

# Part 1: Information Theory Basics

## L02: Basics of Probability (1/2)

[Random experiment, independence, conditional probability, chain rule, Bayes' theorem, random variables]

Wolfgang Gatterbauer

cs7840 Foundations and Applications of Information Theory (fa25)

<https://northeastern-datalab.github.io/cs7840/fa25/>

9/11/2025

# Pre-class conversations

- Last class recapitulation
- Organizational matters: Does every one get Piazza messages?
- First scribe on Piazza. Awesome! I will look over the weekend
- Office hours: Usually right after class, or via email / Teams. Also feel free talk to me regularly during break/ after class about project ideas, or papers/research directions to include towards the end
- New class arrivals
- Today:
  - The basics of probability theory

# Basics of probability theory

Following slides are built upon examples by Jay Aslam from earlier editions of this class. For a more extensive cover of the basics, I recommend "Bertsekas, Tsitsiklis. Introduction to Probability, 2008." It's a solid textbook on probability theory and I regularly find myself going back to this book to look up basic concepts. Working by yourself through chapter 1 on "Sample Space and Probability" is a good investment of your time.

# Sample space / outcomes / events / probability



## PROBABILITY

- A random experiment
- Sample space  $\Omega$ : set of all possible outcomes

## EXAMPLE 1:

roll a fair die with 6 sides

$\Omega =$  ?

# Sample space / outcomes / events / probability



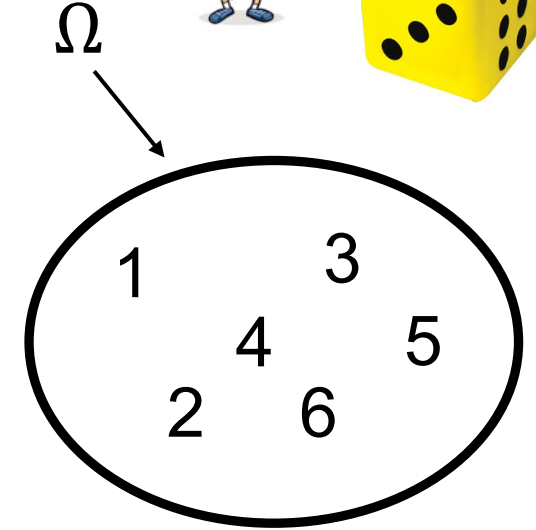
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- Sample space  $\Omega$ : set of all possible outcomes
- An outcome  $\omega \in \Omega$  of the exp.

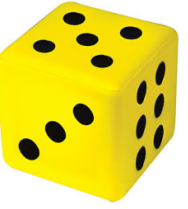
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roll a fair die with 6 sides

$$\Omega = \{1, 2, 3, 4, 5, 6\}$$



# Sample space / outcomes / events / probability



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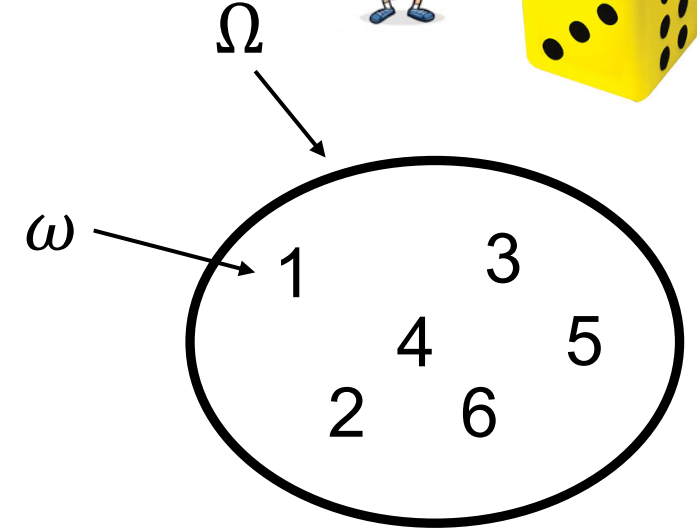
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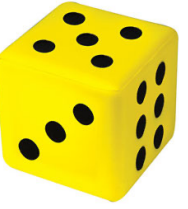
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Say the outcome is a 1



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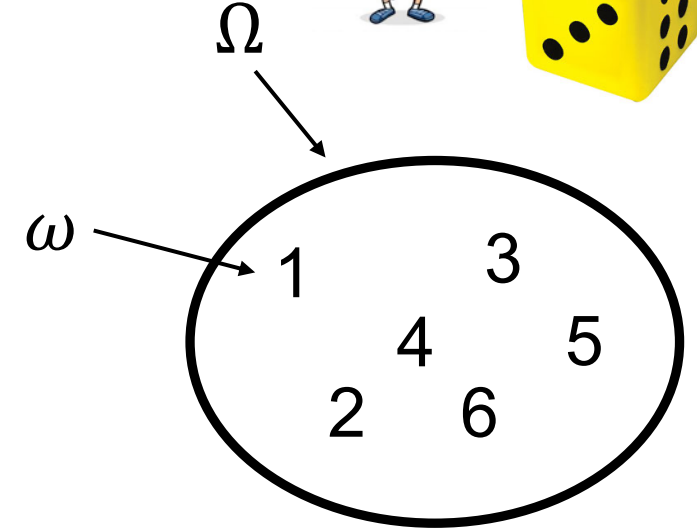
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Say the outcome is a 1

$$E_1 = \text{"even"}$$

$$E_2 = \text{"}\geq 3\text{"}$$



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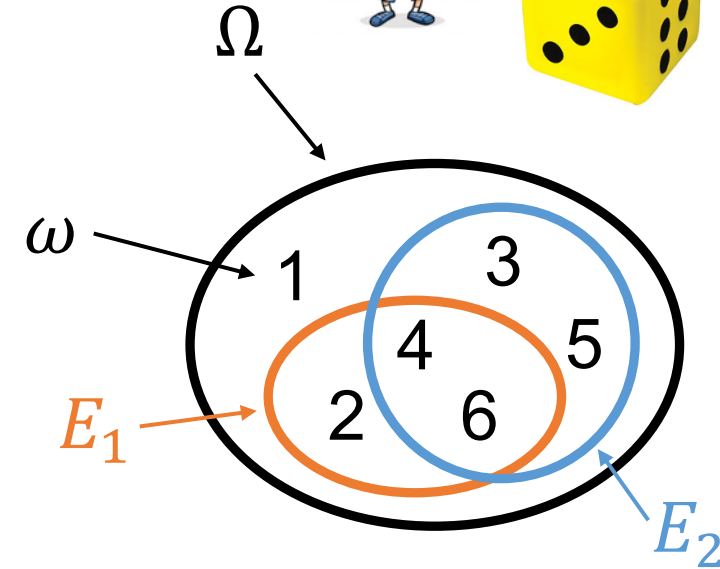
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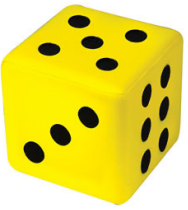
$$E_2 = \text{"}\geq 3\text{"} = \{3, 4, 5, 6\}$$



?



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 ~~$\mathbb{P}: \Omega \rightarrow [0, 1]$~~

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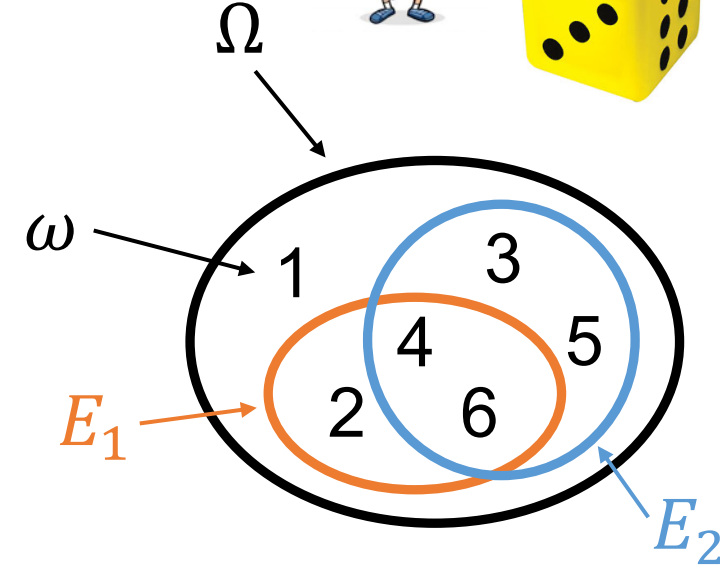
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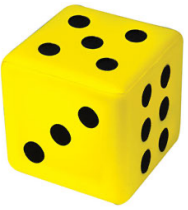
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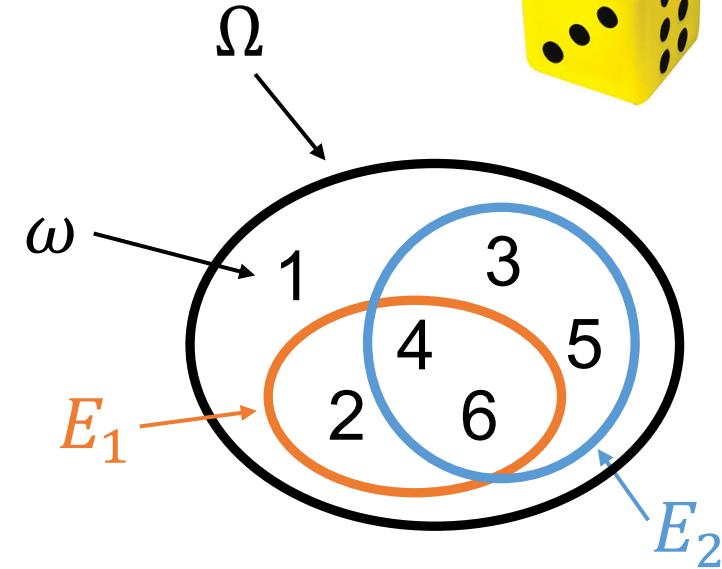
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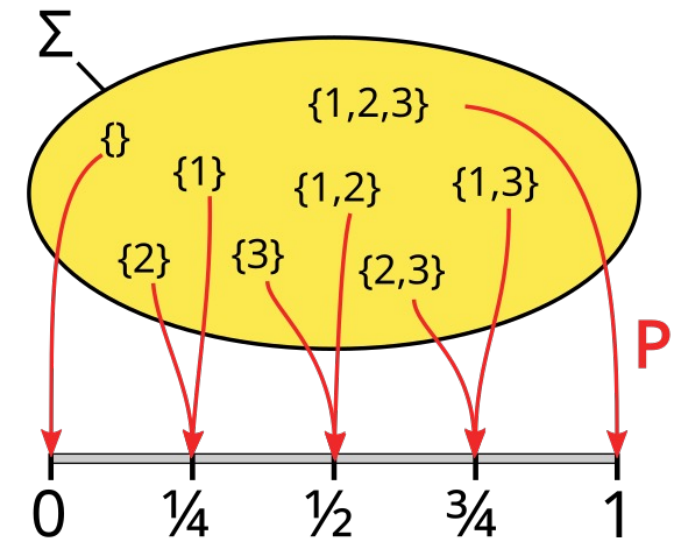
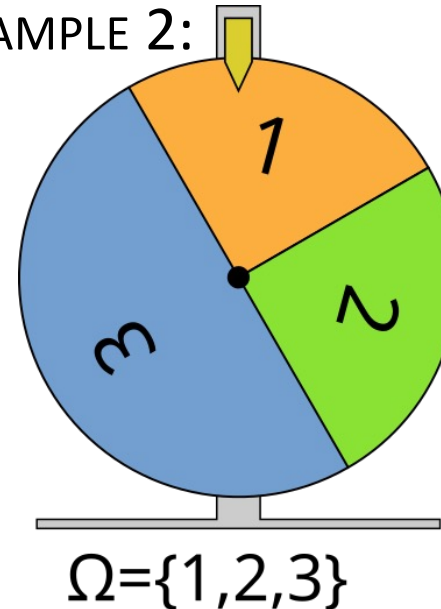
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## EXAMPLE 2:



The set  $\Sigma$  of all events  $E \in \Sigma$  is a  $\sigma$ -algebra, basically the same as the power set  $2^\Omega$ , except for some pathological non-measurable sets like the Vitali-Set ([https://www.youtube.com/watch?v=hs3eDa3\\_DzU](https://www.youtube.com/watch?v=hs3eDa3_DzU)). See also: <https://en.wikipedia.org/wiki/%CE%A3-algebra>

Source of figure to the right: [https://en.wikipedia.org/wiki/Probability\\_measure](https://en.wikipedia.org/wiki/Probability_measure)

Gatterbauer. Foundations and Applications of Information Theory: <https://northeastern-datalab.github.io/cs7840/>

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~~$$\mathbb{P}: \Omega \rightarrow [0, 1]$$~~

$$\mathbb{P}: \Sigma \rightarrow [0, 1]$$

also  $\sum_{\omega \in \Omega} \mathbb{P}(\{\omega\}) = 1$

- 1) Nonnegativity:  $\mathbb{P}(E) \geq 0, \forall E \in \Sigma$
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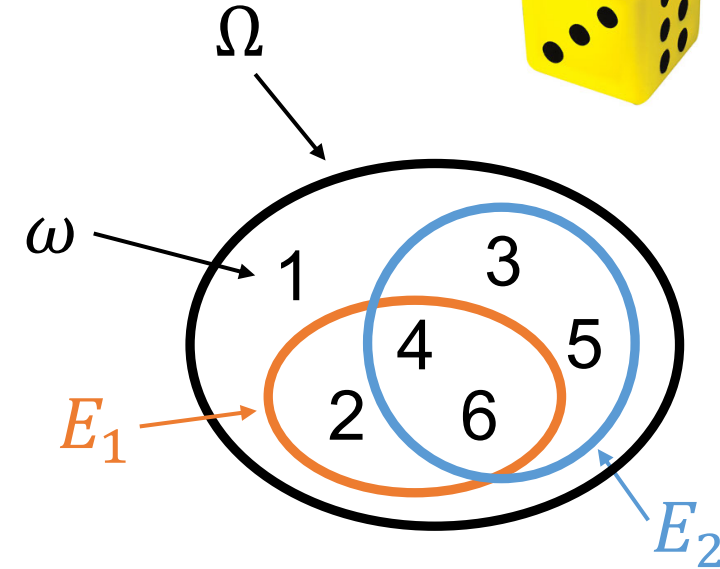
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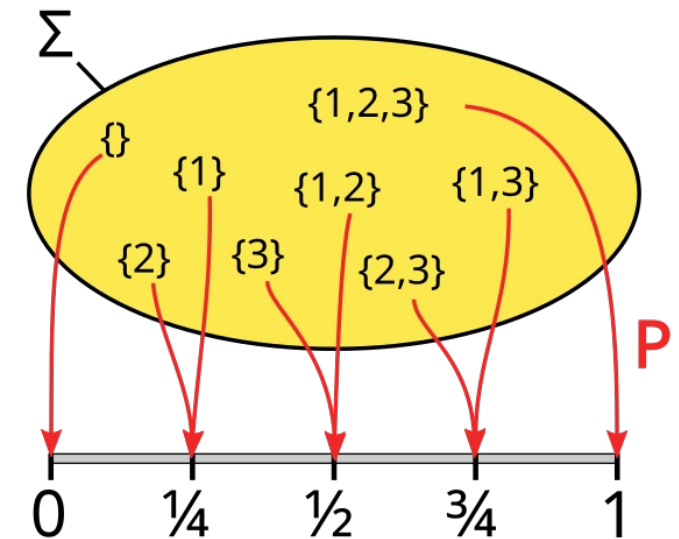
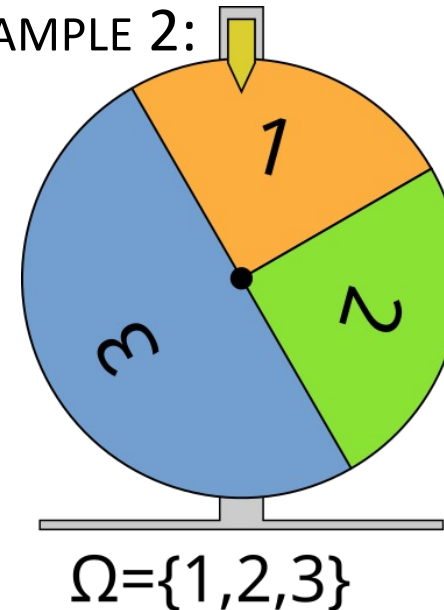
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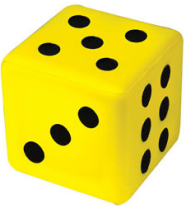
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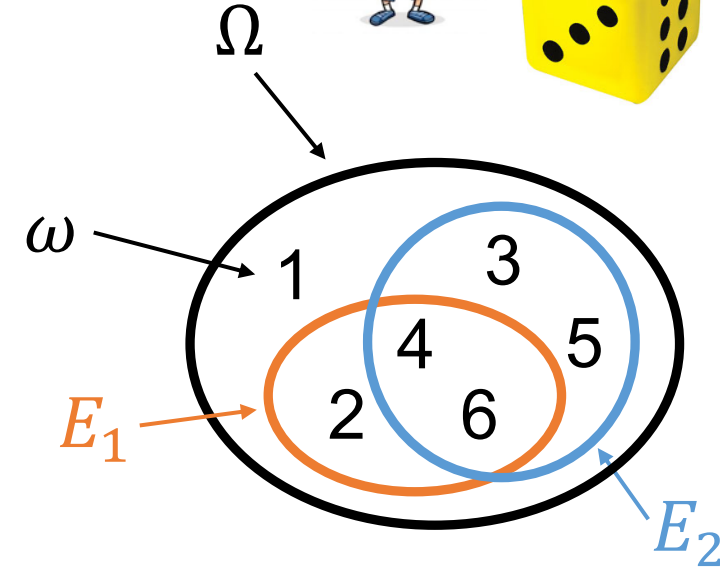
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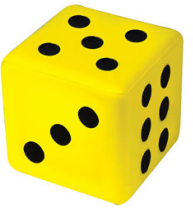
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short for  $\mathbb{P}(\{\omega\})$

- If  $\mathbb{P}(\omega) = \frac{1}{|\Omega|}, \forall \omega \in \Omega$ , then  $\mathbb{P}(E) = ?$

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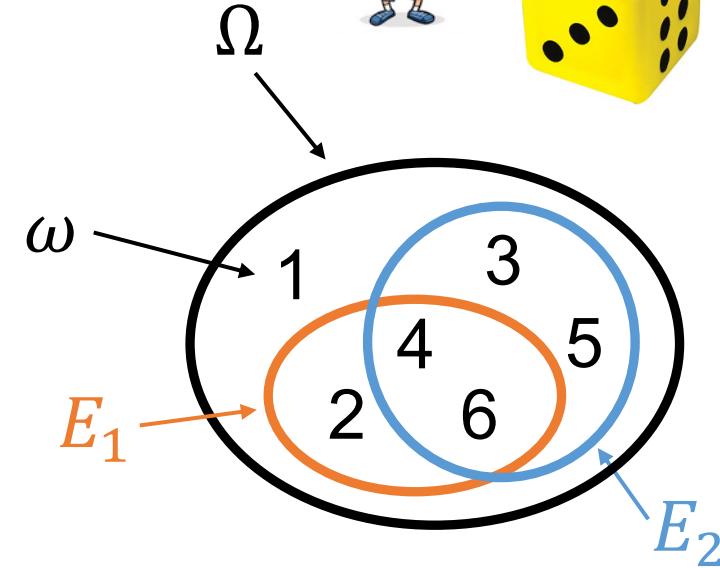
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$$\mathbb{P}(1) = \mathbb{P}(2) = \dots = \mathbb{P}(6) = \frac{1}{6}$$

$$\mathbb{P}(E_1) =$$

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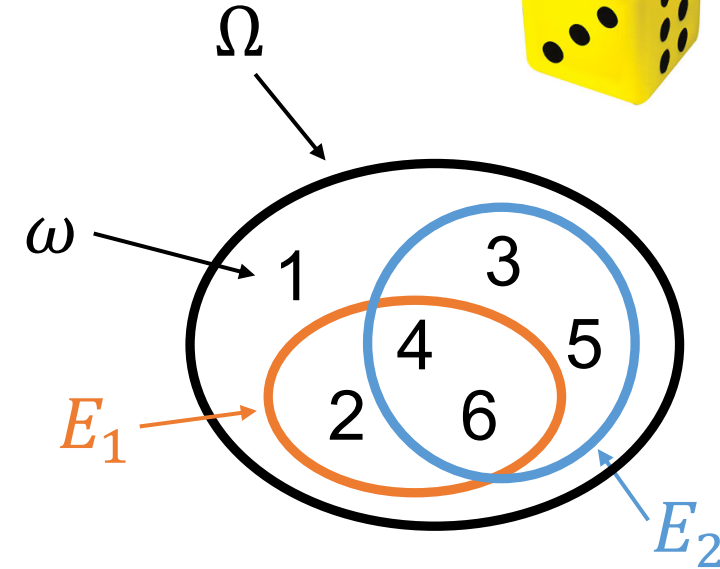
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$$\mathbb{P}(1) = \mathbb{P}(2) = \dots = \mathbb{P}(6) = \frac{1}{6}$$

$$\mathbb{P}(E_1) = \mathbb{P}(2) + \mathbb{P}(4) + \mathbb{P}(6) = \frac{3}{6} = \frac{1}{2}$$

$$\mathbb{P}(E_2) = \frac{4}{6} = \frac{2}{3}$$

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$\Omega =$

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$|\Omega| =$

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$$\begin{aligned}\Omega &= \{1, 2, 3, 4, 5, 6\} \times \{1, 2, 3, 4, 5, 6\} \\ &= \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}\end{aligned}$$

$$|\Omega| = 36$$

		second die					
		1	2	3	4	5	6
first die	1						
	2						
	3						
	4						
	5						
	6						



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		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

$$E_1 = \text{"sum is 7"}$$

$$|E_1| =$$

$$\mathbb{P}(E_1) =$$



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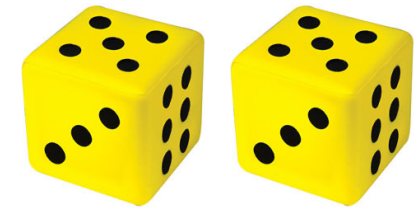
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$$\begin{aligned}E_1 &= \text{"sum is 7"} \\ &= \{(1,6), (2,5), (3,4), (4,3), (5,2), (6,1)\}\end{aligned}$$

$$|E_1| = 6$$

$$\mathbb{P}(E_1) = \frac{|E_1|}{|\Omega|} = \frac{6}{36} = \frac{1}{6}$$

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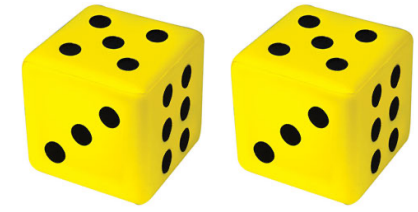
$$E_2 = \text{"total is greater than 8"}$$

$$|E_2| =$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|}$$



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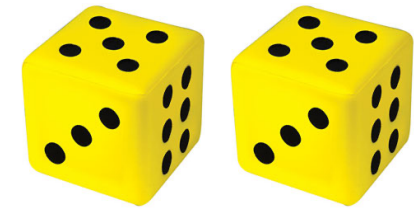
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$$\begin{aligned}E_2 &= \text{"total is greater than 8"} \\ &= \text{"total is 9 or 10 or 11 or 12"}\end{aligned}$$

$$|E_2| =$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|}$$

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$$|\Omega| = 36$$

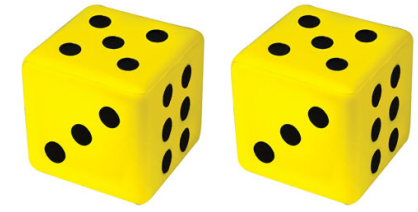
		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

$$\begin{aligned}E_2 &= \text{"total is greater than 8"} \\ &= \text{"total is 9 or 10 or 11 or 12"}\end{aligned}$$

$$|E_2| = 4$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|}$$

# Sample space / outcomes / events / probability



## PROBABILITY

- A random experiment
- Sample space  $\Omega$ : set of all possible outcomes
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$$|\Omega| = 36$$

		second die					
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first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
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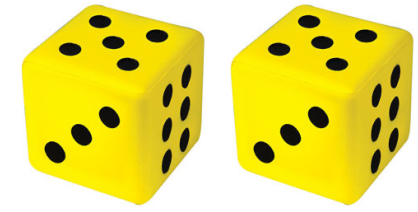
$E_2$  = "total is greater than 8"

= "total is 9 or 10 or 11 or 12"

$$|E_2| = \begin{array}{c} \downarrow \quad \downarrow \\ 4 + 3 \end{array}$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|}$$

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	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
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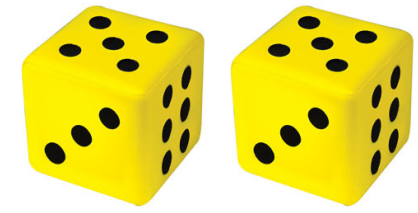
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# Sample space / outcomes / events / probability



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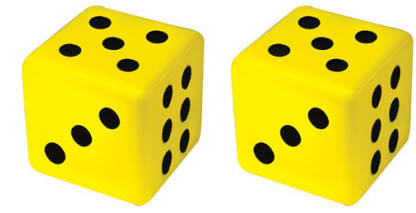
= "total is 9 or 10 or 11 or 12"

$$|E_2| = \begin{array}{ccccccc} & & \downarrow & & \downarrow & & \downarrow & & \downarrow \\ & & 4 & + & 3 & + & 2 & + & 1 \end{array}$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|}$$



# Sample space / outcomes / events / probability



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$E_2$  = "total is greater than 8"

= "total is 9 or 10 or 11 or 12"

additivity of  
disjoint events

$$|E_2| = \begin{array}{ccccccc} & \downarrow & \downarrow & \downarrow & \downarrow & & \\ & 4 & + & 3 & + & 2 & + & 1 & = & 10 \end{array}$$

$$\mathbb{P}(E_2) = \frac{|E_2|}{|\Omega|} = \frac{10}{36} = \frac{5}{18}$$

# Conditional Probabilities / Independence


# Conditional probability

## DEFINITION

Conditional probability  $\mathbb{P}(A|B)$  is the probability that event  $A$  occurred, given that event  $B$  occurred

EXAMPLE 4  
(FAIR 16-SIDED DIE):

$\Omega$



1	2	3	5	6	7	10		11
13				16				
14	4	12		8	9		15	

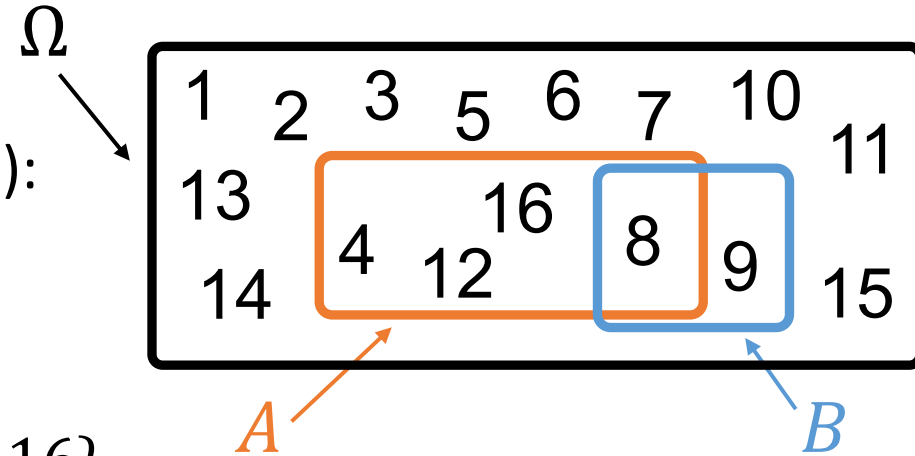
$$\Omega = \{1, 2, \dots, 16\}$$

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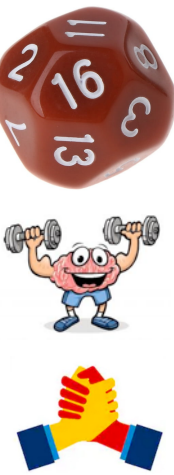


$$\Omega = \{1, 2, \dots, 16\}$$

$$A = \text{"divisible by 4"} = \{4, 8, 12, 16\}$$

$$B = \text{"center numbers"} = \{8, 9\}$$

$$\mathbb{P}(A|B) = ?$$

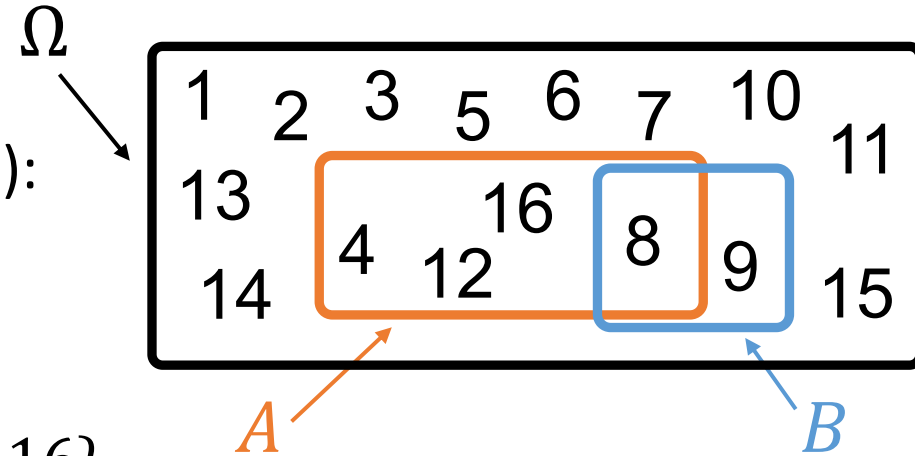


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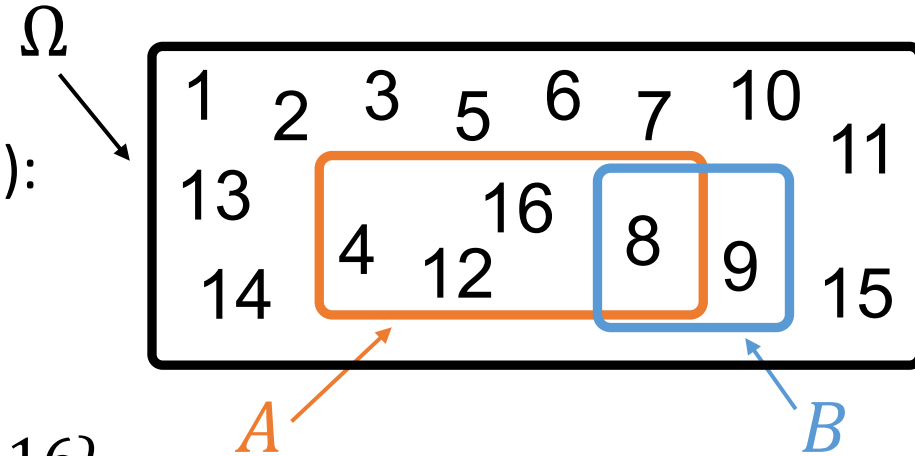
$$\mathbb{P}(A|B) = \frac{1}{2}$$

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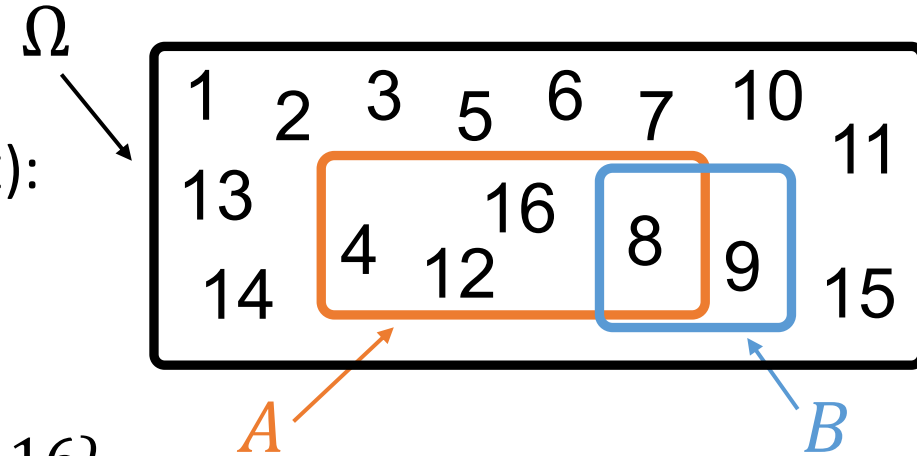
$$\mathbb{P}(A|B) = \frac{\quad}{\mathbb{P}(B)}$$

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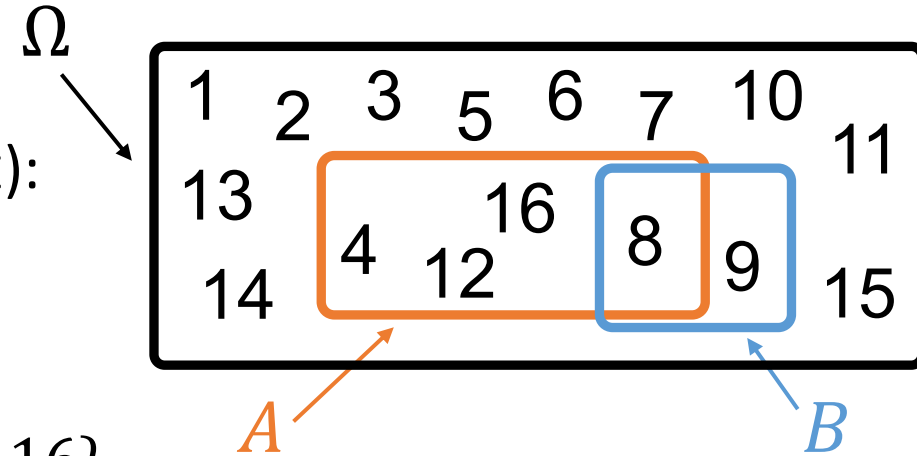
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# Conditional probability

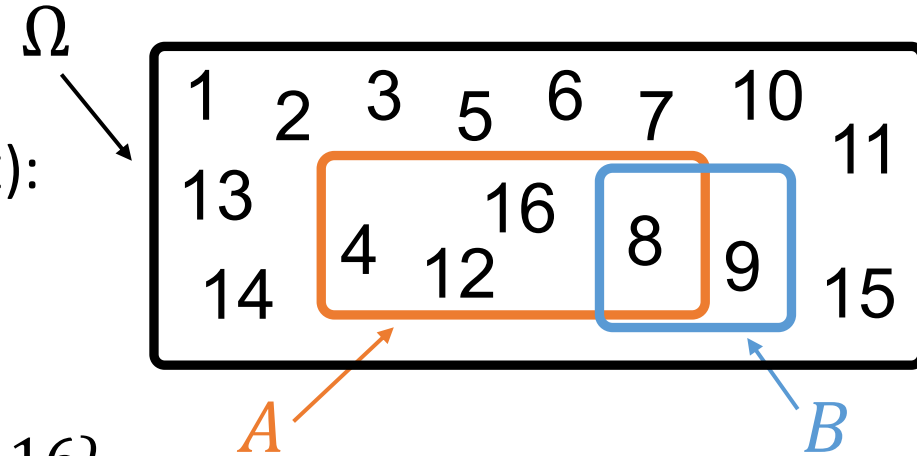
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$\mathbb{P}(A \cap B)$

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(FAIR 16-SIDED DIE):



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$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)} = \frac{\frac{1}{16}}{\frac{2}{16}} = \frac{1}{2}$$

# Independence

## DEFINITION:

Two events are independent if the probability that one occurred is not affected by knowledge that the other occurred.

$$\mathbb{P}(A|B) = \mathbb{P}(A) \quad (\text{if } \mathbb{P}(B) \neq 0)$$

$$\mathbb{P}(B|A) = \mathbb{P}(B) \quad (\text{if } \mathbb{P}(A) \neq 0)$$

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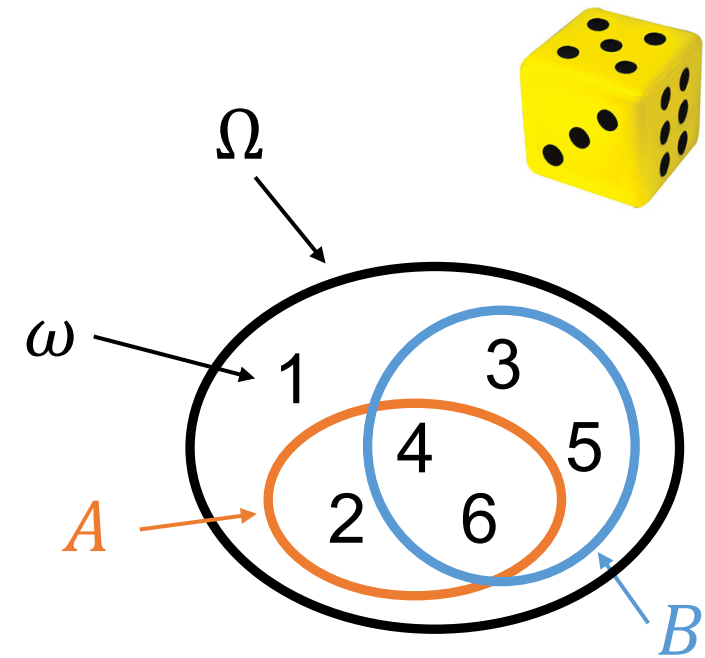
## EXAMPLE 1 (CONTINUED):

roll a fair die with 6 sides  
the outcome is a 1

$$\Omega = \{1, 2, 3, 4, 5, 6\}$$

$$A = \text{"even"} = \{2, 4, 6\}$$

$$B = \text{"}\geq 3\text{"} = \{3, 4, 5, 6\}$$



Are  $A$  and  $B$  independent ?

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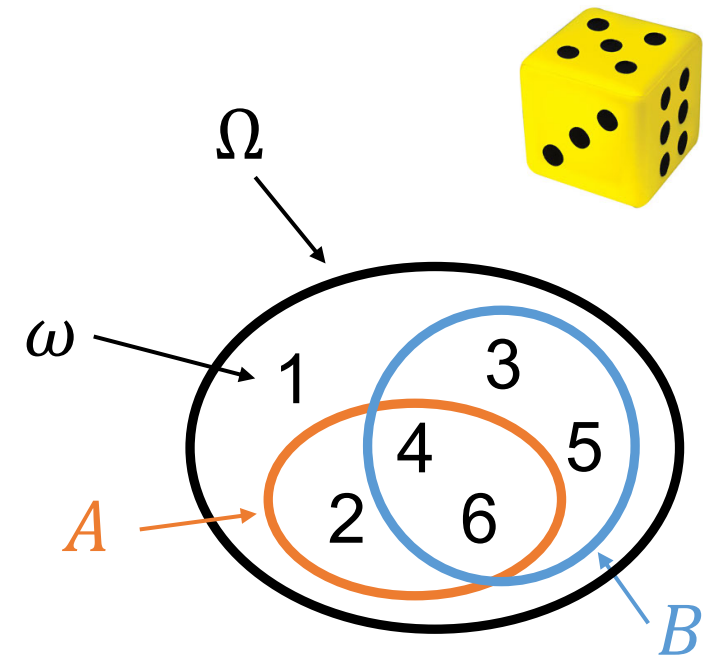
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*A and B are independent !*

$$\mathbb{P}(A) = \frac{1}{2}$$

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)} = \frac{1}{2}$$

1	3	5
2	4	6

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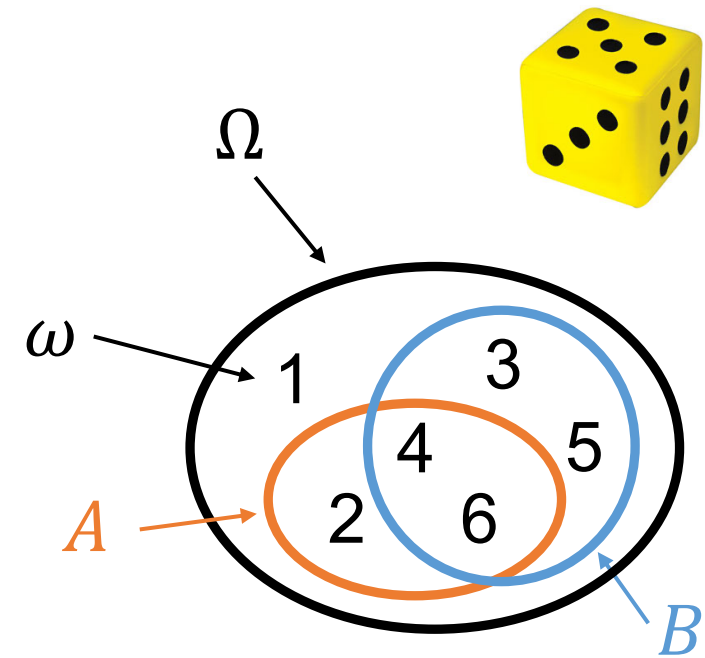
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$$\mathbb{P}(A) = \frac{1}{2}$$

		$\mathbb{P}(B) = \frac{2}{3}$	
	1	3	5
	2	4	6

*Think about a coordinate system:  
x and y are independent*

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Recall the definition of cond. prob.:

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

## CLAIM

Two events  $A$  and  $B$  are independent if and only if  $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$

PROOF (assuming  $\mathbb{P}(A) \neq 0, \mathbb{P}(B) \neq 0$ )



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( $\Rightarrow$ ) If  $\mathbb{P}(A|B) = \mathbb{P}(A)$ :

?

( $\Leftarrow$ ) If  $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$ :

?

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( $\Rightarrow$ ) If  $\mathbb{P}(A|B) = \mathbb{P}(A)$ :

$$\text{Then } \mathbb{P}(A, B) = \mathbb{P}(A|B) \cdot \mathbb{P}(B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$

( $\Leftarrow$ ) If  $\mathbb{P}(A, B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$ :

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)} = \frac{\mathbb{P}(A) \cdot \cancel{\mathbb{P}(B)}}{\cancel{\mathbb{P}(B)}} = \mathbb{P}(A)$$

$\nearrow \neq 0$



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Recall the definition of cond. prob.:

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

Commonly used as definition of independence because it is slightly more general: if one event has probability 0 then the statement holds vacuously (while one conditional probability is undefined). And if both are not 0, then we just proved

DEFINITION: that to be equivalent to the earlier definition.

$$\text{Two events } A \text{ and } B \text{ are independent} \Leftrightarrow \mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$

PROOF (assuming  $\mathbb{P}(A) \neq 0, \mathbb{P}(B) \neq 0$ )

( $\Rightarrow$ ) If  $\mathbb{P}(A|B) = \mathbb{P}(A)$ :

$$\text{Then } \mathbb{P}(A, B) = \mathbb{P}(A|B) \cdot \mathbb{P}(B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$

( $\Leftarrow$ ) If  $\mathbb{P}(A, B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$ :

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)} = \frac{\mathbb{P}(A) \cdot \cancel{\mathbb{P}(B)}}{\cancel{\mathbb{P}(B)}} = \mathbb{P}(A)$$

$\nearrow$   
 $\neq 0$

# Independence

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 $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$

EXAMPLE 3 (CONTINUED):

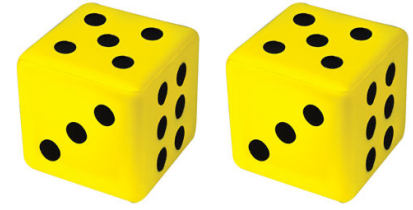
roll two fair dice with 6 sides

$$\Omega = \{1, 2, 3, 4, 5, 6\} \times \{1, 2, 3, 4, 5, 6\}$$

$A$  = "1<sup>st</sup> roll is 1"

$B$  = "sum is 7"

Are  $A$  and  $B$  independent? ?



# Independence

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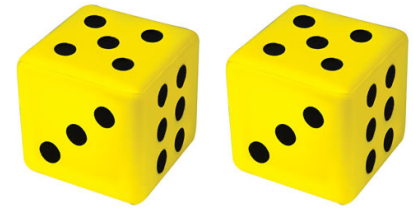
$B$  = "sum is 7"

Are  $A$  and  $B$  independent?

$$\mathbb{P}(A) = ?$$

$$\mathbb{P}(B) = ?$$

$$\mathbb{P}(A, B) = ?$$



# Independence

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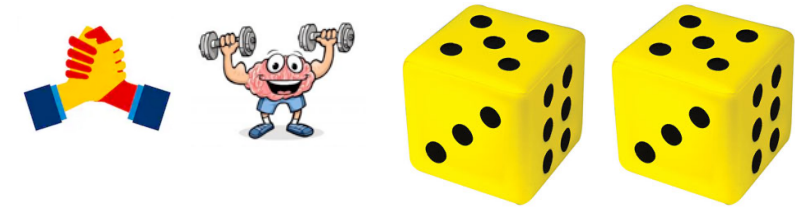
$B$  = "sum is 7"

Are  $A$  and  $B$  independent?

$$\mathbb{P}(A) = \frac{1}{6}$$

$$\mathbb{P}(B) = \frac{1}{6}$$

$$\mathbb{P}(A, B) = ?$$



second die

	1	2	3	4	5	6
1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

first die

# Independence

DEFINITION:

Two events  $A$  and  $B$  are independent if  
 $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$

EXAMPLE 3 (CONTINUED):

roll two fair dice with 6 sides

$$\Omega = \{1, 2, 3, 4, 5, 6\} \times \{1, 2, 3, 4, 5, 6\}$$

$A$  = "1<sup>st</sup> roll is 1"

$B$  = "sum is 7"

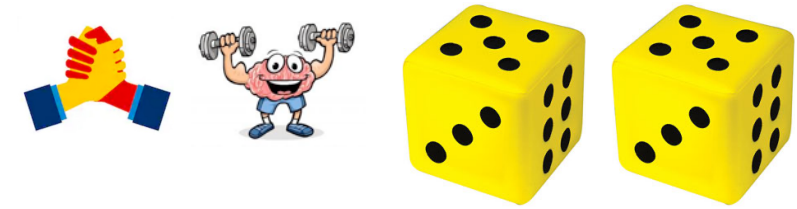
Are  $A$  and  $B$  independent?

$$\mathbb{P}(A) = \frac{1}{6}$$

$$\mathbb{P}(B) = \frac{1}{6}$$

$$\mathbb{P}(A, B) = \frac{1}{36}$$

Yes:  $\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$



second die

	1	2	3	4	5	6
1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

first die

Handwritten annotations: An orange circle highlights the first row (1,1) through (1,6). A blue circle highlights the first column (1,1) through (6,1). A red circle highlights the cell (1,6). A blue diagonal line runs from (1,1) to (6,6). Handwritten fractions: 1/6 next to the first row, 1/6 next to the first column, and 1/36 next to the cell (1,6).

# Chain rules

# Chain Rule

DEFINITION:

$$\mathbb{P}(A, B) = \mathbb{P}(A) \cdot \mathbb{P}(B|A)$$

$$\mathbb{P}(A, B, C) = \text{?}$$

Recall the definition of cond. prob.:

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

# Chain Rule

DEFINITION:

$$\mathbb{P}(A, B) = \mathbb{P}(A) \cdot \mathbb{P}(B|A)$$

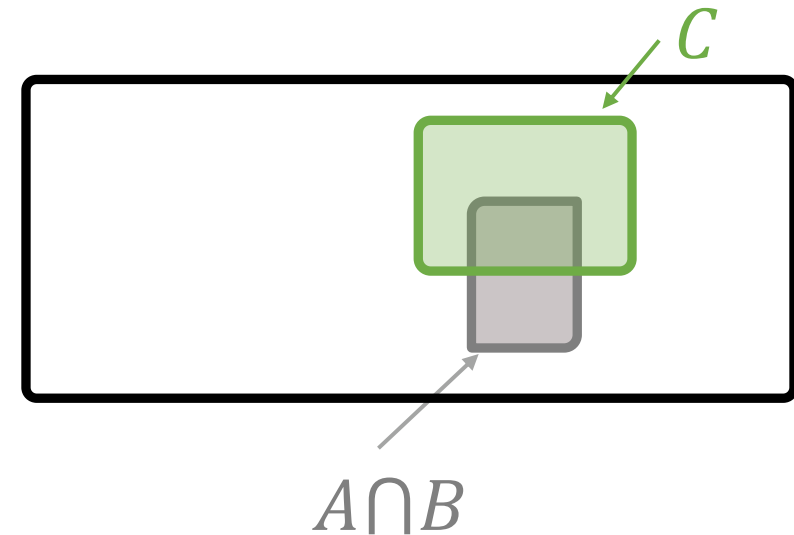
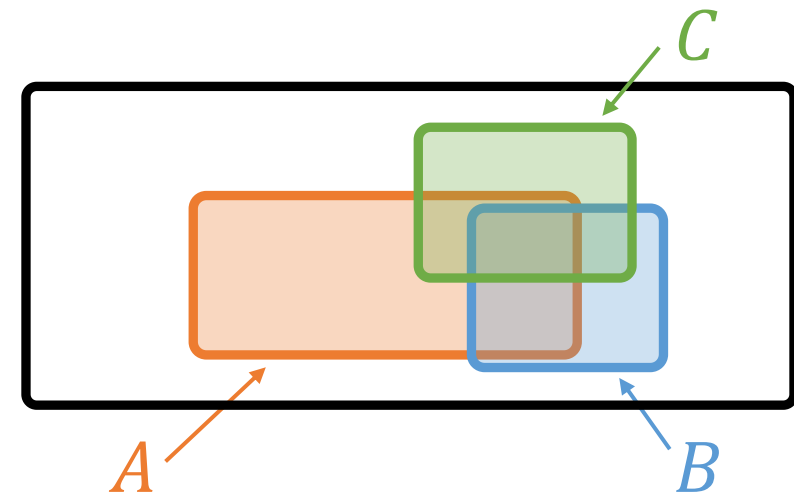
$$\mathbb{P}(\underbrace{A, B}_E, C) = \mathbb{P}(\underbrace{A, B}_E) \cdot \mathbb{P}(C|\underbrace{A, B}_E)$$

*Just treat  $A \cap B$  as a new event  $E$*

$$\mathbb{P}(A, B, C) = \mathbb{P}(A) \cdot \mathbb{P}(B|A) \cdot \mathbb{P}(C|A, B)$$

Recall the definition of cond. prob.:

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$





# Bayes law

# Bayes Law

## DEFINITION

Conditional probability  $\mathbb{P}(A|B)$  is the probability that event  $A$  occurred, given that event  $B$  occurred (defined if  $\mathbb{P}(B) \neq 0$ )

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

$$\mathbb{P}(B|A) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(A)}$$

$$\mathbb{P}(A, B) = \mathbb{P}(B|A) \cdot \mathbb{P}(A)$$

$$= \mathbb{P}(A|B) \cdot \mathbb{P}(B)$$

## BAYES LAW

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(B|A) \cdot \mathbb{P}(A)}{\mathbb{P}(B)}$$

# Bayes Law

## DEFINITION

Conditional probability  $\mathbb{P}(A|B)$  is the probability that event  $A$  occurred, given that event  $B$  occurred (defined if  $\mathbb{P}(B) \neq 0$ )

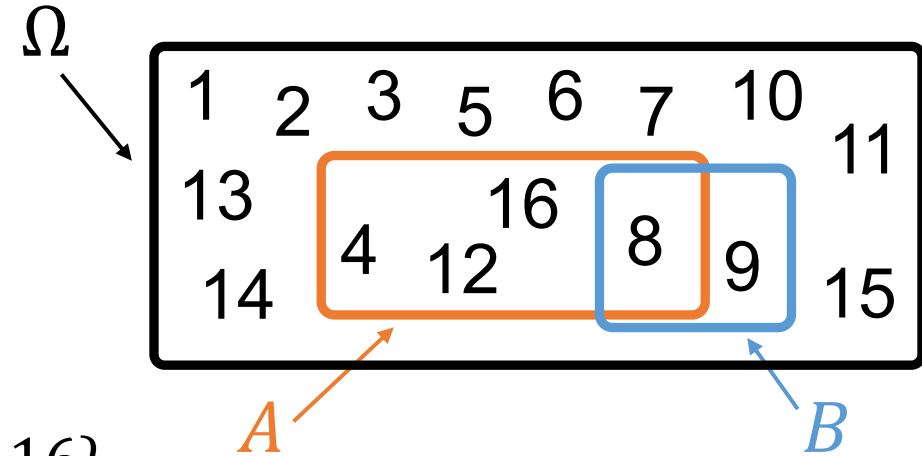
$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

$$\mathbb{P}(B|A) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(A)}$$

$$\begin{aligned}\mathbb{P}(A, B) &= \mathbb{P}(B|A) \cdot \mathbb{P}(A) \\ &= \mathbb{P}(A|B) \cdot \mathbb{P}(B)\end{aligned}$$

## EXAMPLE 4:

$$\Omega = \{1, 2, \dots, 16\}$$



## BAYES LAW

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(B|A) \cdot \mathbb{P}(A)}{\mathbb{P}(B)}$$

?

# Bayes Law

## DEFINITION

Conditional probability  $\mathbb{P}(A|B)$  is the probability that event  $A$  occurred, given that event  $B$  occurred (defined if  $\mathbb{P}(B) \neq 0$ )

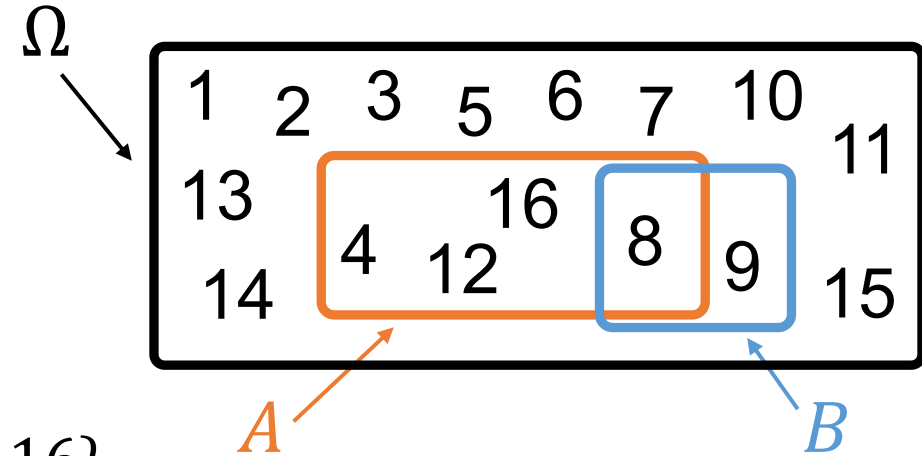
$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(B)}$$

$$\mathbb{P}(B|A) = \frac{\mathbb{P}(A, B)}{\mathbb{P}(A)}$$

$$\begin{aligned}\mathbb{P}(A, B) &= \mathbb{P}(B|A) \cdot \mathbb{P}(A) = \frac{1}{4} \cdot \frac{1}{4} = \frac{1}{16} \\ &= \mathbb{P}(A|B) \cdot \mathbb{P}(B) = \frac{1}{2} \cdot \frac{1}{8} = \frac{1}{16}\end{aligned}$$

## EXAMPLE 4:

$$\Omega = \{1, 2, \dots, 16\}$$



## BAYES LAW

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(B|A) \cdot \mathbb{P}(A)}{\mathbb{P}(B)}$$

# Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k)
- accuracy of blood test is 99%
- A patient has a positive test.

What is the chance they have Zika?



		blood test positive ( $B$ )	
		$B$	$\bar{B}$
patient has Zika ( $Z$ )	$Z$	$\mathbb{P}(B Z) = 0.99$	$\mathbb{P}(\bar{B} Z) = 0.01$
	$\bar{Z}$	$\mathbb{P}(B \bar{Z}) = 0.01$	$\mathbb{P}(\bar{B} \bar{Z}) = 0.99$

# Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k) = **base rate** =  $\mathbb{P}(Z)$

- accuracy of blood test is 99%

- A patient has a positive test.

**What is the chance they have Zika?**



patient has  $Z$   
Zika ( $Z$ )

blood test positive ( $B$ )	
$B$	$\bar{B}$
$\mathbb{P}(B Z) = 0.99$	$\mathbb{P}(\bar{B} Z) = 0.01$
$\mathbb{P}(B \bar{Z}) = 0.01$	$\mathbb{P}(\bar{B} \bar{Z}) = 0.99$

**false negative**

**false positive**

We don't want  $\mathbb{P}(B|Z) = 0.99$ , but rather  $\mathbb{P}(Z|B)$

# Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k) = base rate =  $\mathbb{P}(Z)$
- accuracy of blood test is 99%
- A patient has a positive test.  
**What is the chance they have Zika?**

		blood test positive ( $B$ )	
		$B$	$\bar{B}$
patient has Zika ( $Z$ )	$Z$	$\mathbb{P}(B Z) = 0.99$	$\mathbb{P}(\bar{B} Z) = 0.01$
	$\bar{Z}$	$\mathbb{P}(B \bar{Z}) = 0.01$	$\mathbb{P}(\bar{B} \bar{Z}) = 0.99$

We don't want  $\mathbb{P}(B|Z) = 0.99$ , but rather  $\mathbb{P}(Z|B)$

$$\mathbb{P}(Z|B) = \frac{\mathbb{P}(B, Z)}{\mathbb{P}(B)} = \frac{\mathbb{P}(B|Z) \cdot \mathbb{P}(Z)}{?}$$

# Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

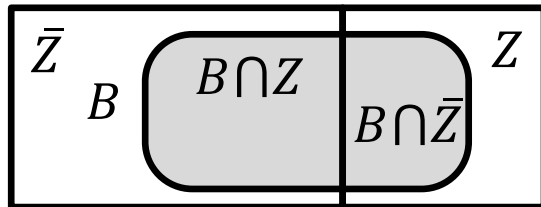
- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k) = base rate =  $\mathbb{P}(Z)$
- accuracy of blood test is 99%
- A patient has a positive test.  
What is the chance they have Zika?

		blood test positive (B)	
		$B$	$\bar{B}$
patient has Zika (Z)	$Z$	$\mathbb{P}(B Z) = 0.99$	$\mathbb{P}(\bar{B} Z) = 0.01$
	$\bar{Z}$	$\mathbb{P}(B \bar{Z}) = 0.01$	$\mathbb{P}(\bar{B} \bar{Z}) = 0.99$

We don't want  $\mathbb{P}(B|Z) = 0.99$ , but rather  $\mathbb{P}(Z|B)$

$$\mathbb{P}(Z|B) = \frac{\mathbb{P}(B, Z)}{\mathbb{P}(B)} = \frac{\mathbb{P}(B|Z) \cdot \mathbb{P}(Z)}{\mathbb{P}(B, Z) + \mathbb{P}(B, \bar{Z})}$$

total probability theorem



$\mathbb{P}(B|Z) \cdot \mathbb{P}(Z)$  ?



# Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k) = base rate =  $\mathbb{P}(Z)$
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We don't want  $\mathbb{P}(B|Z) = 0.99$ , but rather  $\mathbb{P}(Z|B)$

$$\mathbb{P}(Z|B) = \frac{\mathbb{P}(B, Z)}{\mathbb{P}(B)} = \frac{\mathbb{P}(B|Z) \cdot \mathbb{P}(Z)}{\mathbb{P}(B, Z) + \mathbb{P}(B, \bar{Z})} = \frac{0.99 \cdot 10^{-5}}{0.99 \cdot 10^{-5} + 0.01 \cdot (1 - 10^{-5})}$$

*total probability theorem*

$\mathbb{P}(B|Z) \cdot \mathbb{P}(Z)$  (points to  $\mathbb{P}(B, Z)$ )

$\mathbb{P}(B|\bar{Z}) \cdot \mathbb{P}(\bar{Z})$  (points to  $\mathbb{P}(B, \bar{Z})$ )

$$\approx \frac{10^{-5}}{10^{-5} + 0.01} \approx \frac{10^{-5}}{10^{-2}} = 10^{-3} = 0.1\%$$

$10^{-2}$  (points to 0.01)

# Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k) = base rate =  $\mathbb{P}(Z)$
- accuracy of blood test is 99%
- Assume 10M people in FL. **?**  
**Complete the numbers!**

		blood test positive ( $B$ )	
		$B$	$\bar{B}$
patient has Zika ( $Z$ )	$Z$	?	?
	$\bar{Z}$	?	?

$$\mathbb{P}(Z|B) \approx 0.1\%$$

# Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k) = **base rate** =  $\mathbb{P}(Z)$
- accuracy of blood test is 99%
- Assume 10M people in FL.  
**Complete the numbers!**

		blood test positive ( $B$ )		
		$B$	$\bar{B}$	
patient has Zika ( $Z$ )	$Z$			100
	$\bar{Z}$			9,999,900

$$\mathbb{P}(Z|B) \approx 0.1\%$$

# Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k) = base rate =  $\mathbb{P}(Z)$
- accuracy of blood test is 99%

- Assume 10M people in FL.  
**Complete the numbers!**

		blood test positive (B)		
		$B$	$\bar{B}$	
patient has Zika (Z)	$Z$	99	1	100
	$\bar{Z}$			9,999,900

$$\mathbb{P}(Z|B) \approx 0.1\%$$

# Bayes Law ("Base rate fallacy problem")



## EXAMPLE 5: Zika in Florida in 2016

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- Assume 10M people in FL.  
**Complete the numbers!**

		blood test positive (B)		
		$B$	$\bar{B}$	
patient has Zika (Z)	$Z$	99	1	100
	$\bar{Z}$	99,999	9,899,901	9,999,900

$$\mathbb{P}(Z|B) \approx 0.1\%$$

# Bayes Law ("Base rate fallacy problem")



## EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k) = base rate =  $\mathbb{P}(Z)$
- accuracy of blood test is 99%
- Assume 10M people in FL. Complete the numbers!

		blood test positive ( $B$ )		
		$B$	$\bar{B}$	
patient has Zika ( $Z$ )	$Z$	99	1	100
	$\bar{Z}$	99,999	9,899,901	9,999,900
		100,098	9,899,902	

$\mathbb{P}(Z|B) \approx 0.1\%$   
probability that:  
a positive test is correct

# Bayes Law ("Base rate fallacy problem")



## EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k) = **base rate** =  $\mathbb{P}(Z)$
- accuracy of blood test is 99%
- Assume 10M people in FL. Complete the numbers!

$\mathbb{P}(B|Z) = 99\%$  patient has  $Z$   
 $\mathbb{P}(\bar{B}|\bar{Z}) = 99\%$  Zika ( $Z$ )  $\bar{Z}$

blood test positive ( $B$ )		
$B$	$\bar{B}$	
99	1	100
99,999	9,899,901	9,999,900
100,098	9,899,902	

$\mathbb{P}(B|Z) \cdot \mathbb{P}(Z) + \mathbb{P}(\bar{B}|\bar{Z}) \cdot \mathbb{P}(\bar{Z}) = 99\%$   
probability that a random  
test is correct

$\neq$

$\mathbb{P}(Z|B) \approx 0.1\%$   
probability that:  
a positive test is correct

# Bayes Law ("Base rate fallacy problem")



EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k) = **base rate** =  $\mathbb{P}(Z)$
- accuracy of blood test is 99%
- Assume 10M people in FL. Complete the numbers!

		blood test positive ( $B$ )		
		$B$	$\bar{B}$	
patient has Zika ( $Z$ )	$Z$	99	1	100
	$\bar{Z}$	99,999	9,899,901	9,999,900
		100,098	9,899,902	

$\mathbb{P}(B|Z) \cdot \mathbb{P}(Z) + \mathbb{P}(\bar{B}|\bar{Z}) \cdot \mathbb{P}(\bar{Z}) = 99\%$   
 probability that a random test is correct

$\neq$

$\mathbb{P}(Z|B) \approx 0.1\%$   
 probability that: a positive test is correct

  $\mathbb{P}(Z) = 10^{-5}$





# Bayes Law ("Base rate fallacy problem")

## EXAMPLE 5: Zika in Florida in 2016

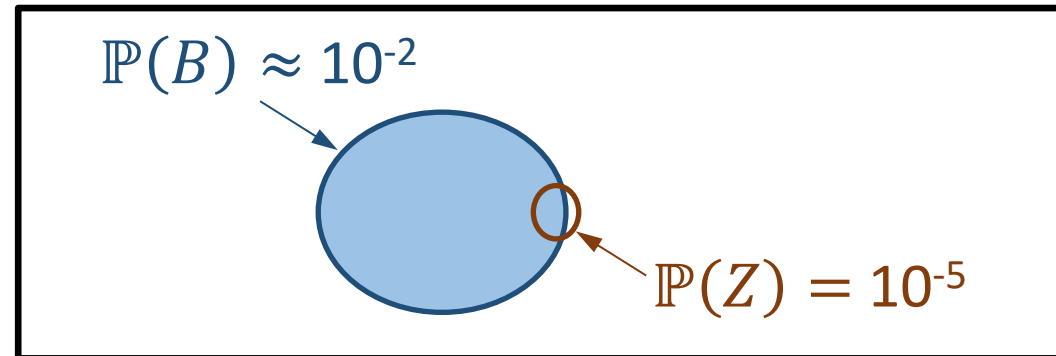
- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k) = **base rate** =  $\mathbb{P}(Z)$
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- Assume 10M people in FL. Complete the numbers!

		blood test positive ( $B$ )		
		$B$	$\bar{B}$	
patient has Zika ( $Z$ )	$Z$	99	1	100
	$\bar{Z}$	99,999	9,899,901	9,999,900
		100,098	9,899,902	

$\mathbb{P}(B|Z) \cdot \mathbb{P}(Z) + \mathbb{P}(\bar{B}|\bar{Z}) \cdot \mathbb{P}(\bar{Z}) = 99\%$   
probability that a random  
test is correct

$\neq$

$\mathbb{P}(Z|B) \approx 0.1\%$   
probability that:  
a positive test is correct





# Bayes Law ("Base rate fallacy problem")

EXAMPLE 5: Zika in Florida in 2016

- prevalence of Zika ( $Z$ ) in Florida is  $10^{-5}$  (1 in 100k) = **base rate** =  $\mathbb{P}(Z)$
- accuracy of blood test is 99%

- Assume 10M people in FL.  
**Complete the numbers**

$$\mathbb{P}(B|Z) = 99\%$$
$$\mathbb{P}(\bar{B}|\bar{Z}) = 99\%$$

patient has  $Z$   
Zika ( $Z$ )  $\bar{Z}$

blood test positive ( $B$ )		
$B$	$\bar{B}$	
99	1	100
99,999	9,899,901	9,999,900
100,098	9,899,902	

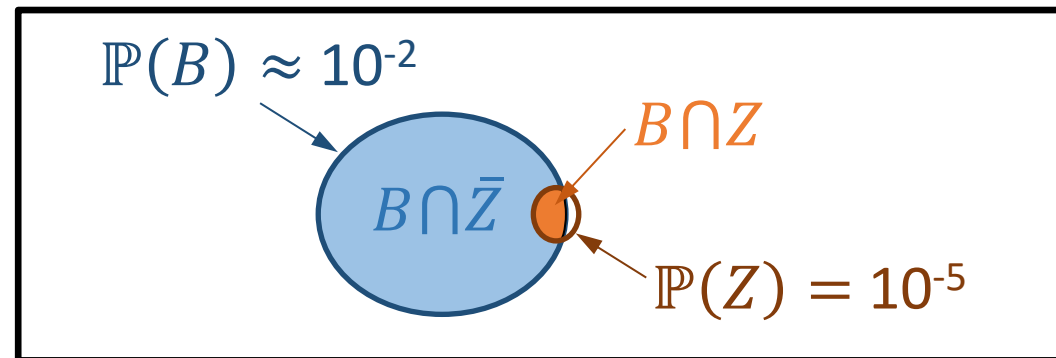
$$\mathbb{P}(B|Z) \cdot \mathbb{P}(Z) + \mathbb{P}(\bar{B}|\bar{Z}) \cdot \mathbb{P}(\bar{Z}) = 99\%$$

probability that a random  
test is correct

$\neq$

$$\mathbb{P}(Z|B) = 0.1\%$$

probability that:  
a positive test is correct

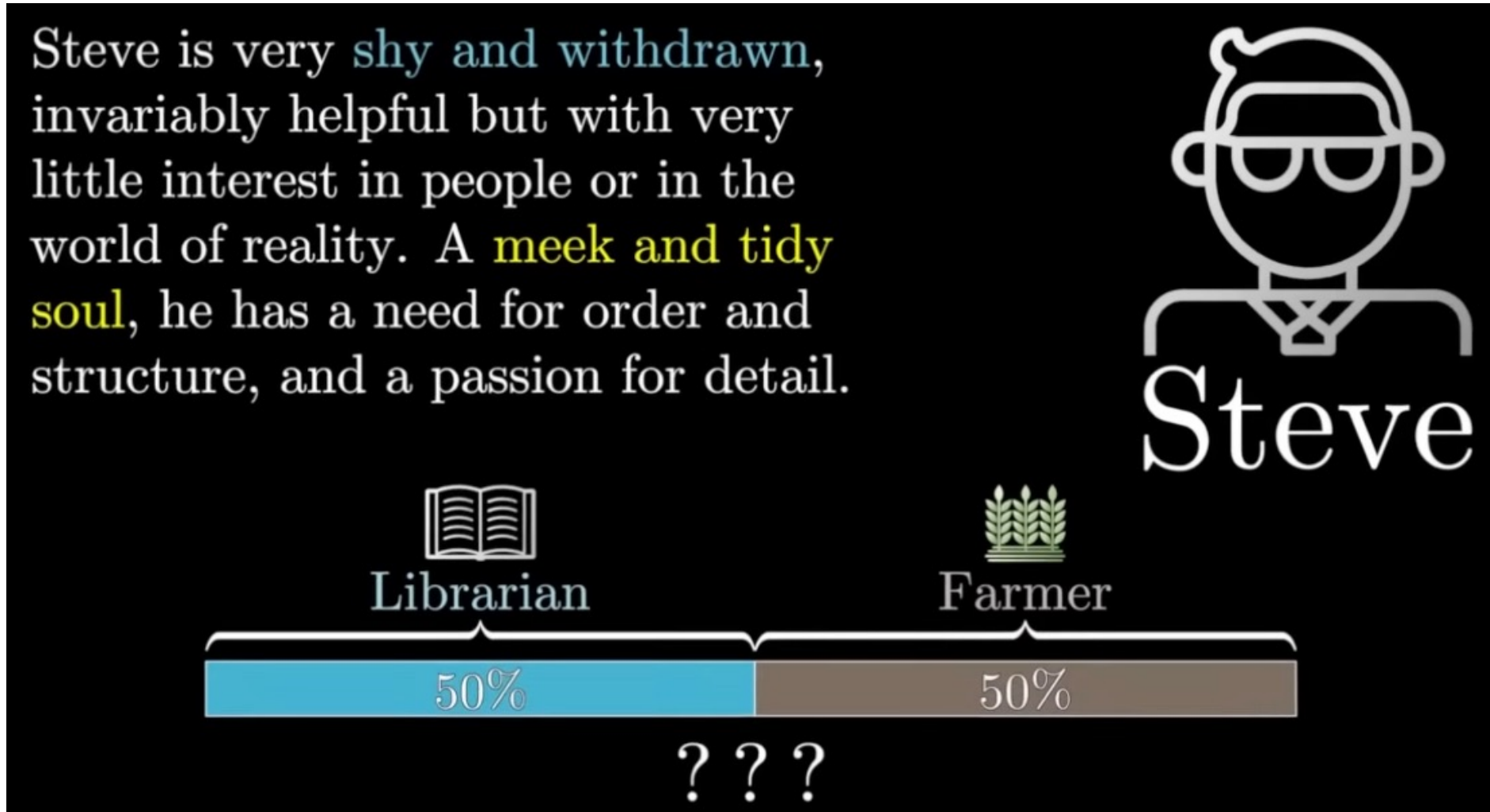


# An unrelated (?) question:

Steve is very shy and withdrawn, invariably helpful but with very little interest in people or in the world of reality. A meek and tidy soul, he has a need for order and structure, and a passion for detail.



# An unrelated (?) question:



# Bayes theorem illustrated by 3Blue1Brown

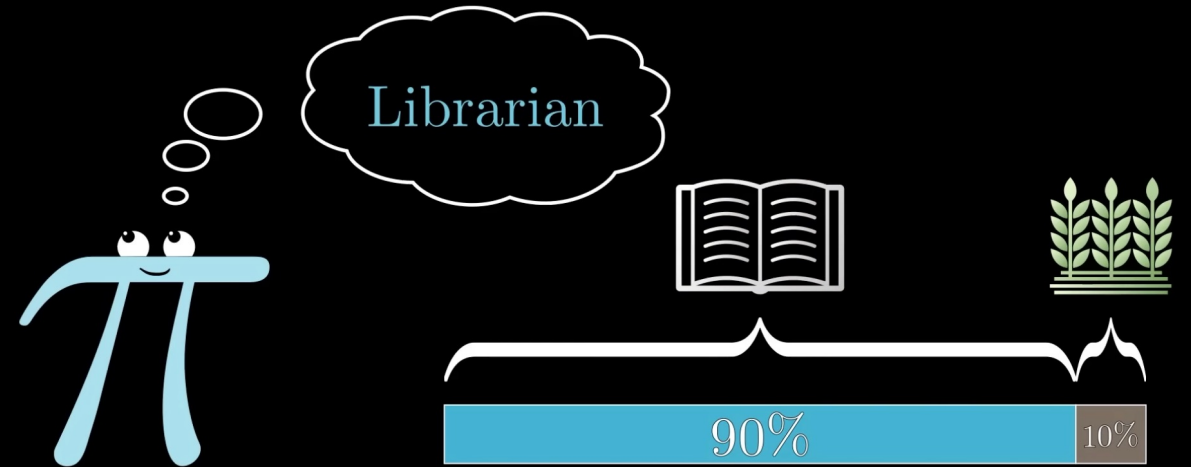


Daniel  
Kahneman

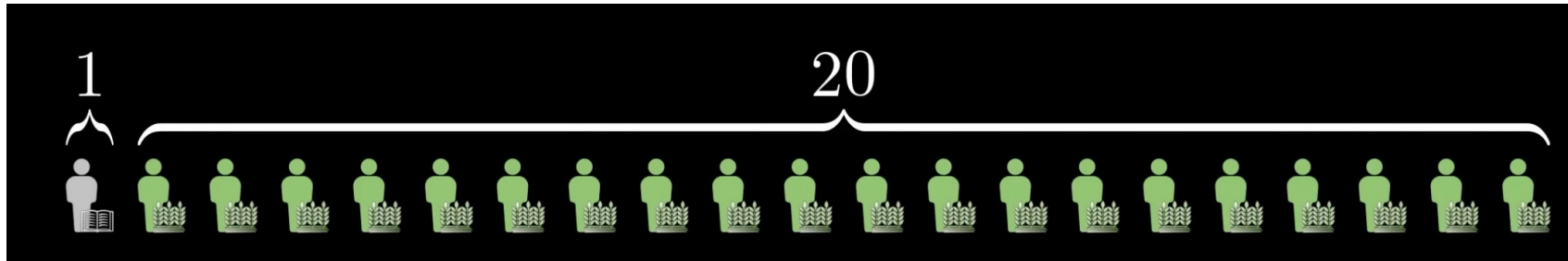


Amos  
Tversky

Steve is very shy and withdrawn, invariably helpful but with very little interest in people or in the world of reality. A meek and tidy soul, he has a need for order and structure, and a passion for detail.

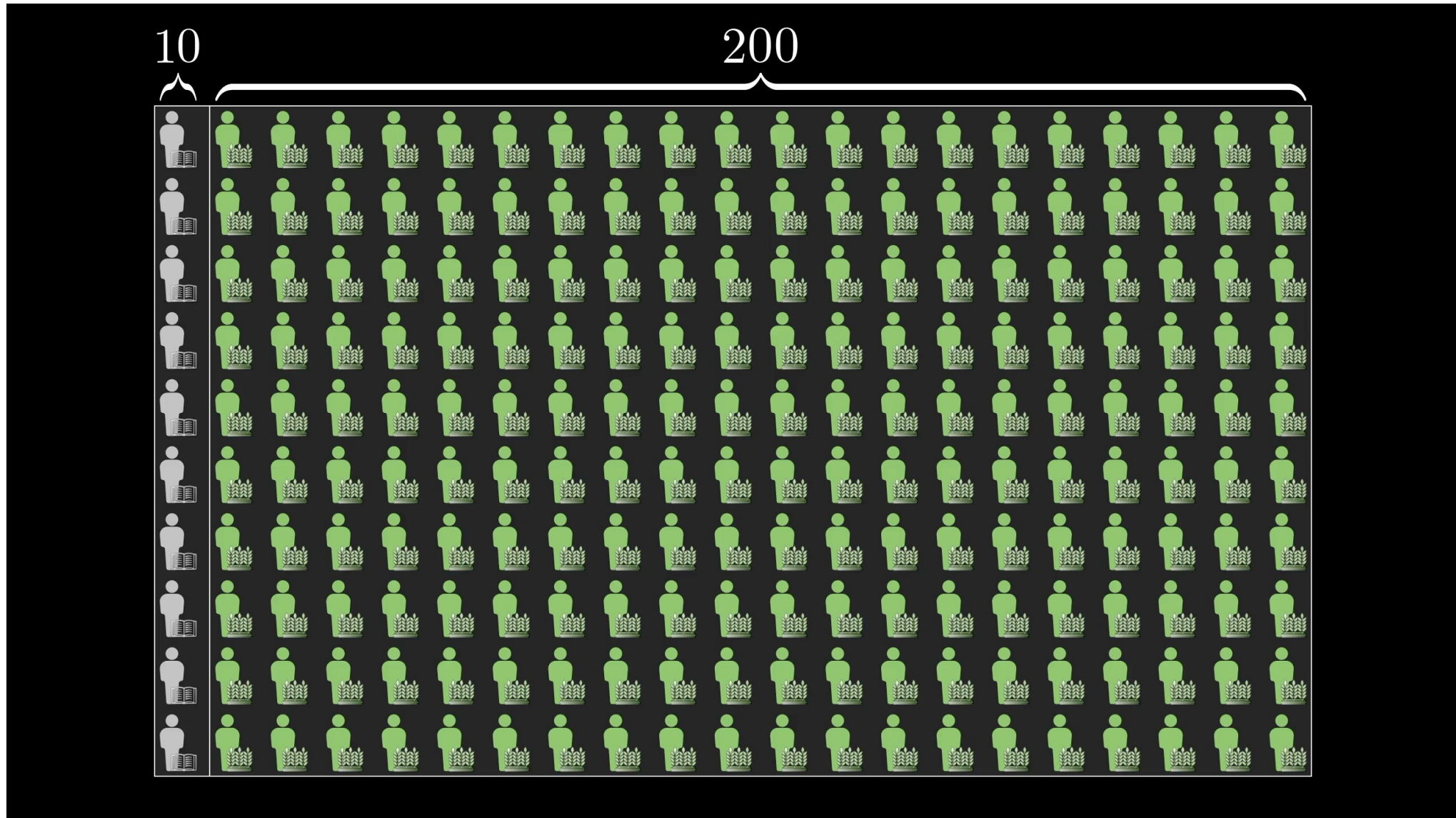


# Bayes theorem illustrated by 3Blue1Brown

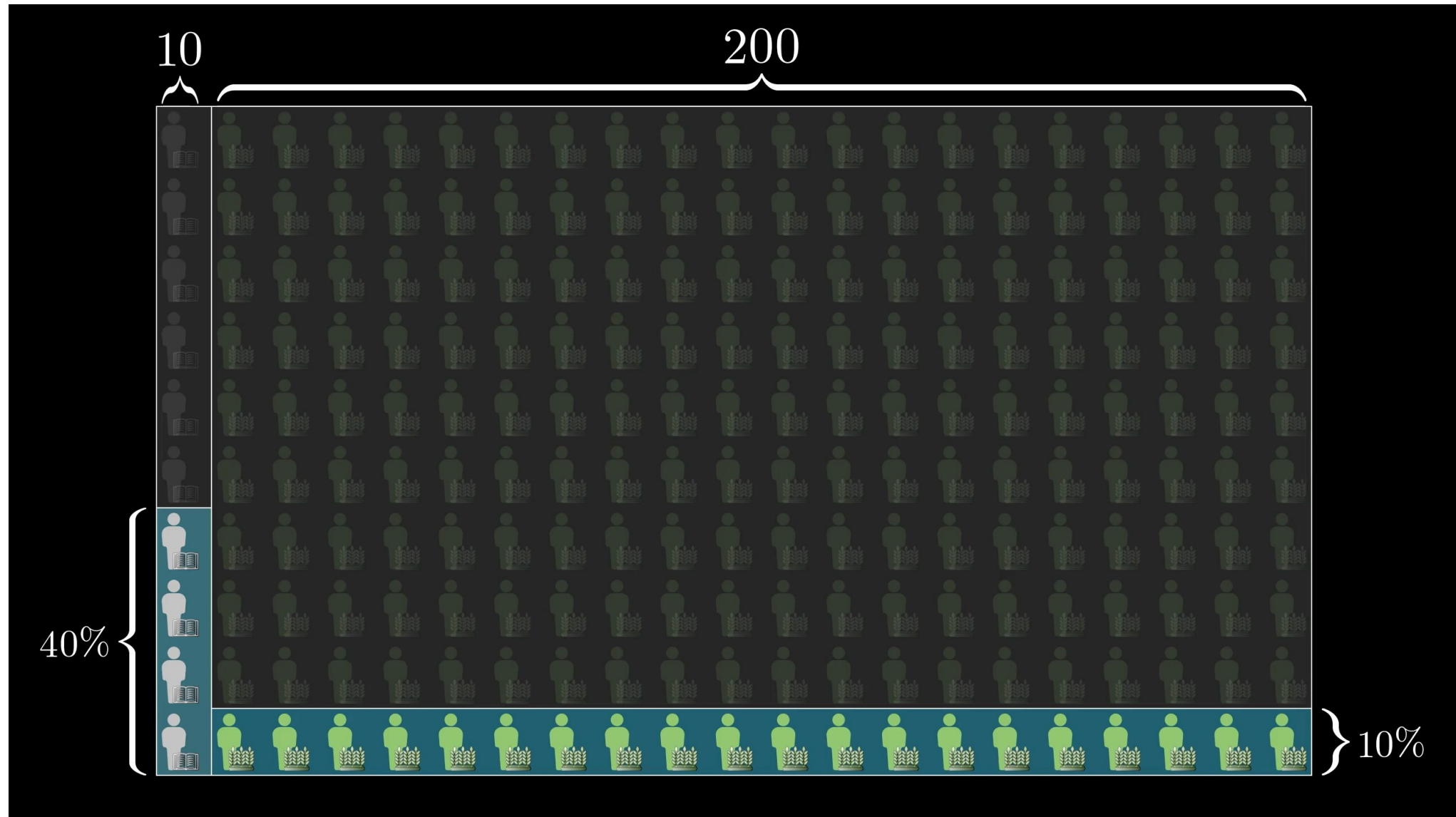




# Bayes theorem illustrated by 3Blue1Brown



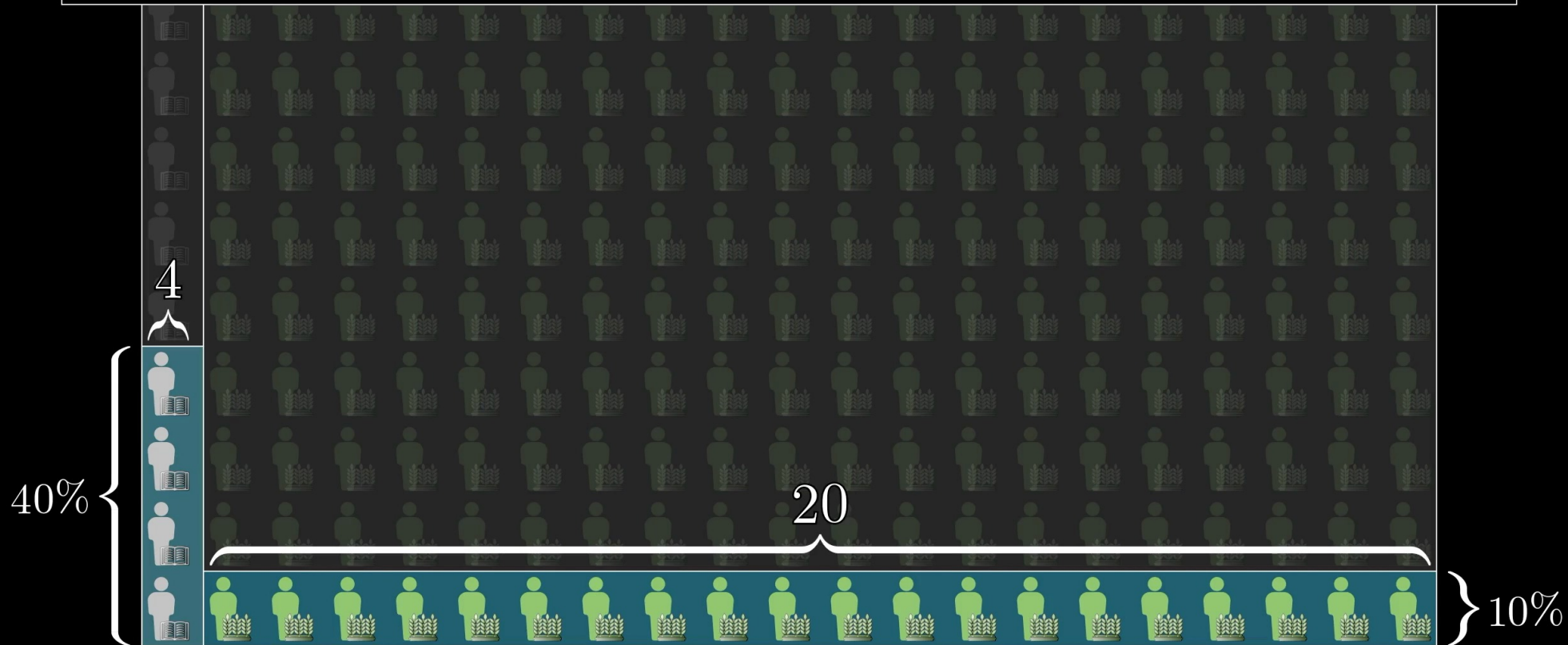
# Bayes theorem illustrated by 3Blue1Brown





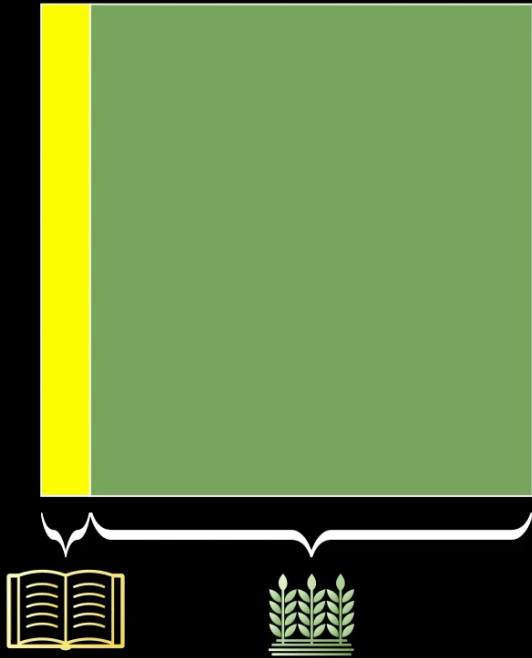
# Bayes theorem illustrated by 3Blue1Brown

$$P(\text{Librarian given description}) = \frac{4}{4 + 20} \approx 16.7\%$$

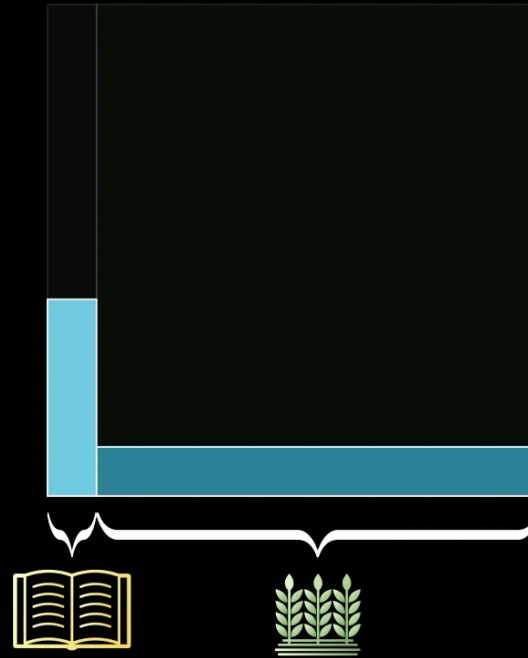


# Bayes theorem illustrated by 3Blue1Brown

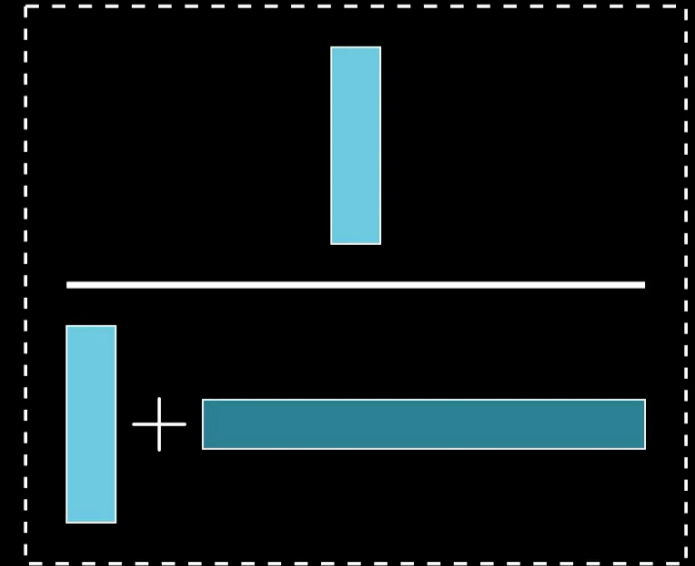
All possibilities



All possibilities  
fitting the evidence



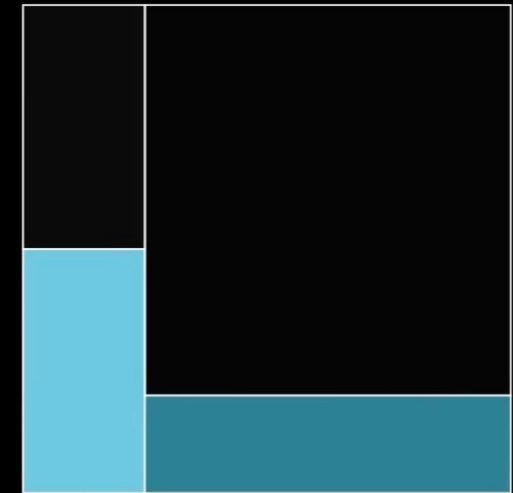
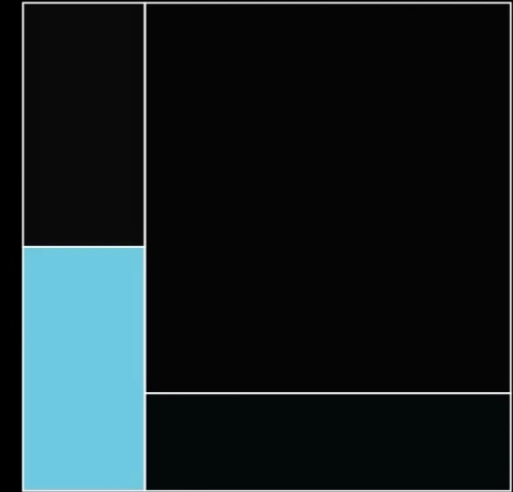
$P$  ( Librarian given  
the evidence )



# Bayes theorem illustrated by 3Blue1Brown

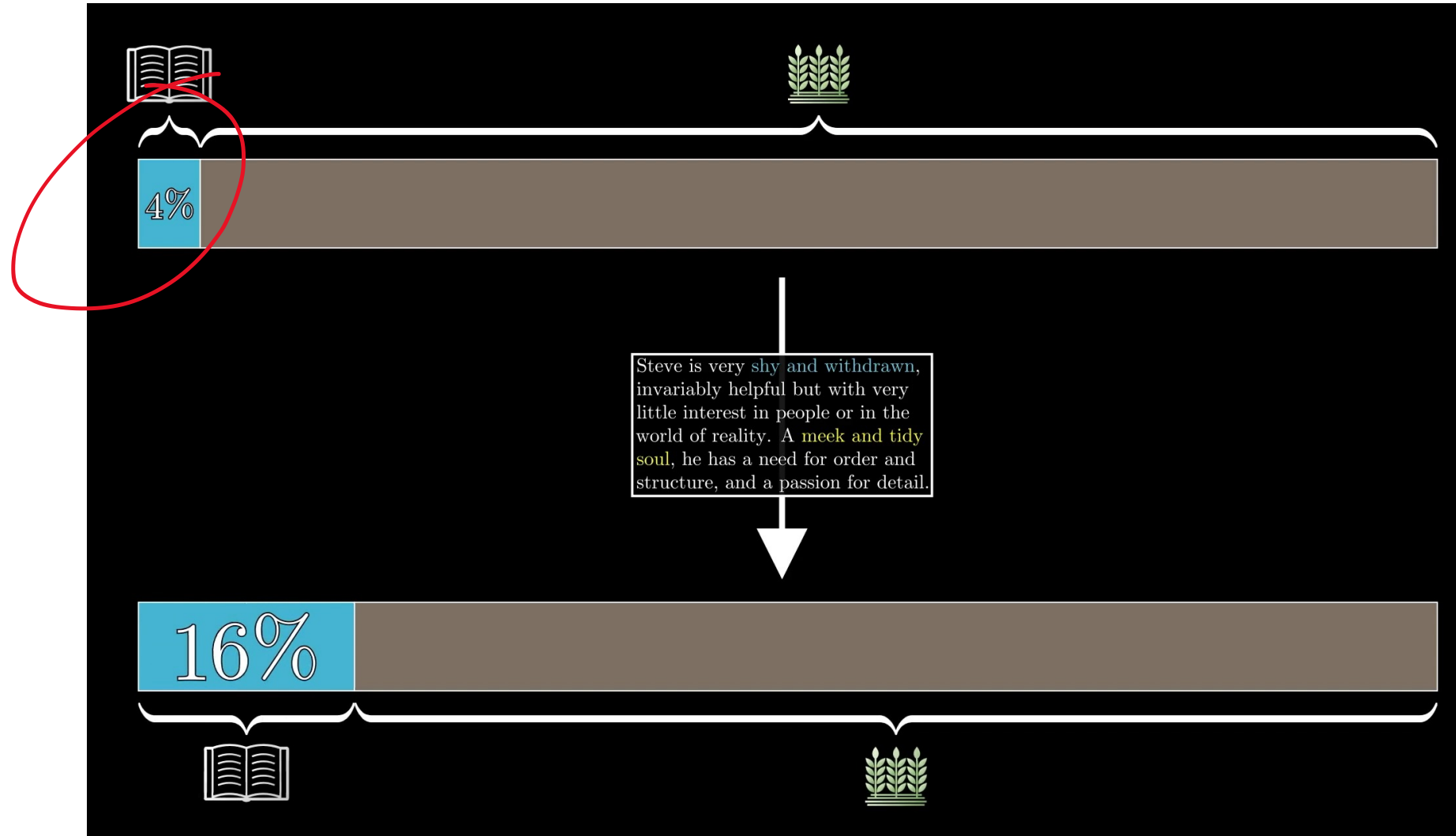
## Bayes' theorem

$$P(H|E) = \frac{P(H)P(E|H)}{P(E)} = \frac{\text{Top Diagram}}{\text{Bottom Diagram}}$$



# Bayes theorem illustrated by 3Blue1Brown

New evidence updates prior beliefs! Evidence does not exist in a vacuum



# Part 1: Information Theory Basics

## L03: Basics of Probability (2/2)

[Random experiment, independence, conditional probability, chain rule, Bayes' theorem, random variables]

Wolfgang Gatterbauer

cs7840 Foundations and Applications of Information Theory (fa25)

<https://northeastern-datalab.github.io/cs7840/fa25/>

9/15/2025

# Pre-class conversations

- Last class recapitulation
- Suggestions on small class projects = "scribes" (e.g. parallelizing compression)
- New class arrivals
- Questions
  
- Today:
  - The basics of probability theory
  - Start of information theory basics

# Random variable, expectation

# Random variable: pmf $p_X(x)$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

Two types:

1. numerical

e.g.  $X(\omega)$  lottery win

2. indicator

e.g.  $X(\omega) = \begin{cases} 1 & \text{if win} \\ 0 & \text{if not} \end{cases}$

$X$  is called a "random variable" (RV) because it depends on the outcome of a random experiment. But the mapping  $X: \Omega \rightarrow \mathbb{R}$  is deterministic.

The underlying probability measure  $\mathbb{P}: \Sigma \rightarrow [0, 1]$  induces a pmf  $p_X$  (probability mass function) over the range of the RV  $X: p_X: \mathbb{R} \rightarrow [0, 1]$

$$p_X(x) = \mathbb{P}(\{X = x\})$$

also written as  $p(x) = \mathbb{P}(X = x)$

EXAMPLE 3 (CONT.):

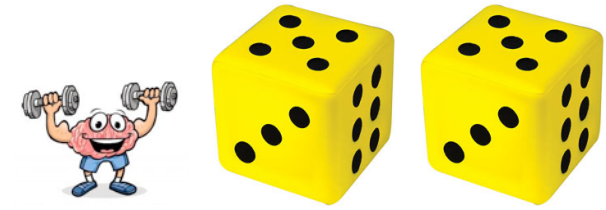
roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let  $X$  be the sum of the rolls.

Then  $X$  is a RV.

What is the pmf?



$\Omega$		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)



# Random variable: pmf $p_X(x)$

Random variable (RV)

$$X: \Omega \rightarrow \mathbb{R}$$

Two types:

1. numerical

e.g.  $X(\omega)$  lottery win

2. indicator

$$\text{e.g. } X(\omega) = \begin{cases} 1 & \text{if win} \\ 0 & \text{if not} \end{cases}$$

$X$  is called a "random variable" (RV) because it depends on the outcome of a random experiment. But the mapping  $X: \Omega \rightarrow \mathbb{R}$  is deterministic.

The underlying probability measure  $\mathbb{P}: \Sigma \rightarrow [0, 1]$  induces a pmf  $p_X$  (probability mass function) over the range of the RV  $X: p_X: \mathbb{R} \rightarrow [0, 1]$

$$p_X(x) = \mathbb{P}(\{X = x\})$$

also written as  $p(x) = \mathbb{P}(X = x)$

EXAMPLE 3 (CONT.):

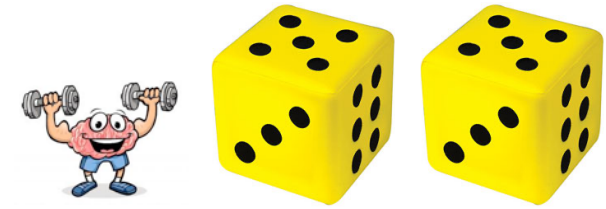
roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let  $X$  be the sum of the rolls.

Then  $X$  is a RV.

What is the pmf?



		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)
		X					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
	6	7	8	9	10	11	12

# Random variable: pmf $p_X(x)$

Random variable (RV)

$$X: \Omega \rightarrow \mathbb{R}$$

Two types:

1. numerical

e.g.  $X(\omega)$  lottery win

2. indicator

$$\text{e.g. } X(\omega) = \begin{cases} 1 & \text{if win} \\ 0 & \text{if not} \end{cases}$$

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EXAMPLE 3 (CONT.):

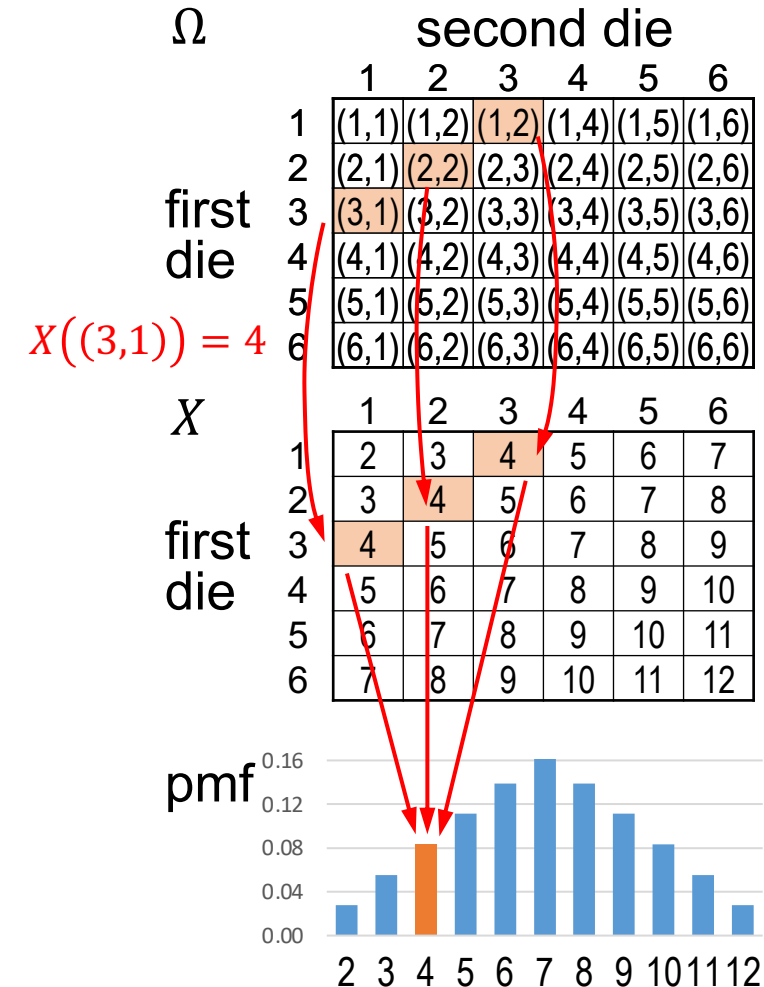
roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let  $X$  be the sum of the rolls.

Then  $X$  is a RV.

What is the pmf?



# Random variable: expectation $\mathbb{E}$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

Expectation: a weighted average  
(in proportion to the probabilities)  
of the possible values of  $X$

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

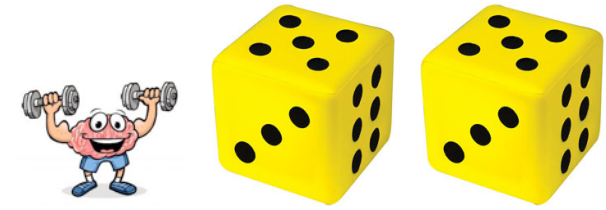
$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$

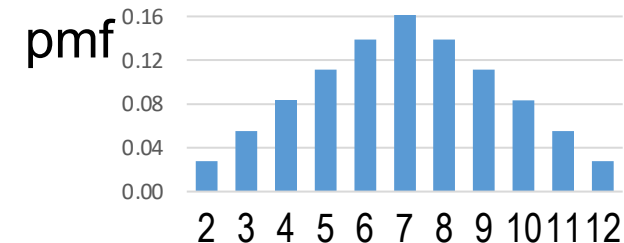
Let  $X$  be the sum of the rolls.



What is  $\mathbb{E}[X]$  ?

		$\Omega$ second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		$X$					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
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# Random variable: expectation $\mathbb{E}$

Random variable (RV)

$$X: \Omega \rightarrow \mathbb{R}$$

Expectation: a weighted average  
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of the possible values of  $X$

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let  $X$  be the sum of the rolls.



What is  $\mathbb{E}[X]$ ?

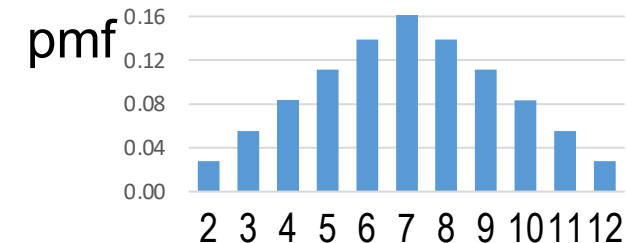
Variant 1:

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

?

		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
	6	7	8	9	10	11	12



# Random variable: expectation $\mathbb{E}$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

Expectation: a weighted average  
(in proportion to the probabilities)  
of the possible values of  $X$

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

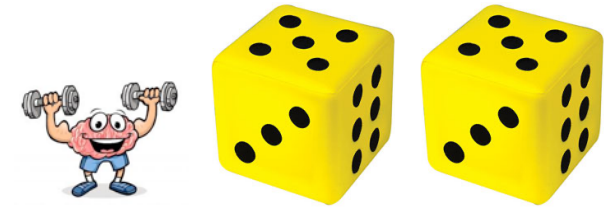
$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$

Let  $X$  be the sum of the rolls.



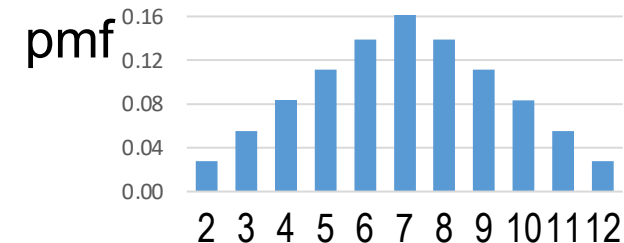
What is  $\mathbb{E}[X]$ ?

Variant 1:

$$\begin{aligned} \mathbb{E}[X] &= \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\}) \\ &= \frac{1}{36} \cdot \sum_{\omega \in \Omega} X(\omega) \\ &= \frac{1}{36} \cdot (\text{"sum of table"}) \\ &= 7 \end{aligned}$$

		$\Omega$ second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		$X$					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
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# Random variable: expectation $\mathbb{E}$

Random variable (RV)

$$X: \Omega \rightarrow \mathbb{R}$$

Expectation: a weighted average  
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$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

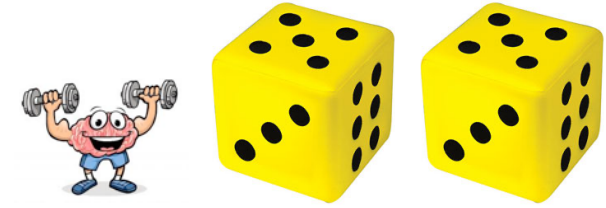
$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

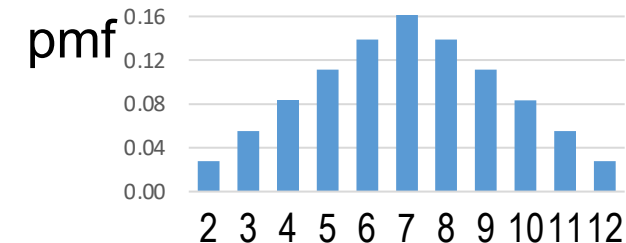
$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let  $X$  be the sum of the rolls.



		$\Omega$ second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		$X$					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
	6	7	8	9	10	11	12



What is  $\mathbb{E}[X]$ ?

Variant 2:

$$\mathbb{E}[X] = \sum_x x \cdot \mathbb{P}(X = x)$$

?

# Random variable: expectation $\mathbb{E}$

Random variable (RV)

$$X: \Omega \rightarrow \mathbb{R}$$

Expectation: a weighted average  
(in proportion to the probabilities)  
of the possible values of  $X$

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

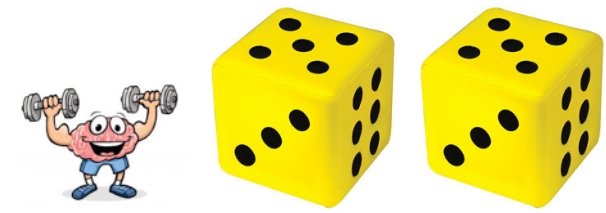
$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let  $X$  be the sum of the rolls.



What is  $\mathbb{E}[X]$ ?

Variant 2:

$$\mathbb{E}[X] = \sum_x x \cdot \mathbb{P}(X = x)$$

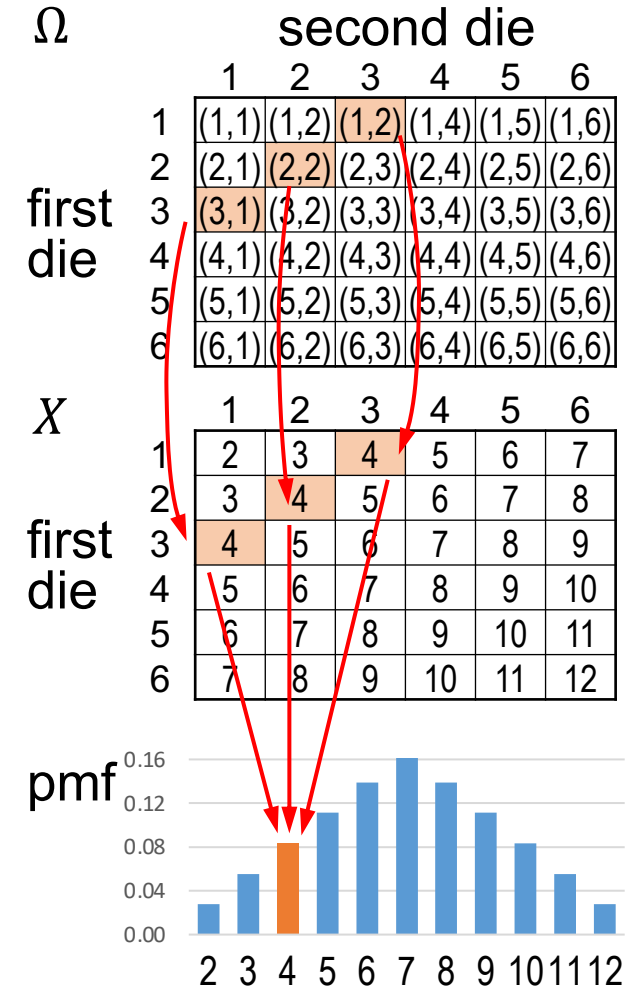
$$X = 2 \quad p_X(2) = \frac{1}{36}$$

$$X = 3 \quad p_X(3) = \frac{2}{36}$$

...

$$X = 12 \quad p_X(12) = \frac{1}{36}$$

$$\mathbb{E}[X] = 2 \cdot \frac{1}{36} + 3 \cdot \frac{2}{36} + \dots + 12 \cdot \frac{1}{36} = 7$$



# Random variable: expectation $\mathbb{E}$

Random variable (RV)

$$X: \Omega \rightarrow \mathbb{R}$$

Expectation: a weighted average  
(in proportion to the probabilities)  
of the possible values of  $X$

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let  $X$  be the sum of the rolls.

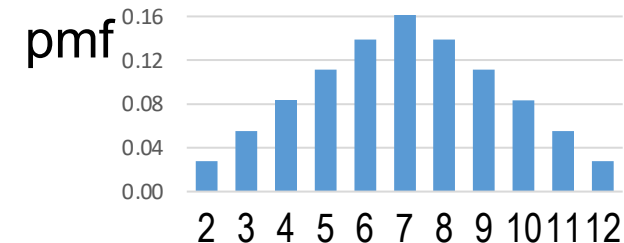


What is  $\mathbb{E}[X]$ ?

Variant 3: *RVs for the outcome of the first and second die rolls*

$$\mathbb{E}[X] = \mathbb{E}[X_1 + X_2] = ?$$

		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)
		X					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
	6	7	8	9	10	11	12





# Random variable: expectation $\mathbb{E}$

Random variable (RV)

$$X: \Omega \rightarrow \mathbb{R}$$

Expectation: a weighted average  
(in proportion to the probabilities)  
of the possible values of  $X$

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$$

Let  $X$  be the sum of the rolls.



What is  $\mathbb{E}[X]$ ?

Variant 3: *RVs for the outcome of the first and second die rolls*

$$\mathbb{E}[X] = \mathbb{E}[X_1 + X_2] = \mathbb{E}[X_1] + \mathbb{E}[X_2]$$

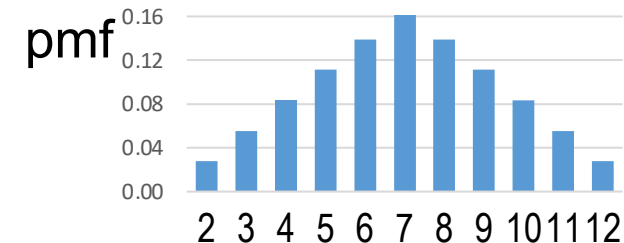
*linearity of expectation!*

$$\mathbb{E}[X_1] = \mathbb{E}[X_2] = 3.5$$

$$\mathbb{E}[X] = 3.5 + 3.5 = 7$$

$\Omega$		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,2)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		$X$					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
	6	7	8	9	10	11	12



# Random variable: expectation $\mathbb{E}$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

Expectation: a weighted average  
(in proportion to the probabilities)  
of the possible values of  $X$

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$

Let  $Y$  be the product of the rolls.



What is  $\mathbb{E}[Y]$  ?

		$\Omega$ second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)
		$Y$					
		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
	2	2	4	6	8	10	12
	3	3	6	9	12	15	18
	4	4	8	12	16	20	24
	5	5	10	15	20	25	30
	6	6	12	18	24	30	36

# Random variable: expectation $\mathbb{E}$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

Expectation: a weighted average  
(in proportion to the probabilities)  
of the possible values of  $X$

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$

Let  $Y$  be the product of the rolls.

What is  $\mathbb{E}[Y]$ ?

$$\mathbb{E}[Y] = \mathbb{E}[X_1 \cdot X_2] = \mathbb{E}[X_1] \cdot \mathbb{E}[X_2]$$

?



		$\Omega$ second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)
		$Y$					
		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
	2	2	4	6	8	10	12
	3	3	6	9	12	15	18
	4	4	8	12	16	20	24
	5	5	10	15	20	25	30
	6	6	12	18	24	30	36

# Random variable: expectation $\mathbb{E}$

Random variable (RV)

$X: \Omega \rightarrow \mathbb{R}$

Expectation: a weighted average  
(in proportion to the probabilities)  
of the possible values of  $X$

$$\mathbb{E}[X] = \sum_{\omega \in \Omega} X(\omega) \cdot \mathbb{P}(\{\omega\})$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

EXAMPLE 3 (CONT.):

roll two fair dice with 6 sides

$\Omega = \{(1,1), (1,2), (1,3) \dots (1,6), (2,1), (2,2), \dots (6,6)\}$

Let  $Y$  be the product of the rolls.



What is  $\mathbb{E}[Y]$ ?

$$\mathbb{E}[Y] = \mathbb{E}[X_1 \cdot X_2] = \mathbb{E}[X_1] \cdot \mathbb{E}[X_2]$$

↑  
*because the  $X_1 \perp X_2$*

$$\mathbb{E}[X_1] = \mathbb{E}[X_2] = 3.5$$

$$\mathbb{E}[Y] = 3.5 \cdot 3.5 = 12.25$$

		$\Omega$ second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)
		$Y$					
		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
	2	2	4	6	8	10	12
	3	3	6	9	12	15	18
	4	4	8	12	16	20	24
	5	5	10	15	20	25	30
	6	6	12	18	24	30	36

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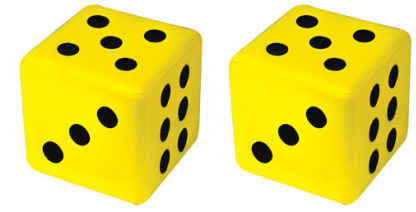
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Let  $Y$  be the product of the rolls.

Let  $X$  be the sum of the rolls.

What is  $\mathbb{E}[X + Y]$  ?



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		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
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	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

		$Y$					
		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
	2	2	4	6	8	10	12
	3	3	6	9	12	15	18
	4	4	8	12	16	20	24
	5	5	10	15	20	25	30
	6	6	12	18	24	30	36

		$X$					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
	3	4	5	6	7	8	9
	4	5	6	7	8	9	10
	5	6	7	8	9	10	11
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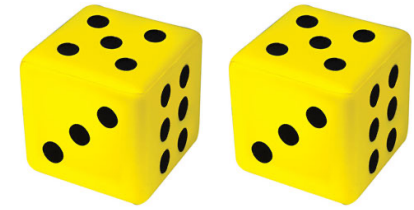
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What is  $\mathbb{E}[X + Y]$ ?

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y] \quad ?$$



		$\Omega$ second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
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		$Y$					
		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
	2	2	4	6	8	10	12
	3	3	6	9	12	15	18
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		$X$					
		1	2	3	4	5	6
first die	1	2	3	4	5	6	7
	2	3	4	5	6	7	8
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What is  $\mathbb{E}[X + Y]$ ?

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y]$$

$X$  and  $Y$  are clearly dependent ☹️



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	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
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	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
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		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
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	3	3	6	9	12	15	18
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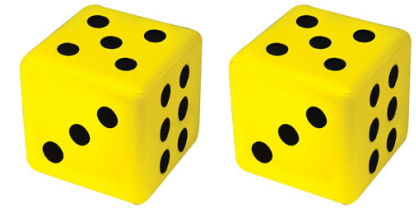
What is  $\mathbb{E}[X + Y]$ ?

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y]$$

$X$  and  $Y$  are clearly dependent ☹️

But linearity of expectation still holds  
even if the RVs are dependent 😊

$$\mathbb{E}[X] = 7 + 12.25 = 19.25$$



$\Omega$		second die					
		1	2	3	4	5	6
first die	1	(1,1)	(1,2)	(1,2)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
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		$Y$					
		1	2	3	4	5	6
first die	1	1	2	3	4	5	6
	2	2	4	6	8	10	12
	3	3	6	9	12	15	18
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		1	2	3	4	5	6
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# $\mathbb{E}[X + Y]$ vs. $\mathbb{E}[X \cdot Y]$



$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y] \text{ (linearity of expectation)}$$

*holds even if  $X$  and  $Y$  are not independent*

PROOF

$$\mathbb{E}[X + Y] = \sum_x \sum_y (x + y) \cdot p(x, y)$$

?

$$= \mathbb{E}[X] + \mathbb{E}[Y]$$

$$\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y], \text{ only if } X \perp Y \text{ (independent)}$$

PROOF

$$\mathbb{E}[X \cdot Y] = \sum_x \sum_y x \cdot y \cdot p(x, y)$$

?

$$= \mathbb{E}[X] \cdot \mathbb{E}[Y]$$

# $\mathbb{E}[X + Y]$ vs. $\mathbb{E}[X \cdot Y]$

$$\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y] \text{ (linearity of expectation)}$$

*holds even if  $X$  and  $Y$  are not independent*

PROOF

$$\begin{aligned}\mathbb{E}[X + Y] &= \sum_x \sum_y (x + y) \cdot p(x, y) \\ &= \sum_x \sum_y x \cdot p(x, y) + \sum_x \sum_y y \cdot p(x, y) \\ &= \sum_x x \cdot \sum_y p(x, y) + \sum_y y \cdot \sum_x p(x, y) \\ &= \sum_x x \cdot p(x) + \sum_y y \cdot p(y) \\ &= \mathbb{E}[X] + \mathbb{E}[Y]\end{aligned}$$

$$\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y], \text{ only if } X \perp Y \text{ (independent)}$$

PROOF

$$\begin{aligned}\mathbb{E}[X \cdot Y] &= \sum_x \sum_y x \cdot y \cdot p(x, y) \\ &= \sum_x \sum_y x \cdot y \cdot p(x) \cdot p(y) \\ &= \left( \sum_x x \cdot p(x) \right) \cdot \left( \sum_y y \cdot p(y) \right) \\ &= \mathbb{E}[X] \cdot \mathbb{E}[Y]\end{aligned}$$

*because  $X \perp Y$*

*distributive law of multiplication over addition (each cross-term appears exactly once on both sides)*

# Linearity of Expectation



EXAMPLE 6 (CONT.):

- The digits 1,2,3, and 4 are randomly arranged to form two two-digit numbers,  $\overline{AB}$  and  $\overline{CD}$ . For example, we could have  $\overline{AB} = 42$  and  $\overline{CD} = 13$ .
- What is the expected value of  $\overline{AB} \cdot \overline{CD}$ ?

?

# Linearity of Expectation



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- What is the expected value of  $\overline{AB} \cdot \overline{CD}$ ?
- What about  $\mathbb{E}[\overline{AB}] \cdot \mathbb{E}[\overline{CD}]$ ?

?

# Linearity of Expectation



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- What is the expected value of  $\overline{AB} \cdot \overline{CD}$ ?
- ~~What about  $\mathbb{E}[\overline{AB}] \cdot \mathbb{E}[\overline{CD}]$ ?~~
- **No!**  $\mathbb{E}[xy] = \mathbb{E}[x] \cdot \mathbb{E}[y]$  only holds when the RVs are independent.

Clearly,  $\overline{AB}$  and  $\overline{CD}$  are not independent since each digit can only be used once (e.g., if  $\overline{AB} = 42$  then we would know that  $\overline{CD}$  can only be 13 or 31).

# Linearity of Expectation

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- Can we instead get some kind of sum?



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- What is the expected value of  $\overline{AB} \cdot \overline{CD}$ ?

- ~~What about  $\mathbb{E}[\overline{AB}] \cdot \mathbb{E}[\overline{CD}]$ ?~~

- Can we instead get some kind of sum?

$$\overline{AB} \cdot \overline{CD} = (10 \cdot A + B) \cdot (10 \cdot C + D) = 100 \cdot A \cdot C + 10 \cdot A \cdot D + 10 \cdot B \cdot C + B \cdot D$$

- Now by linearity of expectation,

?

# Linearity of Expectation

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# Linearity of Expectation

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The expected value of all of these products are the same since there is symmetry among  $A, B, C, D$ .

# Linearity of Expectation

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$$\mathbb{E}[\overline{AB} \cdot \overline{CD}] = 100 \cdot \mathbb{E}[A \cdot C] + 10 \cdot \mathbb{E}[A \cdot D] + 10 \cdot \mathbb{E}[B \cdot C] + \mathbb{E}[B \cdot D] = 121 \cdot \mathbb{E}[A \cdot C] = \frac{4235}{6} = 705.8\dot{3}$$

The expected value of all of these products are the same since there is symmetry among  $A, B, C, D$ .

$$\mathbb{E}[A \cdot C] = \frac{1 \cdot 2 + 1 \cdot 3 + 1 \cdot 4 + 2 \cdot 3 + 2 \cdot 4 + 3 \cdot 4}{6} = \frac{35}{6}$$

# Variance

# Measuring variability

EXAMPLE: average size of mice

case 1:

{4 cm, 5 cm, 6 cm}

case 2:

{9 cm, 10 cm, 11 cm}

case 3:

{8 cm, 10 cm, 12 cm}

*What is a reasonable  
measure of centrality?*

# Measuring variability

EXAMPLE: average size of mice

	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}	What is a reasonable measure of centrality?
$\mathbb{E}[X] =$	5	10	10	$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$  mean = average = "expected value"

# Measuring variability

EXAMPLE: average size of mice

What are possible ways to measure expected "variability"  $\Delta$  around the mean for each point?

$$\mathbb{E}[Y] = \sum_y y \cdot p_Y(y)$$



$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}
$\mathbb{E}[X] =$	5	10	10

# Measuring variability

EXAMPLE: average size of mice

What are possible ways to measure expected "variability"  $\Delta$  around the mean for each point?

$$\mathbb{E}[Y] = \sum_y y \cdot p_Y(y)$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

$$Y_1 = X - \mathbb{E}[X]$$

How can we fix that ?

	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}
$\mathbb{E}[X] =$	5	10	10
$\mathbb{E}[Y_1] =$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ = 0	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ = 0	$-2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ = 0

☹️ cancels out

# Measuring variability

EXAMPLE: average size of mice

What are possible ways to measure expected "variability"  $\Delta$  around the mean for each point?

$$\mathbb{E}[Y] = \sum_y y \cdot p_Y(y)$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

$$Y_1 = X - \mathbb{E}[X]$$

$$Y_2 = |X - \mathbb{E}[X]|$$

absolute deviation

Anything else ?

	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}
$\mathbb{E}[X] =$	5	10	10
$\mathbb{E}[Y_1] =$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= 0$ ☹️ <i>cancels out</i>
$\mathbb{E}[Y_2] =$	$ -1  \cdot \frac{1}{3} +  0  \cdot \frac{1}{3} +  1  \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$	$ -1  \cdot \frac{1}{3} +  0  \cdot \frac{1}{3} +  1  \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$	$ -2  \cdot \frac{1}{3} +  0  \cdot \frac{1}{3} +  2  \cdot \frac{1}{3}$ $= 2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= \frac{4}{3}$ 😊 <i>makes more sense</i>



# Measuring variability

EXAMPLE: average size of mice

What are possible ways to measure expected "variability"  $\Delta$  around the mean for each point?

$$\mathbb{E}[Y] = \sum_y y \cdot p_Y(y)$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

$$Y_1 = X - \mathbb{E}[X]$$

$$Y_2 = |X - \mathbb{E}[X]|$$

absolute deviation

$$Y_3 = (X - \mathbb{E}[X])^2$$

squared error

	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}
$\mathbb{E}[X] =$	5	10	10
$\mathbb{E}[Y_1] =$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= 0$
$\mathbb{E}[Y_2] =$	$ -1  \cdot \frac{1}{3} +  0  \cdot \frac{1}{3} +  1  \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$	$ -1  \cdot \frac{1}{3} +  0  \cdot \frac{1}{3} +  1  \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$	$ -2  \cdot \frac{1}{3} +  0  \cdot \frac{1}{3} +  2  \cdot \frac{1}{3}$ $= 2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= \frac{4}{3}$
$\mathbb{E}[Y_3] =$	$1^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{1}{3}$ $1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$	$1^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{1}{3}$ $1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$	$2^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 2^2 \cdot \frac{1}{3}$ $4 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 4 \cdot \frac{1}{3}$ $= \frac{8}{3}$

☹️ cancels out

😊 makes more sense

What are now the "units" of variability



# Measuring variability

EXAMPLE: average size of mice

What are possible ways to measure expected "variability"  $\Delta$  around the mean for each point?

$$\mathbb{E}[Y] = \sum_y y \cdot p_Y(y)$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

$$Y_1 = X - \mathbb{E}[X]$$

$$Y_2 = |X - \mathbb{E}[X]|$$

absolute deviation

$$Y_3 = (X - \mathbb{E}[X])^2$$

squared error

	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}
$\mathbb{E}[X] =$	5	10	10
$\mathbb{E}[Y_1] =$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= 0$ ☹️ <i>cancels out</i>
$\mathbb{E}[Y_2] =$	$ -1  \cdot \frac{1}{3} +  0  \cdot \frac{1}{3} +  1  \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm	$ -1  \cdot \frac{1}{3} +  0  \cdot \frac{1}{3} +  1  \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm	$ -2  \cdot \frac{1}{3} +  0  \cdot \frac{1}{3} +  2  \cdot \frac{1}{3}$ $= 2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= \frac{4}{3}$ cm 😊 <i>makes more sense</i>
$\mathbb{E}[Y_3] =$	$1^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{1}{3}$ $1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm <sup>2</sup>	$1^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{1}{3}$ $1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm <sup>2</sup>	$2^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 2^2 \cdot \frac{1}{3}$ $4 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 4 \cdot \frac{1}{3}$ $= \frac{8}{3}$ cm <sup>2</sup> ☹️ ???

cm squared is strange. What can we do?

?

# Measuring variability

EXAMPLE: average size of mice

What are possible ways to measure expected "variability"  $\Delta$  around the mean for each point?

$$\mathbb{E}[Y] = \sum_y y \cdot p_Y(y)$$

$$\mathbb{E}[X] = \sum_x x \cdot p_X(x)$$

$$Y_1 = X - \mathbb{E}[X]$$

$$Y_2 = |X - \mathbb{E}[X]|$$

absolute deviation

$$Y_3 = (X - \mathbb{E}[X])^2$$

squared error

Looks pretty complicated. So why is everyone so excited about squared error instead of absolute error?

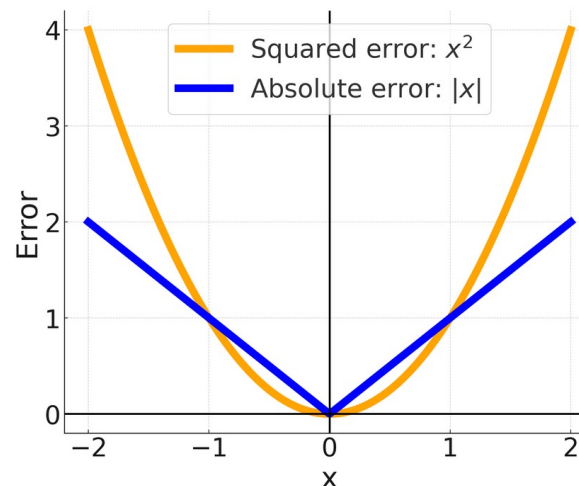
	case 1: {4 cm, 5 cm, 6 cm}	case 2: {9 cm, 10 cm, 11 cm}	case 3: {8 cm, 10 cm, 12 cm}
$\mathbb{E}[X] =$	5	10	10
$\mathbb{E}[Y_1] =$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= 0$	$-2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= 0$ ☹️ <i>cancels out</i>
$\mathbb{E}[Y_2] =$	$  -1   \cdot \frac{1}{3} +   0   \cdot \frac{1}{3} +   1   \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm	$  -1   \cdot \frac{1}{3} +   0   \cdot \frac{1}{3} +   1   \cdot \frac{1}{3}$ $= 1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm	$  -2   \cdot \frac{1}{3} +   0   \cdot \frac{1}{3} +   2   \cdot \frac{1}{3}$ $= 2 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 2 \cdot \frac{1}{3}$ $= \frac{4}{3}$ cm ☺️ <i>makes more sense</i>
$\mathbb{E}[Y_3] =$	$1^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{1}{3}$ $1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm <sup>2</sup>	$1^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{1}{3}$ $1 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 1 \cdot \frac{1}{3}$ $= \frac{2}{3}$ cm <sup>2</sup>	$2^2 \cdot \frac{1}{3} + 0^2 \cdot \frac{1}{3} + 2^2 \cdot \frac{1}{3}$ $4 \cdot \frac{1}{3} + 0 \cdot \frac{1}{3} + 4 \cdot \frac{1}{3}$ $= \frac{8}{3}$ cm <sup>2</sup> ☹️ ???
Just take the square root:	$= \sqrt{\frac{2}{3}}$ cm	$= \sqrt{\frac{2}{3}}$ cm	$= 2\sqrt{\frac{2}{3}}$ cm ???

# Why variance, and not absolute deviation?

But what is the most substantial reason?



- Geometric explanation in Euclidean space (distance = square root of squared components): Projection onto a subspace. Variance is literally the "average squared distance" from the mean.
- Variance plays an important role in the Central Limit Theorem. The variance is a natural parameter for the Normal Distribution.
- Algebraic convenience:  $f(x) = x^2$  is differentiable, but  $g(x) = |x|$  is not.
- Outliers have a bigger influence



$$\Delta_2 = |X - \mathbb{E}[X]|$$

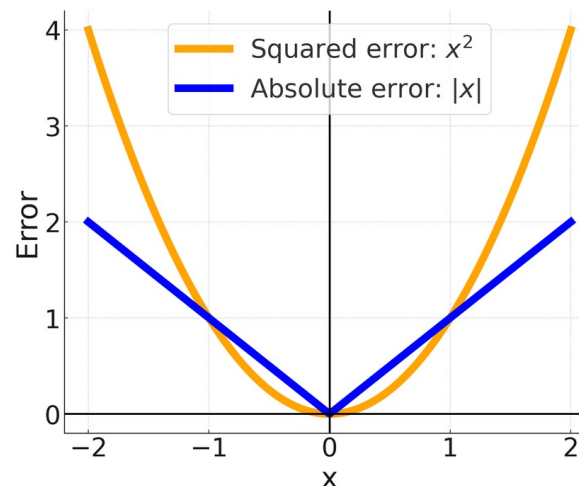
absolute deviation

$$\Delta_3 = (X - \mathbb{E}[X])^2$$

squared error

# Why variance, and not absolute deviation?

- Statistical properties: The mean minimizes the sum of squared errors, while the median minimizes the sum of absolute errors. (Maximum likelihood estimation leads to minimizing squared errors)
- Geometric explanation in Euclidean space (distance = square root of squared components): Projection onto a subspace. Variance is literally the "average squared distance" from the mean.
- Variance plays an important role in the Central Limit Theorem. The variance is a natural parameter for the Normal Distribution.
- Algebraic convenience:  $f(x) = x^2$  is differentiable, but  $g(x) = |x|$  is not.
- Outliers have a bigger influence



$$\Delta_2 = |X - \mathbb{E}[X]|$$

absolute deviation

$$\Delta_3 = (X - \mathbb{E}[X])^2$$

squared error

# Mean (minimizes variance), Median (minimizes absolute error)

EXAMPLE (MEAN VS. MEDIAN):

data = {1, 2, 3, 6, 8}

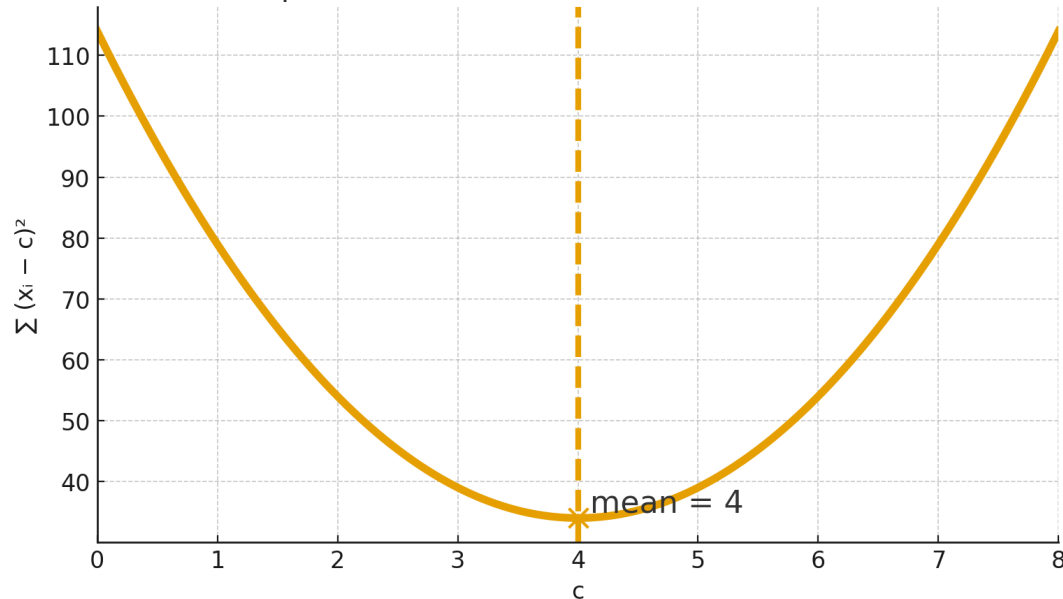
mean = 4

$$\text{mean} = \min_c \sum_i (x_i - c)^2$$

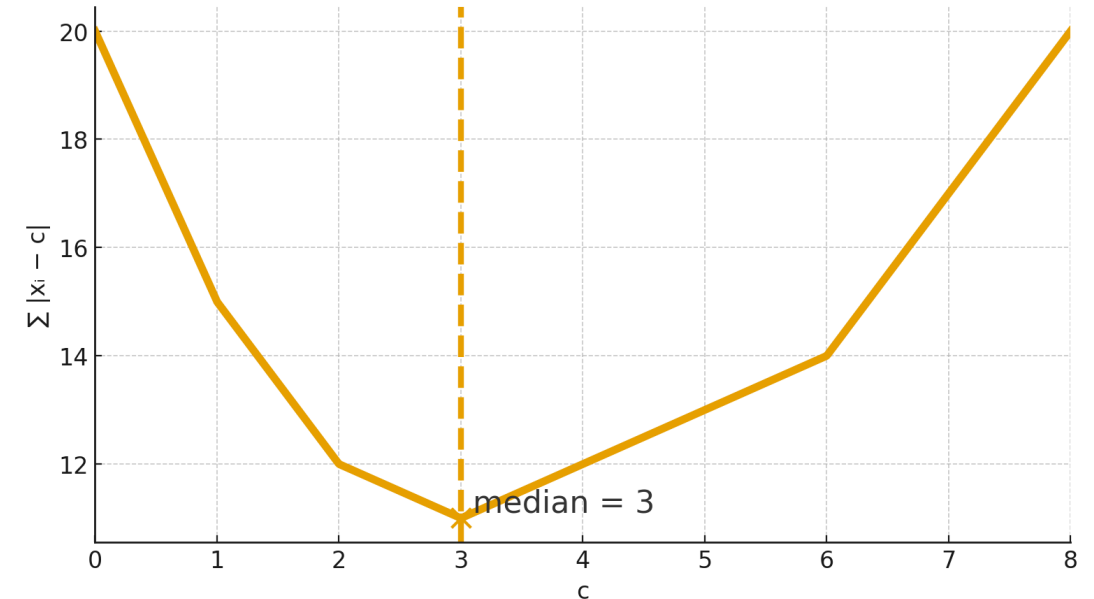
median = 3

$$\text{median} = \min_c \sum_i |x_i - c|$$

Squared-error loss minimized at the mean



Absolute-error loss minimized at the median



# Mean (minimizes variance), Median (minimizes absolute error)

EXAMPLE (MEAN VS. MEDIAN):

data = {1, 2, 3, 6, 8}

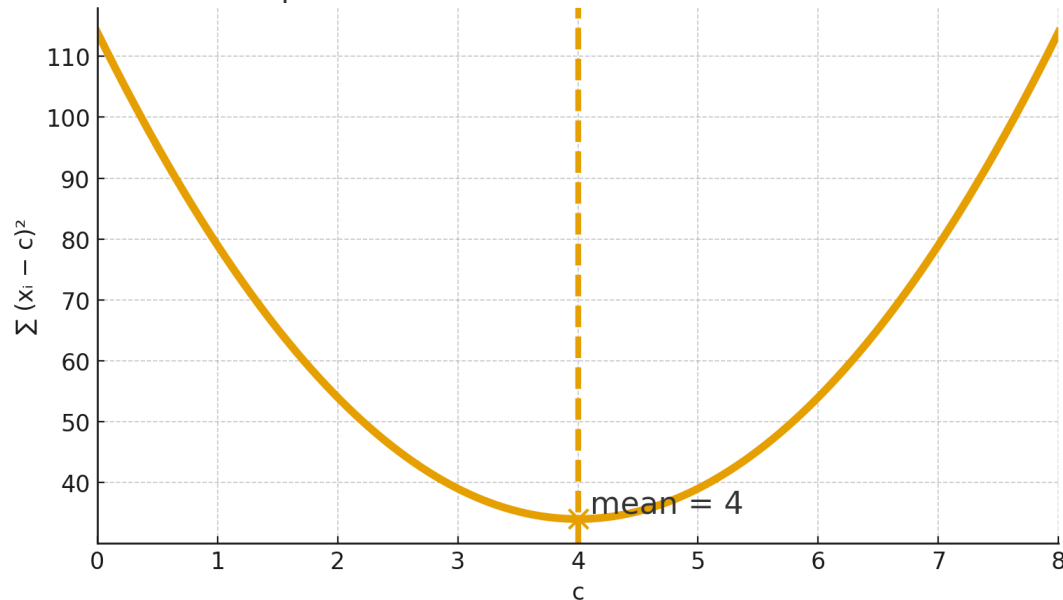
mean = 4

$$\text{mean} = \min_c \sum_i (x_i - c)^2$$

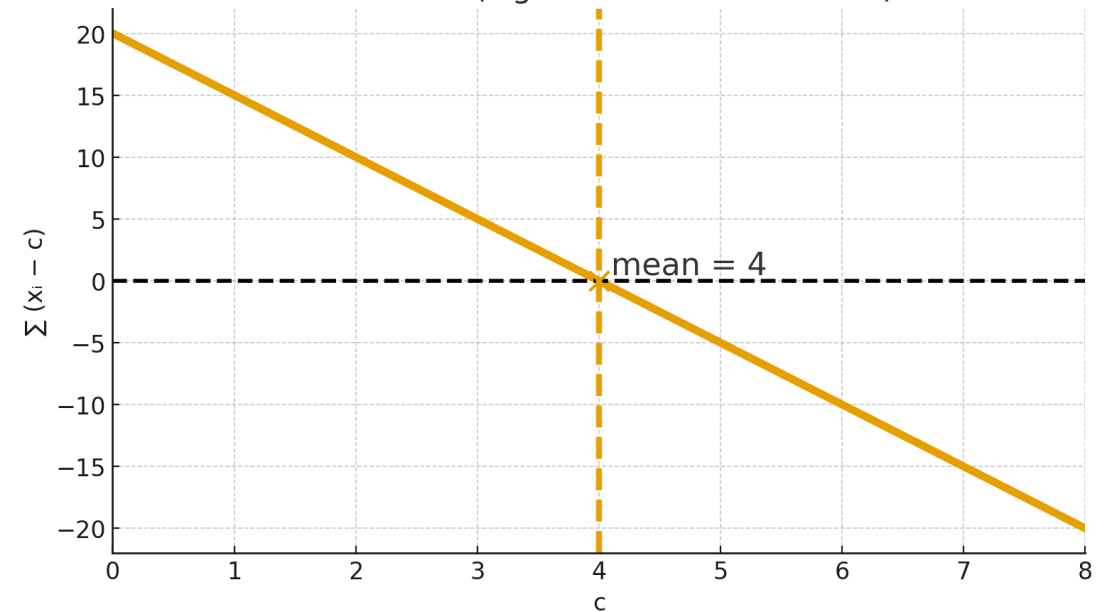
This is not an optimization problem, but a constraint problem: the sum of signed errors = 0.

$$\text{mean: } \sum_i x_i - c = 0$$

Squared-error loss minimized at the mean



Total error (signed sum of deviations)



# Alternative formula for variance

$$\sigma^2 = \text{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] \quad \text{variance}$$

$$\sigma = \sqrt{\text{Var}(X)} = \sqrt{\mathbb{E}[(X - \mathbb{E}[X])^2]} \quad \text{standard deviation (back in original units)}$$

## ALTERNATIVE FORMULA

$$\mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$

## PROOF

$$\begin{aligned} \mathbb{E}[(X - \mathbb{E}[X])^2] &= \mathbb{E}[X^2 - 2 \cdot X \cdot \mathbb{E}[X] + (\mathbb{E}[X])^2] \\ &\quad \text{by linearity of expectation} \\ &= \mathbb{E}[X^2] - \mathbb{E}[2 \cdot X \cdot \mathbb{E}[X]] + \mathbb{E}[(\mathbb{E}[X])^2] \\ &\quad \text{just a constant} \\ &= \mathbb{E}[X^2] - 2 \cdot \mathbb{E}[X] \cdot \mathbb{E}[X] + (\mathbb{E}[X])^2 \\ &= \mathbb{E}[X^2] - 2 \cdot (\mathbb{E}[X])^2 + (\mathbb{E}[X])^2 \\ &= \mathbb{E}[X^2] - (\mathbb{E}[X])^2 \end{aligned}$$

case 3:

{8 cm, 10 cm, 12 cm}

$$\mathbb{E}[X^2] = \frac{8^2 + 10^2 + 12^2}{3} = \frac{308}{3}$$

$$(\mathbb{E}[X])^2 = 100$$

$$\frac{308}{3} - \frac{300}{3} = 8/3$$