

Part 1: Theory

L02: Basics of Probability (1/2)

[Random experiment, independence, conditional probability, chain rule, Bayes' theorem, random variables]

Javed Aslam, Wolfgang Gatterbauer

cs7840 Foundations and Applications of Information Theory (fa24)

<https://northeastern-datalab.github.io/cs7840/fa24/>

9/9/2024

Pre-class conversations

- Last class recapitulation
- We strive to keep it interactive, also among us faculty
- Why class slides need up to 2 days after a class
- Office hours: Usually right after class, or via email / Teams
- Organizational matters: Piazza messages? Canvas pictures did not display correctly?
- New class arrivals
- Today:
 - The basics of probability theory

Last time

- Introduction
- Course logistics
- Entropy
 - examples

Today

- Probability primer

Next time

- Basic concepts
of Info- Theory

Probability

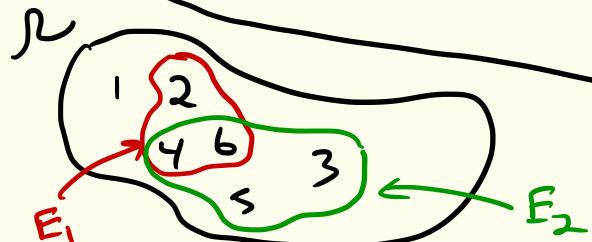
- Random experiment
- Generate outcomes $\omega \in \mathcal{R}$
- sample space: set of \mathcal{R}
all possible outcomes

- Event: subset of sample space

- $p: \mathcal{R} \rightarrow \mathbb{R}$ probability measure

- $0 \leq p(\omega) \leq 1 \quad \forall \omega \in \mathcal{R}$

- $\sum_{\omega \in \mathcal{R}} p(\omega) = 1$



we will assume for now that $p(\omega) = \frac{1}{|\mathcal{R}|} \quad \forall \omega \in \mathcal{R}$

- $E_1 = \text{"even"} = \{2, 4, 6\}$

- $E_2 = \text{"}\geq 3\text{"} = \{3, 4, 5, 6\}$

- $p(1) = p(2) = p(3) = \dots = p(6) = \frac{1}{6}$

$$P(E) = \sum_{\omega \in E} p(\omega)$$

e.g. $P(E_1) = p(2) + p(4) + p(6) = \frac{1}{6} + \frac{1}{6} + \frac{1}{6} = \frac{1}{2}$

$\hat{\equiv}$ $p(\omega) = \frac{1}{|\mathcal{R}|} \quad \forall \omega \in \mathcal{R}$

$$P(E) = |E| / |\mathcal{R}| \quad P(E_1) = \frac{|E_1|}{|\mathcal{R}|} = \frac{3}{6} = \frac{1}{2}$$

Example

- roll a fair six-sided die
- roll a 5
- $\{1, 2, 3, 4, 5, 6\} = \mathcal{R}$

Examples

① Roll one fair die

$$\mathcal{R} = \{1, 2, 3, 4, 5, 6\}$$

② Roll two fair die

$$\mathcal{R} = \{(1,1), (1,2), (1,3), \dots, (2,1), (2,2), \dots, (6,6)\}$$

$$= \{1, 2, 3, \dots, 6\} \times \{1, 2, \dots, 6\}$$

$$E_1 = \text{total is 7} = \{(1,6), (2,5), (3,4), (4,3), (5,2), (6,1)\}$$

$$P(E_1) = \frac{|E_1|}{|\mathcal{R}|} = \frac{6}{36} = \frac{1}{6}$$

$E_2 = \text{total is greater than 8}$

= 9 or 10 or 11 or 12

first
die

		1	2	3	4	5	6
		1	(1,1)	(1,2)			
		2					
		3					
		4					
		5					
		6					(6,6)

$$E_2 = \begin{matrix} 12 \\ \downarrow \\ 1 + 2 + 3 + 4 \end{matrix} = 10$$

$$P(E_2) = \frac{|E_2|}{|\mathcal{R}|} = \frac{10}{36} = \frac{5}{18}$$

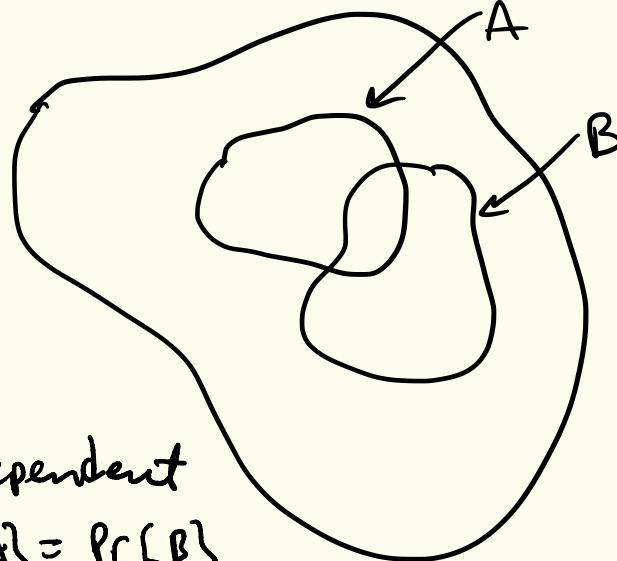
$$\textcircled{1} \quad \Pr\{A|B\} = \frac{\Pr\{A \cap B\}}{\Pr\{B\}}$$

$$\textcircled{2} \quad \Pr\{B|A\} = \frac{\Pr\{A \cap B\}}{\Pr\{A\}}$$

Independence: Two events A & B are independent if $\Pr\{A|B\} = \Pr\{A\}$, $\Pr\{B|A\} = \Pr\{B\}$

\Rightarrow Knowing B does not change your belief in A.
Knowing A does not change your belief in B.

Claim: A & B are independent if and only if $\Pr\{A \cap B\} = \Pr\{A\} \cdot \Pr\{B\}$



$$\textcircled{1} \quad \Pr\{A|B\} = \frac{\Pr\{A \cap B\}}{\Pr\{B\}}$$

Two events A & B are independent
if $\Pr\{A|B\} = \Pr\{A\}$, $\Pr\{B|A\} = \Pr\{B\}$

$$\textcircled{2} \quad \Pr\{B|A\} = \frac{\Pr\{A \cap B\}}{\Pr\{A\}}$$

Claim: A & B are independent if and only if $\Pr\{A \cap B\} = \Pr\{A\} \cdot \Pr\{B\}$

Proof:

(\Rightarrow) If $\Pr\{A|B\} = \Pr\{A\}$, then by ① $\Pr\{A \cap B\} = \Pr\{A|B\} \cdot \Pr\{B\} = \Pr\{A\} \cdot \Pr\{B\}$

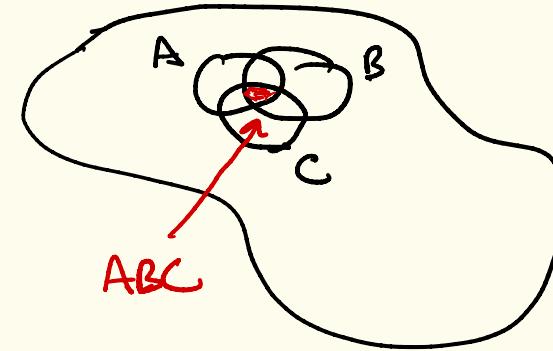
(\Leftarrow) If $\Pr\{A \cap B\} = \Pr\{A\} \cdot \Pr\{B\}$, then by ① & ②

$$\Pr\{A|B\} = \frac{\Pr\{A \cap B\}}{\Pr\{B\}} = \frac{\Pr\{A\} \cdot \Pr\{B\}}{\Pr\{B\}} = \Pr\{A\}$$

$$\Pr\{B|A\} = \frac{\Pr\{A \cap B\}}{\Pr\{A\}} = \frac{\Pr\{A\} \cdot \Pr\{B\}}{\Pr\{A\}} = \Pr\{B\} \quad \therefore$$

Chain Rules

- $\Pr\{ABC\} = ?$
 - treat AB as an event



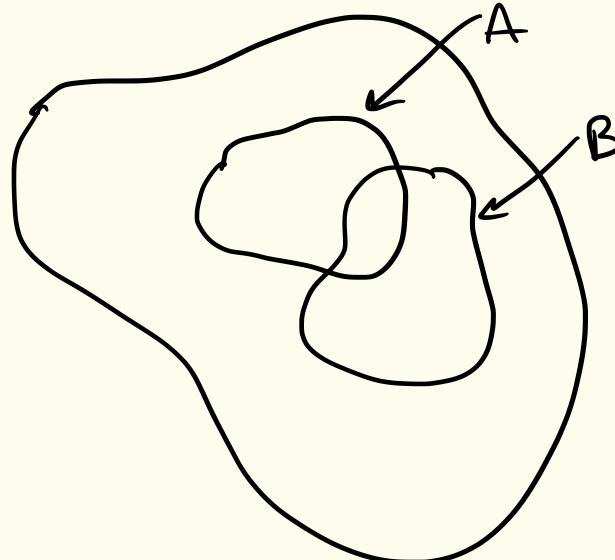
$$\begin{aligned}\Pr\{ABC\} &= \Pr\{AB\} \cdot \Pr\{C|AB\} \\ &= \Pr\{A\} \cdot \Pr\{B|A\} \cdot \Pr\{C|AB\}\end{aligned}$$

Generally : $\Pr\{x_1, x_2, \dots, x_n\} = \Pr\{x_1\} \cdot \Pr\{x_2|x_1\} \cdot \Pr\{x_3|x_1, x_2\} \cdots \Pr\{x_n|x_1, \dots, x_{n-1}\}$

Bayes Law

$$\textcircled{1} \quad \Pr\{A|B\} = \frac{\Pr\{A \cap B\}}{\Pr\{B\}}$$

$$\textcircled{2} \quad \Pr\{B|A\} = \frac{\Pr\{A \cap B\}}{\Pr\{A\}}$$



$$\underbrace{\Pr\{A|B\} \cdot \Pr\{B\}}_{\textcircled{1}} = \Pr\{A \cap B\} = \underbrace{\Pr\{B|A\} \cdot \Pr\{A\}}_{\textcircled{2}}$$

Solving for $\Pr\{A|B\}$...

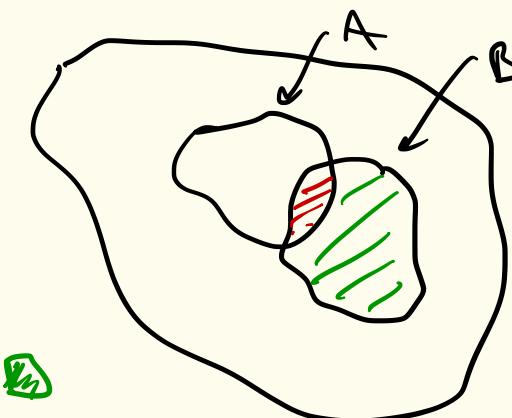
$$\Pr\{A|B\} = \frac{\Pr\{B|A\} \cdot \Pr\{A\}}{\Pr\{B\}}$$

Bayes Law

$$\cdot \Pr\{A|B\} = \frac{\Pr\{B|A\} \cdot \Pr\{A\}}{\Pr\{B\}}$$

• what is $\Pr\{B\}$?

$$\Pr\{A|B\} = \frac{\Pr\{B|A\} \cdot \Pr\{A\}}{\Pr\{B|A\} \cdot \Pr\{A\} + \Pr\{B|\bar{A}\} \cdot \Pr\{\bar{A}\}}$$



$$\begin{aligned}\Pr\{B\} &= \text{Red} + \text{Green} \\ &= \Pr\{A \wedge B\} + \Pr\{\bar{A} \wedge B\} \\ &= \Pr\{B|A\} \cdot \Pr\{A\} + \Pr\{B|\bar{A}\} \cdot \Pr\{\bar{A}\}\end{aligned}$$

$$\Pr\{H|E\} = \frac{\Pr\{E|H\} \cdot \Pr\{H\}}{\Pr\{E|H\} \cdot \Pr\{H\} + \Pr\{E|\bar{H}\} \cdot \Pr\{\bar{H}\}}$$

H : hypothesis — patient has Zika
 E : evidence — patient tested positive on Zika blood test

Example : Zika in FL in 2016

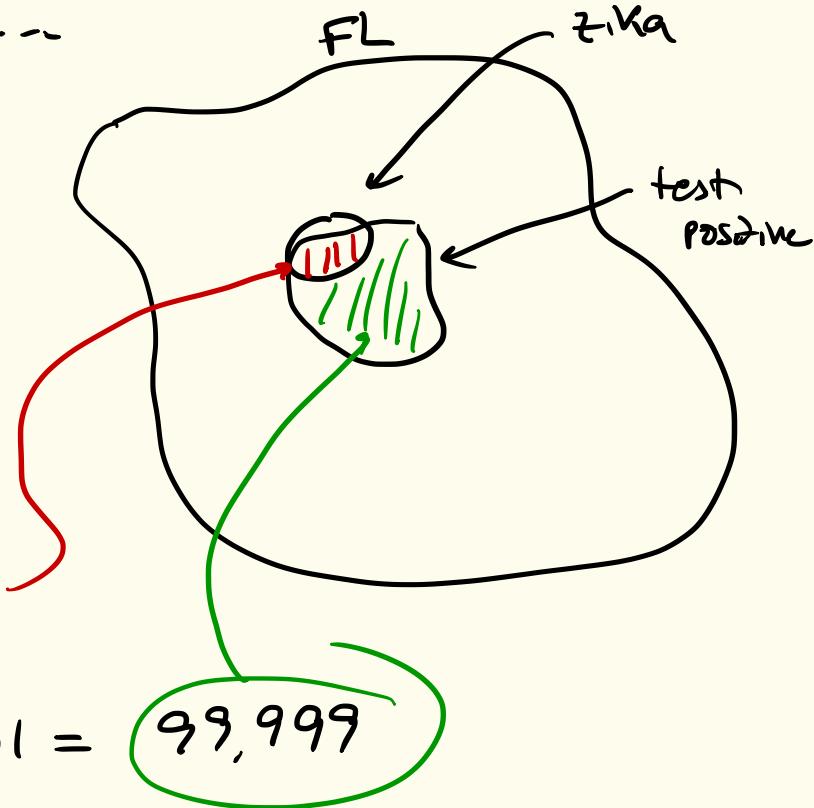
- Prevalence of Zika in FL $\Pr(\text{Zika}) = 10^{-5}$ (1 in 100,000)
- accuracy of blood test is 99 %
 - e.g. $\Pr(\text{pos. test} | \text{Zika}) = 0.99$ or $\Pr(\text{neg. test} | \text{Zika}) = 0.01$
 - $\Pr(\text{pos. test} | \text{no Zika}) = 0.01$ $\Pr(\text{neg. test} | \text{no Zika}) = 0.99$
- patient tests positive : what is chance they have Zika?
 - $\Rightarrow \underline{\text{Not}} \quad \Pr(\text{pos test} | \text{Zika}) = 0.99$
 - $\Rightarrow \text{You want} \quad \Pr(\text{Zika} | \text{pos test})$

$$\Pr\{H|E\} = \frac{\Pr\{E|H\} \cdot \Pr\{H\}}{\Pr\{E|H\} \cdot \Pr\{H\} + \Pr\{E|\bar{H}\} \cdot \Pr\{\bar{H}\}}$$

$$\begin{aligned} \Pr\{\text{zika} | \text{pos test}\} &= \frac{\Pr\{\text{pos.test} | \text{zika}\} \cdot \Pr\{\text{zika}\}}{\Pr\{\text{pos.test} | \text{zika}\} \cdot \Pr\{\text{zika}\} + \Pr\{\text{pos.test} | \text{not zika}\} \cdot \Pr\{\text{not zika}\}} \\ &= \frac{0.99 \cdot 10^{-5}}{0.99 \cdot 10^{-5} + 0.01 \cdot (1 - 10^{-5})} \\ &= \frac{0.000099}{0.000099 + 0.0099999} \\ &\approx 0.00099 \\ &\approx 0.1\% \quad \text{i.e., about 1 in 1,000!} \end{aligned}$$

Seems wildly counterintuitive, but...

- 10,000,000 people in FL 10^7
- w/ zika? $10^7 \cdot 10^{-5} = 10^2 = 100$
- w/o zika $10^7 - 10^2 = 9,999,900$



$$\text{test pos w/zika : } 100 \cdot 0.99 = 99$$

$$\text{test pos. w/o zika : } 9,999,900 \cdot 0.01 = 99,999$$

∴ among those who test pos,
only 99 out of (99 + 99,999)
actually have zika - about 1 in 1000.

Random Variable

$$X: \Omega \rightarrow \mathbb{R}$$

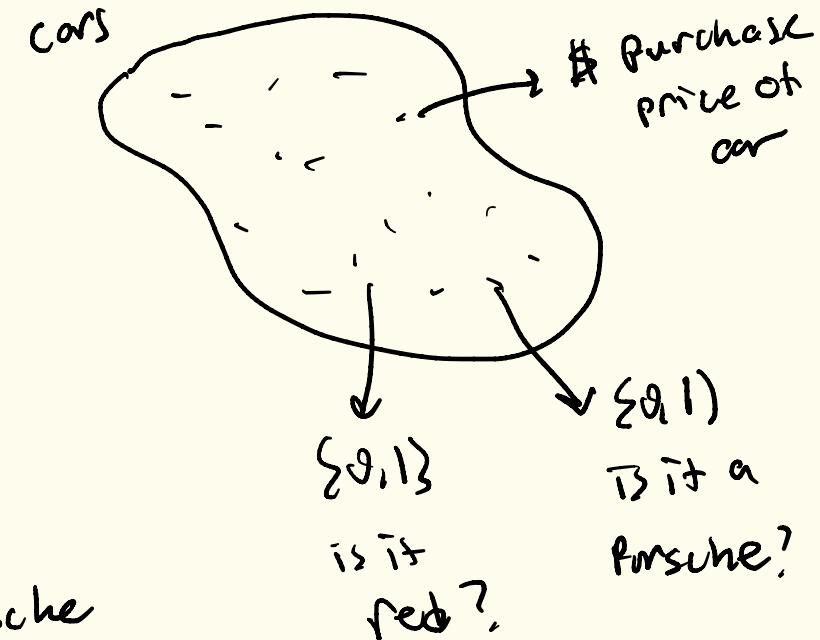
Two primary types:

① "numerical"

e.g. $X(\omega) = \text{price of } \omega \text{ (dollars)}$

② "indicator"

e.g. $X(\omega) = \begin{cases} 1 & \text{if } \omega \text{ is a Porsche} \\ 0 & \text{if not} \end{cases}$



$\Rightarrow X$ is a "random" variable because it depends on the outcome of a random experiment

\Rightarrow Underlying probability measure $p: \Omega \rightarrow \mathbb{R}$ induces a distribution D over the range of the random variable

$$D: \mathbb{R} \rightarrow \mathbb{R}$$

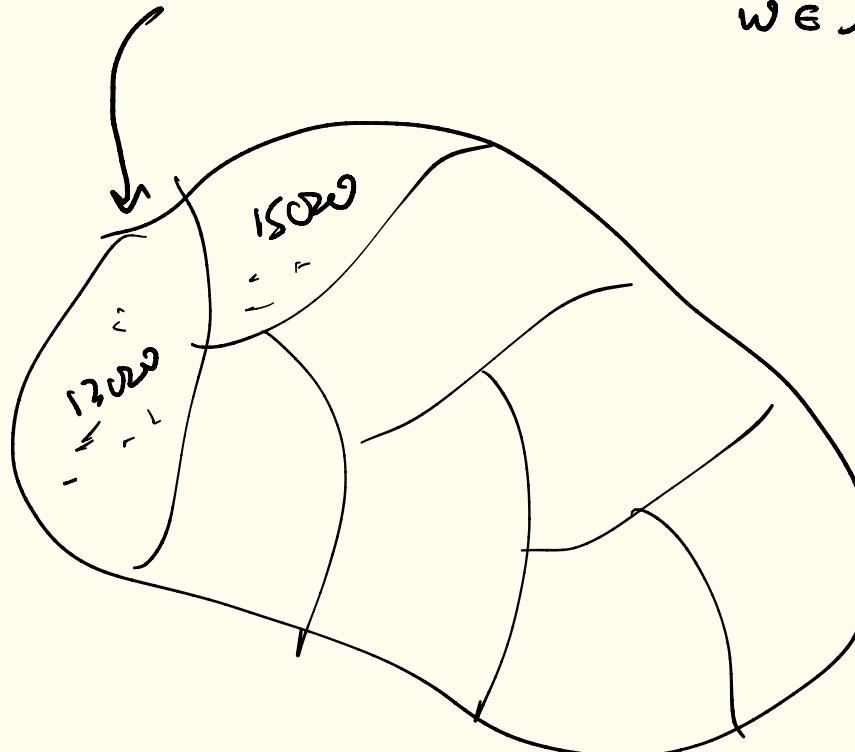
Expectation

- Random variable

$$X: \Omega \rightarrow \mathbb{R}$$

- $E[X] = \sum_x x \cdot \Pr[X=x]$

$$E[X] = \sum_{w \in \Omega} x(w) \cdot p(w)$$



Part 1: Theory

L03: Basics of Probability (2/2)

[Expectation, Variance, Markov chains]

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9/11/2024

Pre-class conversations

- Last class recapitulation
- PDF Class slides organized by topic (with subsections for classes)
- Any organizational matters: Piazza messages? Organizational matters?
- New class arrivals?
- Today:
 - The basics of probability theory
 - Intuition behind "information" (and "information measures")

Expectation

• Random variable

$$X: \Omega \rightarrow \mathbb{R}$$

$$E\{X\} = \sum_x x \cdot \Pr[X=x]$$



expected or

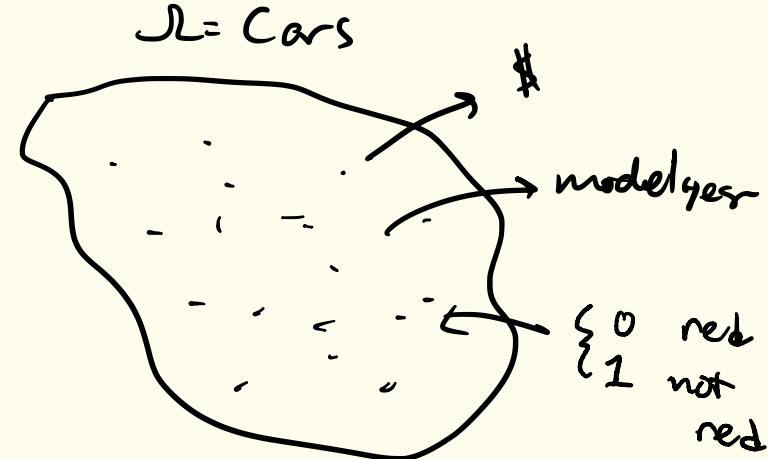
"average" value

$$\frac{1}{2} \text{ cars } \$27,000$$

$$\frac{1}{3} \text{ cars } \$10,000$$

$$\frac{1}{6} \text{ cars } \$15,000$$

$$E[X] = \sum_x x \cdot \Pr[X=x] = \frac{1}{2} \cdot 27000 + \frac{1}{3} \cdot 10000 + \frac{1}{6} \cdot 15000 = \$19,333.33$$



$$E[X] = \sum_{\omega \in \Omega} X(\omega) \cdot p(\omega)$$

Example

- Pay \$6 to play game
- Roll two fair 6-sided dice
- Pay you sum of die faces
except if double, then 0
- $X = \text{winnings (Profit)}$

$$\cdot E[X] = \sum_x x \cdot \Pr\{X=x\}$$

$$x = -6 \quad \Pr\{X=-6\} = \frac{6}{36} = \frac{1}{6}$$

$$x = -3 \quad \Pr\{X=-3\} = \frac{2}{36} = \frac{1}{18}$$

$$\vdots$$

$$x = +5 \quad \Pr\{X=5\} = \frac{2}{36} = \frac{1}{18}$$

$$E[X] = \sum_x x \cdot \Pr\{X=x\}$$

$$= (-6) \cdot \frac{1}{6} + (-3) \cdot \frac{1}{18} + \dots + 5 \left(\frac{1}{18} \right) = -0.16\bar{6}$$

dice 1

	1	2	3	4	5	6
1	-6	-3	-2	-1	0	1
2	-3	-6	-1	0	1	2
3	-2	-1	-6	1	2	3
4	-1	0	1	-6	3	4
5	0	1	2	3	-6	5
6	1	2	3	4	5	-6

$$\cdot E[X] = \sum_{\omega \in \Omega} X(\omega) \cdot p(\omega)$$

$$\text{induced distribution} = \frac{1}{36} \sum_{\omega \in \Omega} X(\omega)$$

$$= \frac{1}{36} (\text{"sum whole table"})$$

$$= -6/36 = -0.16\bar{6}$$

\Rightarrow lose 16¢ per play, on average

More on Expectation

- Example:
- Roll two fair 6-sided die
 - Let $X = \text{sum of die faces}$
 - Q: $E\{X\}?$

$$X: \Omega \rightarrow \mathbb{R}$$

e.g. $X(1,2,4) \rightarrow 6$

$$E\{X\} = \sum_x x \cdot \Pr\{X=x\}$$

$\left. \begin{array}{l} x=2 \\ x=3 \\ x=4 \\ \vdots \\ x=12 \end{array} \right\}$	$\Pr\{X=2\} = 1/36$ $\Pr\{X=3\} = 2/36$ $\Pr\{X=4\} = 3/36$ \vdots $\Pr\{X=12\} = 1/36$
--	---

$$E\{X\} = \sum_x x \cdot \Pr\{X=x\}$$

$$= 2 \cdot 1/36 + 3 \cdot 2/36 + \dots + 12 \cdot 1/36 = 7$$

		die 1		die 2			
		1	2	3	4	5	6
die 1	1	(1,1)	(1,2)				
	2	(2,1)			(2,4)		
3							
4							
5							
6							(6,6)

$$\begin{aligned} E\{X\} &= \sum_{w \in \Omega} X(w) \cdot \rho(w) \\ &= \sum_{w \in \Omega} x(w) \cdot 1/36 \\ &= 1/36 \cdot \sum_{w \in \Omega} X(w) \\ &= 1/36 \cdot \{\text{sum of table}\} \\ &\vdots \\ &= 7 \end{aligned}$$

Linearity of Expectation

Let $X_1 = \text{r.v. for first die roll}$

Let $X_2 = \text{r.v. for second die roll}$

Let $X = X_1 + X_2$ *

$$E\{X\} = E\{X_1 + X_2\} = E\{X_1\} + E\{X_2\} = 3.5 + 3.5 = 7$$

Variance & standard deviation

case 1

$$4'10'' \quad 5' \quad 5'2''$$

$$E\{x_1\} = 5'$$

case 2

$$4' \quad 5' \quad 6'$$

$$E\{x_2\} = 5'$$

case 3

$$3' \quad 5' \quad 7'$$

$$E\{x_3\} = 5'$$

case 3 How to measure "variability"

$$\begin{aligned} \textcircled{1} \quad E\{y_1\} &= \sum_{w \in \Omega} y(w) \cdot p(w) \\ &= -12'' \cdot 1/3 + 0'' \cdot 1/3 + (+12'') \cdot 1/3 \\ &= 0'' \end{aligned}$$

$$\begin{aligned} \textcircled{2} \quad E\{y_2\} &= (-12'') \cdot 1/3 + |0''| \cdot 1/3 + |+12''| \cdot 1/3 \\ &= 12 \cdot 1/3 + 0 \cdot 1/3 + 12 \cdot 1/3 = 8'' \end{aligned}$$

$$\begin{aligned} \textcircled{3} \quad E\{y_3\} &= (-12'')^2 \cdot 1/3 + (0'')^2 \cdot 1/3 + (12'')^2 \cdot 1/3 \\ &= 96 \text{ in}^2 \end{aligned}$$

3 ways

$$\textcircled{1} \quad y_1 = x - E\{x\}$$

X nes &
pos
cancel

$$\textcircled{2} \quad y_2 = |x - E\{x\}|$$

- mean
absolute
deviation

$$\textcircled{3} \quad y_3 = (x - E\{x\})^2$$

- variance

\Rightarrow take square root, set standard deviation $\sqrt{96 \text{ in}^2} = 9.8 \text{ in}$

$$\sigma^2 = \text{Var}(x) = E\{(x - E[x])^2\} \quad - \underline{\text{variance}}$$

$$\Rightarrow \sigma = \sqrt{\text{Var}(x)} = \sqrt{E\{(x - E[x])^2\}} \quad - \underline{\text{standard deviation}}$$

- back in original units

Claim: $E\{(x - E[x])^2\} = E[x^2] - (E[x])^2$

example: case 2 $E(x) = s' = 60''$

$$E[x^2] = \frac{(48'')^2 + (60'')^2 + (72'')^2}{3}$$

$$\text{Claim: } 3696 - 60^2 = 3696 - 3600 = \underline{\underline{96}}$$

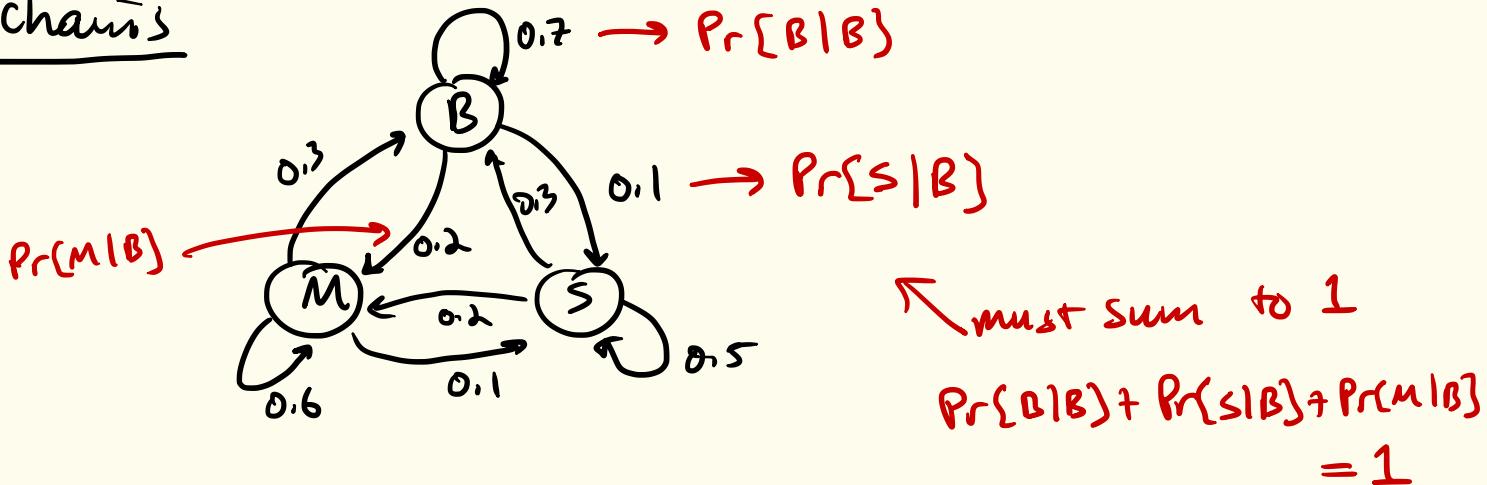
$$\begin{aligned} \text{Pf: } E\{(x - E[x])^2\} &= E[x^2 - 2x E[x] + (E[x])^2] \\ &= E[x^2] - E[2x E[x]] + E[(E[x])^2] \\ &= E[x^2] - 2E[x] \cdot E[x] + (E[x])^2 \\ &= E[x^2] - 2(E[x])^2 + (E[x])^2 \\ &= E[x^2] - (E[x])^2 \end{aligned}$$

Markov chains

B: Bertucci's

M: Margaritas

S: Sato



State transition matrix:

$$P = \begin{matrix} & \begin{matrix} B & M & S \end{matrix} \\ \begin{matrix} B \\ M \\ S \end{matrix} & \begin{pmatrix} 0.7 & 0.2 & 0.1 \\ 0.3 & 0.6 & 0.1 \\ 0.1 & 0.2 & 0.5 \end{pmatrix} \end{matrix} \quad \begin{matrix} \cdot \text{ all rows sum to 1} \\ \cdot \text{ stochastic matrix} \end{matrix}$$

Q : If I were to "run" the Markov Chain, what is the long-term fraction of time I would spend visiting each of the nodes?

A: Stationary distribution

$$\vec{\pi} = \langle \pi_B, \pi_M, \pi_S \rangle \quad \pi_B + \pi_M + \pi_S = 1$$

```
[jay@jay-mac-2021 Downloads % ./restaurantMC.pl 10
```

BMBBSSSBMB

Counts: B=5; M=2; S=3

Prob: B=0.5; M=0.2; S=0.3

```
[jay@jay-mac-2021 Downloads % ./restaurantMC.pl 100
```

BSBMBSSSMMMMMMMMMMSSBBBBSMMMMMSSSMMMBBMMMBBBBMMMBBBBBBSSSSMBBBSSMBBBMMMBBMMMBBMMMBSSMMMBBBBBBBBB

Counts: B=37; M=46; S=17

Prob: B=0.37; M=0.46; S=0.17

```
[jay@jay-mac-2021 Downloads % ./restaurantMC.pl 1000
```

Counts: B=488; M=366; S=146

Prob: B=0.488; M=0.366; S=0.146

⇒ way too long to converge, and just for 3 nodes!

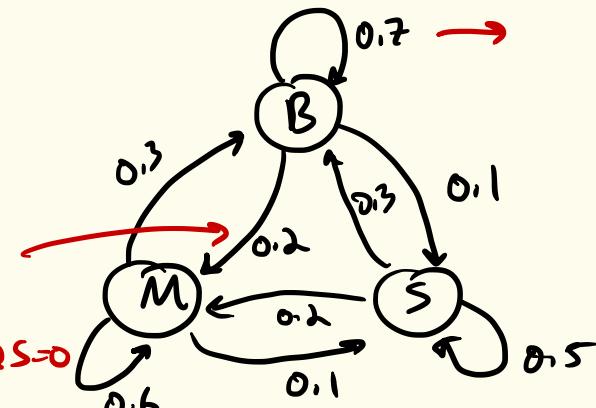
① simulate - takes too long to converge

② solve algebraically

$$① B = .7 \cdot B + .3 \cdot M + .3 \cdot S$$

$$② M = 0.2 \cdot B + 0.6 M + 0.2 \cdot S \rightarrow -0.2B + 0.4M - 0.2S = 0$$

$$③ S = 0.1 \cdot B + 0.1 \cdot M + 0.5 \cdot S \rightarrow -0.1B - 0.1M + 0.5S = 0$$



$$① .3B - .3M - .3S = 0$$

But also have

$$② + ③ .3B - .3M - .3S = 0$$

$$\cancel{④} B + M + S = 1$$

$$B + M + S = 1$$

$$5 \times ② -B + 2M - S = 0$$

$$10 \times ③ -B - M + 5S = 0$$

$$3M = 1 \Rightarrow M = \frac{1}{3}$$

$$M = \frac{1}{3}$$

$$6S = 1 \Rightarrow S = \frac{1}{6}$$

$$S = \frac{1}{6}$$

Works, but in general, solving n equations in n unknowns takes $O(n^3)$ time, which is infeasible for large n .

Guess an answer for
stationary distribution

$$B_0 = \frac{1}{3}$$

$$M_0 = \frac{1}{3}$$

$$S_0 = \frac{1}{3}$$

Try it

$$\textcircled{1} \quad B = .7 \cdot B + .3 \cdot M + .3 \cdot S$$

$$\textcircled{2} \quad M = .2 \cdot B + .6 \cdot M + .2 \cdot S$$

$$\textcircled{3} \quad S = .1 \cdot B + .1 \cdot M + .5 \cdot S$$

$$B_1 = .7 \times \frac{1}{3} + .3 \times \frac{1}{3} + .3 \times \frac{1}{3} = \frac{4}{3} \text{ new guess } \textcircled{1} \quad \frac{1}{2}$$

$$M_1 = .2 \times \frac{1}{3} + .6 \times \frac{1}{3} + .2 \times \frac{1}{3} = \frac{3}{3} \text{ } \textcircled{2} \quad \frac{1}{3}$$

$$S_1 = .1 \times \frac{1}{3} + .1 \times \frac{1}{3} + .5 \times \frac{1}{3} = \frac{2}{3} \text{ } \textcircled{3} \quad \frac{1}{6}$$

$$B_2 = .7 \times \frac{4}{3} + .3 \times \frac{3}{3} + .3 \times \frac{2}{3} = \frac{47}{30} \text{ new guess } \textcircled{1}$$

$$M_2 = \dots = \frac{3}{3} \text{ } \textcircled{2}$$

$$S_2 = \dots = \frac{19}{30} \text{ } \textcircled{3}$$

new guess
⇒ iterate
until convergence

```
[jay@jay-mac-2021 Downloads % ./markovIterate.pl 20 transition.txt  
Processing transition matrix...
```

```
0.333333333333333 0.333333333333333 0.333333333333333  
0.433333333333333 0.333333333333333 0.233333333333333  
0.473333333333333 0.333333333333333 0.193333333333333  
0.489333333333333 0.333333333333333 0.177333333333333  
0.495733333333333 0.333333333333333 0.170933333333333  
0.498293333333333 0.333333333333333 0.168373333333333  
0.499317333333333 0.333333333333333 0.167349333333333  
0.499726933333333 0.333333333333333 0.166939733333333  
0.499890773333333 0.333333333333333 0.166775893333333  
0.499956309333333 0.333333333333333 0.166710357333333  
0.499982523733333 0.333333333333333 0.166684142933333  
0.499993009493333 0.333333333333333 0.166673657173333  
0.499997203797333 0.333333333333333 0.166669462869333  
0.499998881518933 0.333333333333333 0.166667785147733  
0.499999552607573 0.333333333333333 0.166667114059093  
0.499999821043029 0.333333333333333 0.166666845623637  
0.499999928417212 0.333333333333333 0.166666738249455  
0.499999971366884 0.333333333333333 0.166666695299782  
0.49999998546754 0.333333333333333 0.166666678119913  
0.49999995418701 0.333333333333333 0.166666671247965  
0.4999999816748 0.333333333333333 0.16666668499186
```

⇒ rapid convergence;
very efficient