mice

	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}	<i>What are possible ways to measure expected "variability" Δ around the mean for each point?</i>
$\mathbb{E}[X] =$	5	10	10	$\mathbb{E}[Y] = \sum_y y \cdot p_Y(y)$ $\mathbb{E}[X] = \sum_x x \cdot p_X(x)$ <i>?</i>

Measuring variability

EXAMPLE: average size of mice

What are possible ways to measure expected "variability" Δ around the mean for each point?

$$\mathbb{E}[Y] = \sum_y y \cdot p_Y(y)$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

$$Y_1 = X - \mathbb{E}[X]$$

How can we fix that ?

	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}
$\mathbb{E}[X] =$	5	10	10
$\mathbb{E}[Y_1] =$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ = 0	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ = 0	$-2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ = 0

☹️ cancels out

Measuring variability

EXAMPLE: average size of mice

What are possible ways to measure expected "variability" Δ around the mean for each point?

$$\mathbb{E}[Y] = \sum_y y \cdot p_Y(y)$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

$$Y_1 = X - \mathbb{E}[X]$$

$$Y_2 = |X - \mathbb{E}[X]|$$

absolute deviation

Anything else ?

	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}
$\mathbb{E}[X] =$	5	10	10
$\mathbb{E}[Y_1] =$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= 0$ ☹️ <i>cancels out</i>
$\mathbb{E}[Y_2] =$	$ -1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$	$ -1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$	$ -2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= 2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= \frac{4}{3}$ 😊 <i>makes more sense</i>

Measuring variability

EXAMPLE: average size of mice

What are possible ways to measure expected "variability" Δ around the mean for each point?

$$\mathbb{E}[Y] = \sum_y y \cdot p_Y(y)$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

$$Y_1 = X - \mathbb{E}[X]$$

$$Y_2 = |X - \mathbb{E}[X]|$$

absolute deviation

$$Y_3 = (X - \mathbb{E}[X])^2$$

squared error

	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}
$\mathbb{E}[X] =$	5	10	10
$\mathbb{E}[Y_1] =$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= 0$
$\mathbb{E}[Y_2] =$	$ -1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$	$ -1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$	$ -2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= 2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= \frac{4}{3}$
$\mathbb{E}[Y_3] =$	$1^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{1}{3}$ $1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$	$1^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{1}{3}$ $1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$	$2^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 2^2 \cdot \frac{1}{3}$ $4 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 4 \cdot \frac{1}{3}$ $= \frac{8}{3}$

☹️ cancels out

😊 makes more sense

What are now the "units" of variability



Measuring variability

EXAMPLE: average size of mice

What are possible ways to measure expected "variability" Δ around the mean for each point?

$$\mathbb{E}[Y] = \sum_y y \cdot p_Y(y)$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

$$Y_1 = X - \mathbb{E}[X]$$

$$Y_2 = |X - \mathbb{E}[X]|$$

absolute deviation

$$Y_3 = (X - \mathbb{E}[X])^2$$

squared error

	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}
$\mathbb{E}[X] =$	5	10	10
$\mathbb{E}[Y_1] =$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= 0$ ☹️ <i>cancels out</i>
$\mathbb{E}[Y_2] =$	$ -1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm	$ -1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm	$ -2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= 2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= \frac{4}{3}$ cm 😊 <i>makes more sense</i>
$\mathbb{E}[Y_3] =$	$1^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{1}{3}$ $1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm ²	$1^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{1}{3}$ $1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm ²	$2^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 2^2 \cdot \frac{1}{3}$ $4 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 4 \cdot \frac{1}{3}$ $= \frac{8}{3}$ cm ² ☹️ ???

cm squared is strange. What can we do?

?

Measuring variability

EXAMPLE: average size of mice

What are possible ways to measure expected "variability" Δ around the mean for each point?

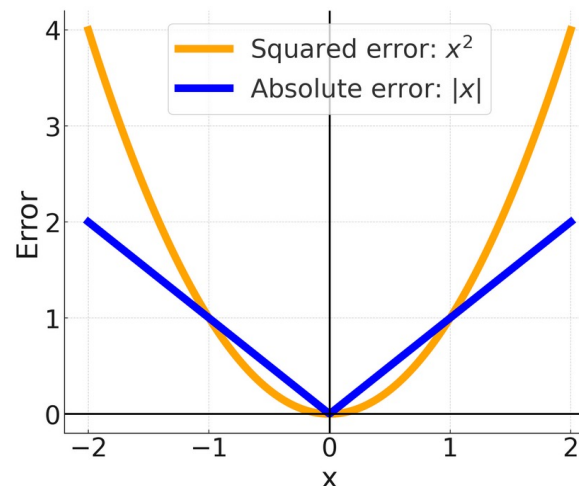
	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}	$\mathbb{E}[Y] = \sum_y y \cdot p_Y(y)$
$\mathbb{E}[X] =$	5	10	10	$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$
$\mathbb{E}[Y_1] =$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= 0$ ☹️ <i>cancels out</i>	$Y_1 = X - \mathbb{E}[X]$
$\mathbb{E}[Y_2] =$	$ -1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm	$ -1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm	$ -2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= 2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= \frac{4}{3}$ cm ☺️ <i>makes more sense</i>	$Y_2 = X - \mathbb{E}[X] $ absolute deviation
$\mathbb{E}[Y_3] =$	$1^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{1}{3}$ $1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm ²	$1^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{1}{3}$ $1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm ²	$2^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 2^2 \cdot \frac{1}{3}$ $4 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 4 \cdot \frac{1}{3}$ $= \frac{8}{3}$ cm ² ☹️ ???	$Y_3 = (X - \mathbb{E}[X])^2$ squared error
Just take the square root:	$= \sqrt{\frac{2}{3}}$ cm	$= \sqrt{\frac{2}{3}}$ cm	$= 2\sqrt{\frac{2}{3}}$ cm ???	Looks pretty complicated. So why is everyone so excited about squared error instead of absolute error? ?

Why variance, and not absolute deviation?

But what is the most substantial reason?



- Geometric explanation in Euclidean space (distance = square root of squared components): Projection onto a subspace. Variance is literally the "average squared distance" from the mean.
- Variance plays an important role in the Central Limit Theorem. The variance is a natural parameter for the Normal Distribution.
- Algebraic convenience: $f(x) = x^2$ is differentiable, but $g(x) = |x|$ is not.
- Outliers have a bigger influence



$$\Delta_2 = |X - \mathbb{E}[X]|$$

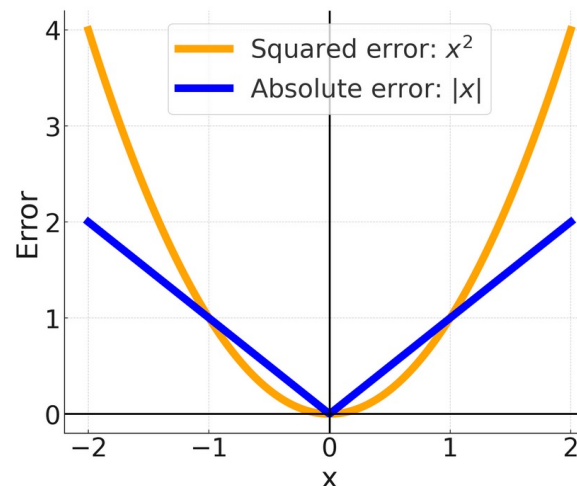
absolute deviation

$$\Delta_3 = (X - \mathbb{E}[X])^2$$

squared error

Why variance, and not absolute deviation?

- Statistical properties: The mean minimizes the sum of squared errors, while the median minimizes the sum of absolute errors. (Maximum likelihood estimation leads to minimizing squared errors)
- Geometric explanation in Euclidean space (distance = square root of squared components): Projection onto a subspace. Variance is literally the "average squared distance" from the mean.
- Variance plays an important role in the Central Limit Theorem. The variance is a natural parameter for the Normal Distribution.
- Algebraic convenience: $f(x) = x^2$ is differentiable, but $g(x) = |x|$ is not.
- Outliers have a bigger influence



$$\Delta_2 = |X - \mathbb{E}[X]|$$

absolute deviation

$$\Delta_3 = (X - \mathbb{E}[X])^2$$

squared error

Mean (minimizes variance), Median (minimizes absolute error)

EXAMPLE (MEAN VS. MEDIAN):

data = {1, 2, 3, 6, 8}

mean = ?

median = ?

Mean (minimizes variance), Median (minimizes absolute error)

EXAMPLE (MEAN VS. MEDIAN):

data = {1, 2, 3, 6, 8}

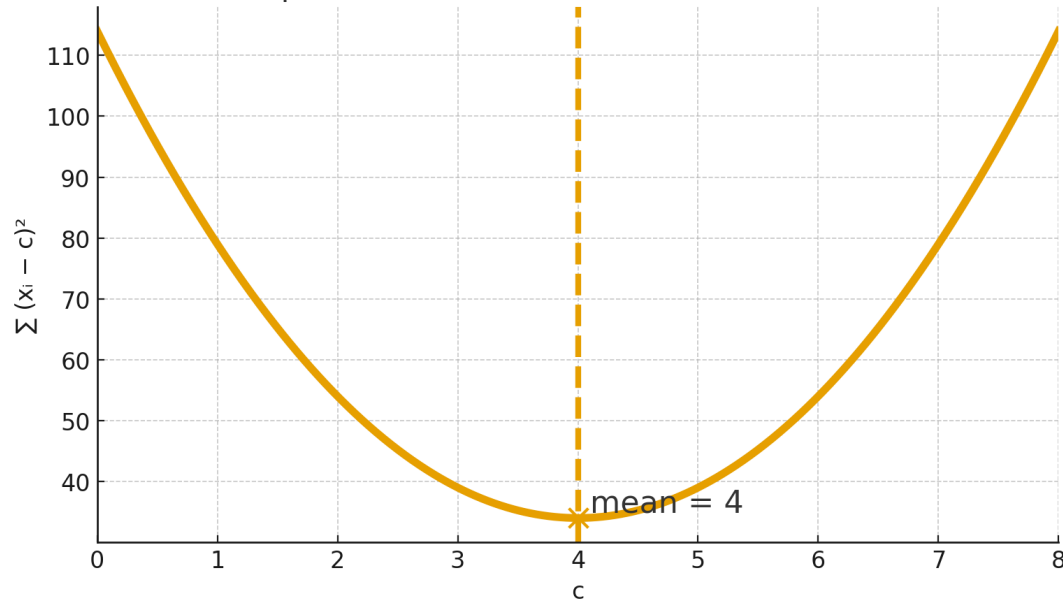
mean = 4

$$\text{mean} = \min_c \sum_i (x_i - c)^2$$

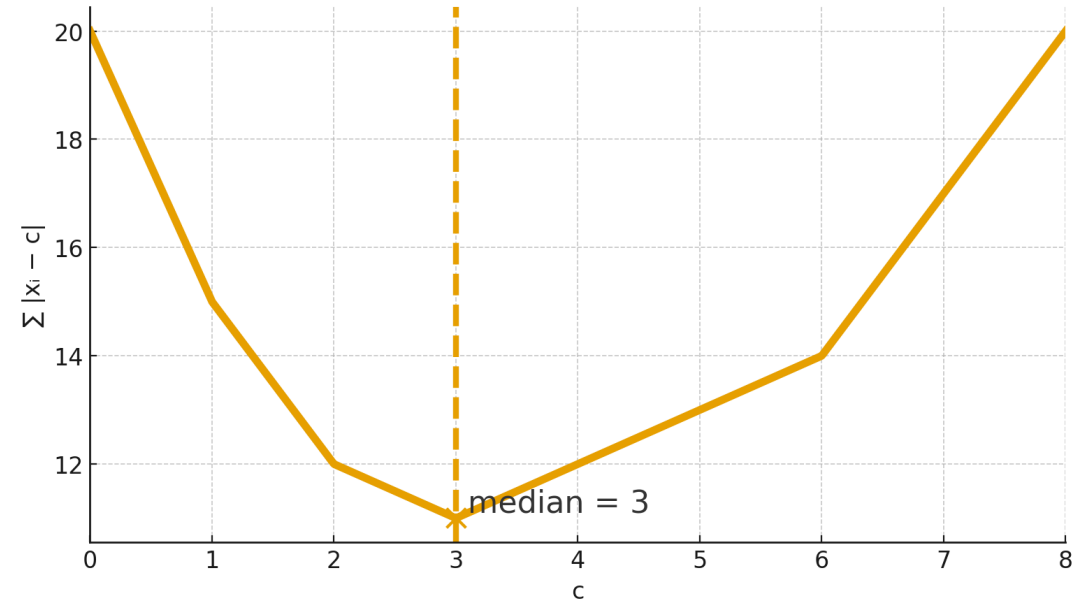
median = 3

$$\text{median} = \min_c \sum_i |x_i - c|$$

Squared-error loss minimized at the mean



Absolute-error loss minimized at the median



Mean (minimizes variance), Median (minimizes absolute error)

EXAMPLE (MEAN VS. MEDIAN):

data = {1, 2, 3, 6, 8}

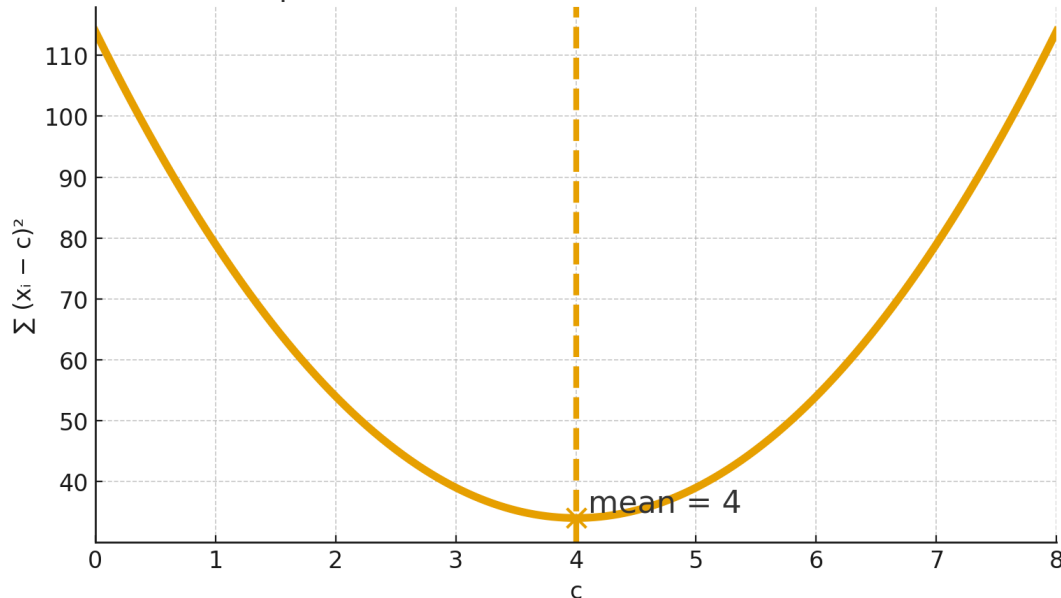
mean = 4

$$\text{mean} = \min_c \sum_i (x_i - c)^2$$

$$\text{mean: } \sum_i x_i - c = 0$$

?

Squared-error loss minimized at the mean



Mean (minimizes variance), Median (minimizes absolute error)

EXAMPLE (MEAN VS. MEDIAN):

data = {1, 2, 3, 6, 8}

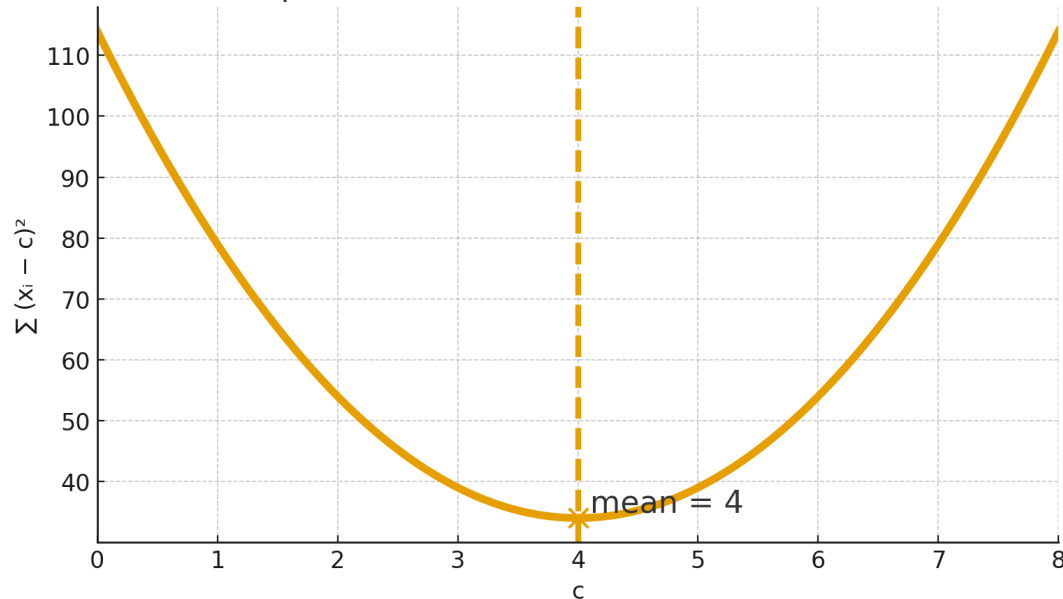
mean = 4

$$\text{mean} = \min_c \sum_i (x_i - c)^2$$

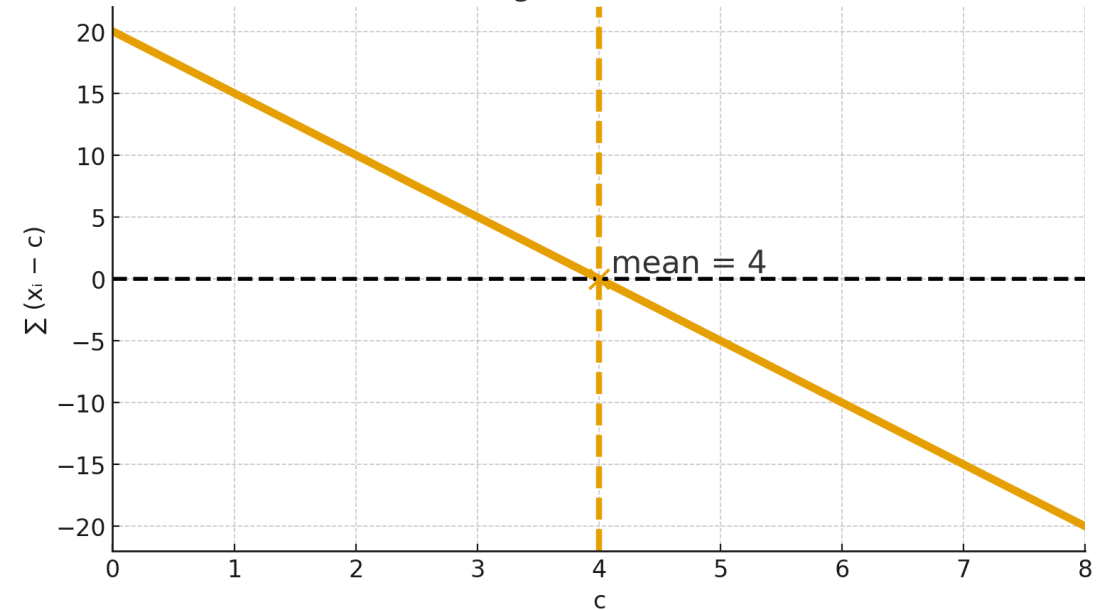
This is not an optimization problem, but a constraint problem: the sum of signed errors = 0.

$$\text{mean: } \sum_i x_i - c = 0$$

Squared-error loss minimized at the mean



Total error (signed sum of deviations)



Alternative formula for variance

$$\sigma^2 = \text{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] \quad \text{variance}$$

$$\sigma = \sqrt{\text{Var}(X)} = \sqrt{\mathbb{E}[(X - \mathbb{E}[X])^2]} \quad \text{standard deviation (back in original units)}$$

ALTERNATIVE FORMULA

$$\mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$

PROOF

$$\mathbb{E}[(X - \mathbb{E}[X])^2] = \text{?}$$

case 3:

{8 cm, 10 cm, 12 cm}

$$\mathbb{E}[X^2] = \frac{8^2 + 10^2 + 12^2}{3} = \frac{308}{3}$$

$$(\mathbb{E}[X])^2 = 100$$

$$\frac{308}{3} - \frac{300}{3} = 8/3$$

Alternative formula for variance

$$\sigma^2 = \text{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] \quad \text{variance}$$

$$\sigma = \sqrt{\text{Var}(X)} = \sqrt{\mathbb{E}[(X - \mathbb{E}[X])^2]} \quad \text{standard deviation (back in original units)}$$

ALTERNATIVE FORMULA

$$\mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$

PROOF

$$\begin{aligned} \mathbb{E}[(X - \mathbb{E}[X])^2] &= \mathbb{E}[X^2 - 2 \cdot X \cdot \mathbb{E}[X] + (\mathbb{E}[X])^2] \\ &\quad \text{by linearity of expectation} \\ &= \mathbb{E}[X^2] - \mathbb{E}[2 \cdot X \cdot \mathbb{E}[X]] + \mathbb{E}[(\mathbb{E}[X])^2] \\ &\quad \text{just a constant} \\ &= \mathbb{E}[X^2] - 2 \cdot \mathbb{E}[X] \cdot \mathbb{E}[X] + (\mathbb{E}[X])^2 \\ &= \mathbb{E}[X^2] - 2 \cdot (\mathbb{E}[X])^2 + (\mathbb{E}[X])^2 \\ &= \mathbb{E}[X^2] - (\mathbb{E}[X])^2 \end{aligned}$$

case 3:

{8 cm, 10 cm, 12 cm}

$$\mathbb{E}[X^2] = \frac{8^2 + 10^2 + 12^2}{3} = \frac{308}{3}$$

$$(\mathbb{E}[X])^2 = 100$$

$$\frac{308}{3} - \frac{300}{3} = 8/3$$